

A NOTE ON PERMANENTS

BY
MORTON ABRAMSON

Let $A=(a_{ij})$ be an $m \times n$ matrix and let $K=\{s_1, \dots, s_k\}$ be a k -subset from $\{1, 2, \dots, n\}$. For $0 \leq t \leq k \leq n$ define the (t, K) -permanent of A to be

$$(1) \quad \text{per}_{(t,K)}(A) = \sum a_{1i_1} a_{2i_2} \dots a_{mi_m}$$

the summation taken over all m -tuples (i_1, i_2, \dots, i_m) (repetitions allowed) of $1, 2, \dots, n$ each containing exactly t distinct entries from K and any number of distinct entries from the remaining $n-k$ integers. For example, $(4, 4, 7, 1, 1, 2)$, $(4, 4, 6, 6, 6, 5)$ are 6-tuples, each containing exactly two distinct entries from $K=\{2, 4, 5\}$ for $n \geq 7$. We define the t -permanent of A to be the case $K=\{1, 2, \dots, n\}$ and write

$$(2) \quad \text{per}_t(A) = \text{per}_{(t,K)}(A), \quad K = \{1, 2, \dots, n\},$$

each m -tuple (i_1, \dots, i_m) in the summation in (1) containing exactly t distinct entries of $\{1, 2, \dots, n\}$. When $t=m=n$, (2) is the "ordinary" permanent of a square matrix (see [1] for a survey article), while $t=m \leq n$ is the generalization to rectangular matrices described by Ryser [3].

Let A_r denote a matrix obtained from A by replacing r of the k columns s_1, s_2, \dots, s_k of A by zeros, $S(A_r)$ the product of the row sums of A_r and $\sum S(A_r)$ the sum of all the $\binom{k}{r}$ numbers $S(A_r)$. Then $\text{per}_{(t,K)}(A)$ can be evaluated by

$$(3) \quad \text{per}_{(t,K)}(A) = \sum_{i=0}^t (-1)^i \binom{k-t+i}{i} \sum S(A_{k-t+i})$$

and hence

$$(4) \quad \text{per}_t(A) = \sum_{i=0}^t (-1)^i \binom{n-t+i}{i} \sum S(A_{n-t+i}).$$

Formula (4) in the case $m=t \leq n$ was first observed by Ryser [3, p. 26] and his elegant proof (based on the Principle of Inclusion and Exclusion) suffices, with minor modification, to establish (3), so we omit the proof.

When $t=k$, (3) becomes

$$(5) \quad \text{per}_{(k,K)}(A) = \sum_{i=0}^k (-1)^i \sum S(A_i).$$

Let ${}_tA$ be a submatrix of A obtained by deleting $n-t$ columns of A . It follows that

$$(6) \quad \text{per}_t(A) = \sum \text{per}_t({}_tA),$$

Received by the editors March 18, 1970.

the summation taken over all the $\binom{n}{t}$ choices of ${}_tA$. In the case $t=m \leq n$, $\text{per}_m(A)$ is therefore equal to the sum of permanents of square submatrices of order m . It is clear that

$$(7) \quad \sum_{t=1}^n \text{per}_t(A) = A_0,$$

A_0 being the product of the row sums of A . In the case $m < t \leq n$, $\text{per}_t(A) = 0$.

We describe several of many possible applications.

Suppose that m distinct objects d_1, \dots, d_m are to be distributed into n distinct cells c_1, \dots, c_n . An $m \times n$ (0, 1) matrix $A = (a_{ij})$ can be interpreted as describing restrictions on the distribution, namely object d_i can be placed in cell c_j if and only if $a_{ij} = 1$. Suppose furthermore that t is given, and we insist that exactly t cells are nonempty. Then the number of such distributions (each satisfying conditions described by A and exactly t cells nonempty) is $\text{per}_t(A)$. For a fixed subset $K = \{s_1, \dots, s_k\}$ of $\{1, 2, \dots, n\}$, $\text{per}_{(t,k)}(A)$ is the number of distributions each satisfying conditions described by A and exactly t of the cells $c_{s_1}, c_{s_2}, \dots, c_{s_k}$ nonempty.

