



Motivated political reasoning: Testing the emotion regulation account in the case of perceptual divides over politically relevant facts

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Abstract

Motivated political reasoning is a central phenomenon in political psychology, but no scholarly consensus exists as to its cause. In one influential account, motivated political reasoning is caused by goals to control emotional states. This explanation is often assumed, but has rarely been tested empirically. It implies, I argue, that individual differences in how people control their emotions (i.e., in emotion regulation strategies) should influence outcomes caused by motivated political reasoning, such as perceptual divides over politically relevant facts. I hypothesize such perceptual divides to be negatively associated with emotional acceptance and positively associated with cognitive reappraisal—two key emotion regulation strategies. I test these hypotheses in the specific context of reasoning about facts related to immigration politics in Denmark using a mix of experimental and cross-sectional survey data from three nationally representative samples of the Danish voter population (total $N = 4186$). In the specific context of the present study, the results do not support the often-assumed idea that motivated political reasoning is driven by efforts to regulate emotions. These findings raise important questions about the conditions under which emotion regulation might play a role in motivated political reasoning.

Keywords: partisan perceptual bias; motivated reasoning; emotion regulation; cognitive dissonance; public opinion; preregistration

Introduction

The theory of motivated political reasoning is crucial to much contemporary research in political psychology (Bayes et al., 2020; Reynolds, 2018) and is frequently used to explain a variety of important and diverse outcomes (e.g., Baekgaard et al., 2019; Druckman et al., 2013). One such is partisan divides in beliefs about politically relevant facts (Baekgaard et al., 2017; Bisgaard, 2015, 2019; Christensen & Moynihan, 2020; Clayton et al., 2019; Dickerson & Ondercin, 2017; Flynn et al., 2017; Flynn & Krupnikov, 2019; Hartman & Newmark, 2012; Jerit & Barabas, 2012; Khanna & Sood, 2018; Kraft et al., 2015; Miller et al., 2016; Nyhan & Reifler, 2010; Peterson & Iyengar, 2020; Robison, 2018; Schaffner & Roche, 2016; Tworzecki & Markowski, 2014; Vegetti & Mancosu, 2020), a particularly normatively concerning phenomenon, which may undermine citizens' ability to hold political elites accountable (e.g., Robison, 2018; Shapiro & Bloch-Elkon, 2008).

There are several different explanations for why motivated political reasoning occurs (see e.g., Arceneaux & Vander Wielen, 2017; Lodge & Taber, 2013; Mercier & Sperber, 2011, 2017; Stanovich, 2021; Stanovich et al., 2013; Taber & Lodge, 2006, 2016), including the prominent JQP model (Lodge &



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Taber, 2013; Taber & Lodge, 2016), where affect toward sociopolitical concepts is stored in online tallies that exert strong downstream influence on conscious deliberation (Lodge et al., 1995; Redlawsk, 2002, 2006; Redlawsk et al., 2010).

Another major theoretical explanation views motivated political reasoning as occurring to alleviate cognitive dissonance (Baekgaard et al., 2017, p. 1119; Bisgaard, 2019, p. 3; Christensen et al., 2018; Groenendyk, 2013, 2018; Kuklinski et al., 2000, p. 794; Peterson & Iyengar, 2020, p. 135; Ploger et al., 2021). Because the cognitive dissonance explanation implies that motivated political reasoning occurs to downregulate discomfort (dissonance), it casts motivated reasoning as an emotion regulation process (Cancino-Montecinos et al., 2018, 2020; Gross, 2015, 2024; Jarcho et al., 2011; Kiil, 2024; Ploger et al., 2021; Thibodeau et al., 2015; Westen et al., 2006; Westen & Blagov, 2007) where directional reasoning goals are pursued to maximize positive affective states and minimize negative affective states.

Although the emotion regulation account of motivated political reasoning has received much scholarly attention (Bianchi et al., 2016; Casado-Aranda et al., 2020; Cohen et al., 2019; Kaplan et al., 2016; Kiil, 2024; Munro et al., 2020; Thibodeau et al., 2015; Westen et al., 2006; Westen & Blagov, 2007), and is often assumed (e.g., Baekgaard et al., 2017, p. 1119; Bisgaard, 2019, p. 3; Christensen et al., 2018; Groenendyk, 2013, 2018), or argued against on theoretical grounds (Mercier & Sperber, 2017, pp. 244–246), direct empirical tests of the explanation have been very rare. Most of these are neuroimaging studies indicating that motivated political reasoning correlates with activation in brain areas linked to both implicit and explicit emotion regulation (Casado-Aranda et al., 2020; Kaplan et al., 2016; Westen et al., 2006). I therefore contend that it remains an open question whether motivated political reasoning is truly driven by emotion regulation, as is often assumed. I address this puzzle in the current study, where I investigate whether two widely studied explicit emotion regulation strategies—emotional acceptance and cognitive reappraisal—influence partisan divides in citizens' perceptions of politically relevant facts.

I draw on emotion regulation theory, which investigates how humans differ in terms of how they regulate their emotions (Gross, 2015, 2024). These differences have been found to have important consequences for a variety of psychological and behavioral variables, including well-being, social functioning, attitudes toward stigmatized groups, political tolerance, emotional reactance to political information, and political participation (Aldao et al., 2010; Alkoby et al., 2017; Ford et al., 2018; Halperin et al., 2013, 2014). Two key emotion regulation strategies are *emotional acceptance* and *cognitive reappraisal*. Emotional acceptance (known simply as “acceptance” in the emotion regulation literature) refers to noticing and embracing emotional experiences rather than taking steps to alter or avoid them (Hayes et al., 2012, p. 982; Hofmann & Asmundson, 2008, p. 5). Cognitive reappraisal refers to: “Modifying one’s appraisal of a situation in order to alter its emotional impacts” (Gross, 2015, p. 9). I theorize motivated political reasoning about political facts to be a type of cognitive reappraisal, and I theorize emotional acceptance to provide an alternative way to regulate emotions when reasoning about political facts. I therefore hypothesize that individuals who tend to use emotional acceptance to display small partisan divides in perceptions of political facts, compared to individuals who tend not to do so. I hypothesize that individuals who tend to use cognitive reappraisals display larger partisan divides in perceptions of political facts than individuals who tend not to do so.

I test these expectations through a series of cross-sectional and experimental analyses, using data from three nationally representative online surveys (total $N = 4186$) collected in Denmark. I focus on facts related to immigration politics—a highly salient and emotion-laden issue in Danish politics (Hansen & Stubager, 2017). Contrary to expectations, I find that individuals high in emotional acceptance do not display smaller perceptual divides than others do. Also contrary to expectations, I find individuals high in cognitive reappraisal to display the same or a lower level of perceptual divides than others do. Furthermore, I find that an experimental manipulation, which successfully increases emotional acceptance to have no discernible effect on perceptions of political facts.

The results suggest that motivated political reasoning might not be driven by efforts to regulate emotions as has previously been argued (Kiil, 2024; Thibodeau et al., 2015; Westen et al., 2006; Westen & Blagov, 2007) in the specific contexts of perceptual divides over politically relevant facts in Denmark. This finding is surprising as it is at odds with a large body of previous research, which assumes that voters

engage in motivated political reasoning to avoid psychological discomfort (Baekgaard et al., 2017, p. 1119; Bisgaard, 2019, p. 3; Christensen et al., 2018; Groenendyk, 2013, 2018; Kuklinski et al., 2000, p. 794; Peterson & Iyengar, 2020, p. 135; Ploger et al., 2021). The findings raise important questions about the scope of the emotion regulation explanation of motivated political reasoning, in particular, its applicability to the case of perceptual divides.

In the following section, I review existing work linking motivated political reasoning with emotion regulation before developing an account of motivated political reasoning integrated with Gross's (2015) emotion regulation framework. I derive expectations and describe the specifics of the studies before presenting the empirical results and discussing my findings.

The emotion regulation account of motivated political reasoning

Theories of motivated reasoning start with the assumption that all reasoning is motivated (Kunda, 1990), and distinguish between accuracy and directional goals—accuracy goals referring to motivations to reach correct conclusions, and directional goals referring to motivations to reach desired conclusions, such as conclusions consistent with preexisting political attitudes (e.g., Leeper & Slothuus, 2014). Following previous work, directionally motivated reasoning should lead voters to be skeptical toward information that is incongruent with their political attitudes (e.g., Taber & Lodge, 2006). In the context of perceptions of politically relevant facts, this should lead to skepticism toward facts that are incongruent with voters' political attitudes, which again should lead to perceptual divides where voters are more likely to assess information as true if it is congruent with their political attitudes than if it is incongruent (e.g., Flynn et al., 2017; Hartman & Newmark, 2012; Jerit & Barabas, 2012; Miller et al., 2016; Nyhan & Reifler, 2010; Peterson & Iyengar, 2020; Schaffner & Roche, 2016).