It is easy to see that $\text{per}_n(A) > 0$ if and only if n is equal to the term rank of A , i.e., there is a way of distributing the objects with no cell empty if and only if the maximal number of 1's in A , no two in a row or column, is n . Equivalently, $\text{per}_n(A) > 0$ if and only if the n subsets of $\{1, 2, \dots, n\}$ whose incidence matrix is the transpose of A have a system of distinct representatives.

If any object is permitted into any cell, the matrix A becomes $J_{m,n}$ all of whose entries are 1. Then the number of distributions each with exactly t of the first k cells nonempty (the remaining $n - k$ cells may or may not be empty) is, by (3) with $K = \{1, 2, \dots, k\}$, $k \leq n$,

$$(8) \quad \begin{aligned} \text{per}_{(t,k)}(J_{m,n}) &= \sum_{i=0}^t (-1)^i \binom{k-t+i}{i} \binom{k}{k-t+i} (t+n-i-k)^m \\ &= \binom{k}{t} \sum_{i=0}^t (-1)^i \binom{t}{i} (t+n-i-k)^m. \end{aligned}$$

In the case $t = k$, (8) becomes,

$$(9) \quad \text{per}_{(k,k)}(J_{m,n}) = \sum_{i=0}^k (-1)^i \binom{k}{i} (n-i)^m = \begin{cases} 0 & \text{if } 1 \leq m < k \\ k! & \text{if } m = k. \end{cases}$$

Formula (9) may also be obtained directly using the principle of inclusion and exclusion or by noting that $\text{per}_{(k,k)}(J_{m,n})$ is the coefficient of $x^m/m!$ in the expression $(e^x - 1)^k (e^x)^{n-k}$, $k \leq n$. The number of distributions with none of the n cells empty is formula (9) with $k = n$, namely the well-known formula

$$(10) \quad \text{per}_n(J_{m,n}) = \sum_{i=0}^{n-1} (-1)^i \binom{n}{i} (n-i)^m = \begin{cases} 0 & \text{if } 1 \leq m < n, \\ n! & \text{if } m = n. \end{cases}$$

This is also the number of ordered partitions of a finite set of m objects into n disjoint nonempty subsets, while the number of unordered partitions is

$$(11) \quad S(m, n) = (1/n!) \text{per}_n(J_{m,n})$$

the numbers $S(m, n)$ usually being called Stirling numbers of the second kind [2]. By (6) and (7) we have

$$(12) \quad \sum_{t=1}^n \text{per}_t(J_{m,n}) = \sum_{t=1}^n \binom{n}{t} \text{per}_t(J_{m,t}) = n^m$$

and therefore using (11)

$$(13) \quad \sum_{t=1}^n S(m, t)(n)_t = n^m, \quad (n)_t = n(n-1)\dots(n-t+1), \quad (\text{see [2]}).$$

If object d_i is placed into cell c_j with probability p_{ij} , then, letting $P=(p_{ij})$, the probability that exactly t of the cells are nonempty is $\text{per}_t(P)$.

A generalization of Montmort's "problème des rencontres" is obtained by taking $A=(a_{ij})$ with $a_{ii}=0, i=1, 2, \dots, k, k \leq \min(m, n)$ and $a_{ij}=1$ otherwise. Then the number of distributions of m objects into the n cells such that for $i=1, 2, \dots, k$ cell c_i is nonempty and object d_i is not in cell c_i , is by (5) with $K=\{1, 2, \dots, k\}$,

$$(14) \quad \text{per}_{(k,K)}(A) = D_k(m, n) = \sum_{i=0}^k (-1)^i \binom{k}{i} (n-i)^{m-k+i} (n-i-1)^{k-i},$$

with $D_0(m, n)=n^m$ while,

$$(15) \quad D(m, n) = \text{per}_n(A) = D_n(m, n) \\ = \sum_{i=0}^n (-1)^i \binom{n}{i} (n-i)^{m-n+i} (n-i-1)^{n-i}, \quad m \geq n,$$