The argument that motivated reasoning can distort perceptions of politically relevant facts has been made in a large number of domains (Baekgaard et al., 2017; Bisgaard, 2015, 2019; Christensen & Moynihan, 2020; Clayton et al., 2019; Dickerson & Ondercin, 2017; Flynn et al., 2017; Flynn & Krupnikov, 2019; Khanna & Sood, 2018; Kraft et al., 2015; Robison, 2018; Tworzecki & Markowski, 2014; Vegetti & Mancosu, 2020), and is largely parallel to Taber and Lodge's (2006) and Taber et al. (2009) classical finding that motivated reasoning can distort evaluations of the strength of political arguments (Kraft et al., 2015).

The existing literature provides several different explanations for why motivated political reasoning occurs (see e.g., Arceneaux & Vander Wielen, 2017; Lodge & Taber, 2013; Mercier & Sperber, 2011, 2017; Stanovich, 2021; Stanovich et al., 2013; Taber & Lodge, 2006, 2016).

Among the most prominent explanations is Taber & Lodge's JQP model (Lodge et al., 1995; Lodge & Taber, 2013; Taber et al., 2009; Taber & Lodge, 2006, 2016), where affect toward sociopolitical concepts is stored in online tallies that exert a strong downstream influence on conscious deliberation (Lodge et al., 1995; Redlawsk, 2002, 2006; Redlawsk et al., 2010). In JQP, motivated political reasoning is fundamentally driven by an automatic and unconscious process called *affective contagion* (Erisen et al., 2014; Lodge & Taber, 2013, p. 135; Taber & Lodge, 2016); a memory process that is "[...] *hard wired* into the very architecture of the brain [...]" (Lodge & Taber, 2013, p. 227 emphasis in original). This implies that debiasing strategies are very unlikely to be successful (Lodge & Taber, 2013, p. 229).

It is clear from Taber and Lodge's writings (Lodge & Taber, 2013, p. 220, 227, 233) that they view JQP as qualitatively different from another influential explanation, which holds that people engage in motivated political reasoning to alleviate cognitive dissonance (e.g., Baekgaard et al., 2017, p. 1119; Bisgaard, 2019, p. 3; Christensen et al., 2018; Festinger, 1962; Groenendyk, 2013, 2018; Kuklinski et al., 2000, p. 794; Kunda, 1990; Peterson & Iyengar, 2020, p. 135; Ploger et al., 2021), i.e., to control what emotions they experience (Kiil, 2024; Thibodeau et al., 2015; Westen et al., 2006; Westen & Blagov, 2007). According to this account, humans generally seek to avoid (approach) behaviors and stimuli that become associated with negative (positive) affect. This leads to directionally motivated reasoning because people "...approach and avoid judgments based on their emotional associations" (Westen

et al., 2006, p. 1947). According to the emotion regulation account, then, one central mechanism behind motivated political reasoning is a desire to minimize negative affective states and to maximize positive affective states.

This process could potentially coexist with other mechanisms—such as the automatic activation of affective contagion emphasized in the JQP model—which could operate in parallel or under different conditions, and people might even differ in the extent to which one or the other mechanism dominates, depending on contextual cues or personal dispositions, such as cognitive style (cf. Arceneaux & Vander Wielen, 2017).

Mercier and Sperber (2017, pp. 244–246) argue against the emotion regulation explanation (which they refer to as the “feel-good account”) on theoretical grounds, stressing that it would be maladaptive for humans to adopt erroneous beliefs to obtain positive emotion, and that such a psychological adaption is therefore unlikely to have developed through the evolutionary history of humans.

Direct empirical evidence for the emotion regulation account of motivated political reasoning comes from a limited number of studies. Westen et al. (2006), Kaplan et al. (2016), and Casado-Aranda et al. (2020) found evidence linking emotion regulation and motivated political reasoning through neuroimaging studies, showing that confrontation with attitude-incongruent information increases activity in brain regions associated with implicit and explicit emotion regulation (though some of this evidence was mixed, and some of the results contradicted each other).