is the number of distributions with none of the cells empty and object d_i not in cell $c_i, i=1, 2, \dots, n$. Thus, the ordinary rencontres numbers are given by

$$(16) \quad D(n, n) = \sum_{i=0}^n (-1)^i \binom{n}{i} (n-i)^i (n-i-1)^{n-i},$$

an expression noted by Ryser [3, p. 28]. It is easily seen that

$$(17) \quad \sum_{t=1}^n \binom{n}{t} D(m, t) = (n-1)^n n^{m-n}.$$

The expression for $D(m, n)$ given by (15) may be easily obtained directly using the principle of inclusion and exclusion. Denote by $g(m, n)$ the number of distributions of m objects into n cells with object d_i not in cell $c_i, i=1, \dots, n$. Clearly

$$g(m, n) = (n-1)^n n^{m-n} \quad \text{and} \quad D(m, n) = \sum_{i=0}^n (-1)^i \binom{n}{i} g(m, n-i), \quad m \geq n,$$

giving (15). Similarly (14) may be obtained directly. A second expression may be obtained for $D(m, n)$ (also by the use of inclusion and exclusion), namely

$$(18) \quad D(m, n) = \sum_{k=0}^n (-1)^k \binom{n}{k} \text{per}_{(n-k,K)}(J_{m-k,n}), \quad K = \{1, 2, \dots, n-k\}, \\ = \sum_{k=0}^n (-1)^k \binom{n}{k} \sum_{j=0}^{n-k} (-1)^j \binom{n-k}{j} (n-j)^{m-k},$$

so

$$\begin{aligned}
 (19) \quad D(n, n) &= \sum_{k=0}^n (-1)^k \binom{n}{k} \sum_{j=0}^{n-k} (-1)^j \binom{n-k}{j} (n-j)^{n-k} \\
 &= \sum_{k=0}^n (-1)^k \binom{n}{k} (n-k)! = n! \sum_{k=0}^n (-1)^k / k!.
 \end{aligned}$$

The last expression of (19) is the most common one given for the rencontres numbers [3, p. 23]. Denote by $D(m, n, r)$, $m \geq n$, the number of m -permutations of $1, 2, \dots, n$, repetitions allowed, such that each of the n integers appears at least once in each permutation with exactly r of them being in natural position. Then $D(m, n, 0) = D(m, n)$ and

$$(20) \quad D(m, n, r) = \binom{n}{r} D_{n-r}(m-r, n), \quad r = 0, 1, \dots, n,$$

while

$$(21) \quad \sum_{r=0}^n D(m, n, r) = n! S(m, n)$$

the numbers $D_k(m, n)$ and $S(m, n)$ given by (14) and (11) respectively.

In the case A is a $(0, 1)$ matrix with $m < n$ and all entries on the main diagonal zero and one elsewhere,

$$(22) \quad \text{per}_n(A) = \sum_{k=0}^{n-1} (-1)^k \sum_{u=0}^k \binom{m}{u} \binom{n-m}{k-u} (n-k)^u (n-k-1)^{m-u} = 0.$$

With regard to matching problems we have the following. Let A_1, \dots, A_m be subsets of an n -set S and define a system of t -representatives of (A_1, \dots, A_m) to be an m -tuple (a_1, \dots, a_m) containing exactly t distinct elements of S with $a_i \in A_i$, $i = 1, \dots, m$. Then the number of systems of t -representatives of (A_1, \dots, A_m) is equal to the t -permanent of the corresponding $(0, 1)$ incidence matrix of size m by n . Putting $t = \min(m, n)$, we have the number of systems of maximum distinct representatives and when $t = m \leq n$, the number of systems of distinct representatives [3, p. 54, Theorem 4.1].

REFERENCES

1. Marvin Marcus and Henryk Minc, *Permanents*, Amer. Math. Monthly, **72** (1965), 577–591.
2. J. Riordan, *An introduction to combinatorial analysis*, Wiley, New York, 1958.
3. Herbert John Ryser, *Combinatorial mathematics*, Carus Math. Monograph No. 14, 1963.

YORK UNIVERSITY,
TORONTO, ONTARIO