Other evidence for the emotion regulation account comes from studies that successfully use self-affirmations to reduce outcomes associated with motivated political reasoning (Binning et al., 2010; Cohen et al., 2000, 2007). This supports the emotion regulation account because self-affirmations provide an alternative source of global self-worth (Steele, 1988), which is thought to work through reducing the need to use motivated political reasoning to bolster self-worth (a positive emotion). However, a meta-study of more recent well-powered studies on the efficacy of self-affirmations on motivated reasoning (Lyons et al., 2021; see also Nyhan & Reifler, 2011) found no consistent debiasing effect, leaving the empirical status of this literature unclear.

Thus, while motivated political reasoning is often assumed to be driven by desires to avoid discomfort, the emotion regulation explanation has been subject to very few direct empirical tests. I argue that a testable implication of this explanation is that emotion regulation strategies should influence the extent to which people engage in motivated political reasoning, which, in turn, should influence perceptual divides over political facts. In the following, I theorize how two important emotion regulation strategies (cognitive reappraisal and emotional acceptance) should be associated with perceptual divides if the emotion regulation account of motivated political reasoning is correct. In Gross’s (Gross, 2015) seminal emotion regulation framework, cognitive reappraisals are thoughts—or patterns of thoughts—that target the meaning of a potentially emotion-eliciting situation or its self-relevance. This, I argue, must include motivated political reasoning if—as the emotion regulation account implies—accepting attitude-incongruent political information leads to negative emotions (and accepting attitude-congruent political information leads to positive emotions) all else equal (e.g., Cohen, 2003; Cohen et al., 2000; Groenendyk, 2013, 2018). The main thrust of the argument is that voters who strongly support a policy or a party will experience cognitive dissonance (or related types of discomfort such as self-threats) if they are led to believe something that reflects negatively on that party or policy (Cohen, 2003; Cohen et al., 2000; Devine et al., 1999; Festinger, 1962). This discomfort can be alleviated by altering one’s attitude toward the policy or the party in question, but this may be costly and difficult for voters with strong attitudes (e.g., Groenendyk, 2013, pp. 16–17). Instead, they can alleviate the discomfort by pursuing directional reasoning goals to avoid ending up with beliefs that contradict their political attitudes. This implies that directionally motivated political reasoning corresponds to the emotion regulation strategy of cognitive reappraisal since it involves modifying cognitive patterns to alter the emotional impact of political information (cf. Gross, 2015).

The emotion regulation account of motivated political reasoning, therefore, implies that the more people tend to use cognitive reappraisals, the more they should tend to engage in directional motivated reasoning. Since directional motivated reasoning is generally used to explain perceptual divides

(cf. e.g. Flynn et al., 2017; Hartman & Newmark, 2012; Jerit & Barabas, 2012; Miller et al., 2016; Nyhan & Reifler, 2010; Peterson & Iyengar, 2020; Schaffner & Roche, 2016), perceptual divides should tend to be larger the more people tend to use cognitive reappraisal.

Moving on to emotional acceptance, this concerns the counterintuitive idea of calmly focusing one's attention on uncomfortable feelings and thoughts rather than trying to alter or avoid them (Hayes et al., 2012, p. 982; Hofmann & Asmundson, 2008, p. 5). Emotional acceptance thus represents a fundamentally different way of regulating emotions than cognitive reappraisals, as one does not have to alter one's interpretation of the situation at hand; rather, one has to alter the way one relates to thoughts and emotions involved in and evoked by one's appraisal of the situation.

While cognitive reappraisal and emotional acceptance can operate at both implicit and explicit levels, the present study examines these strategies in their explicit form—that is, as consciously endorsed and effortful ways of managing emotional responses. Implicit processes, on the other hand, occur automatically and without conscious intent (Bargh & Williams, 2007; Tamir et al., 2007). Indeed, emotion regulation itself can operate as an unconscious goal (Hopp et al., 2011), shaping how individuals attend to and interpret emotionally salient political information. Importantly, the distinction between implicit and explicit emotion regulation is not always clear-cut (Gross, 2015; Gyurak et al., 2011). Prior research suggests that frequent conscious use of a strategy—such as emotional acceptance—can lead to automatization, meaning that over time it becomes habitual and potentially automatic (Desbordes et al., 2012; Gyurak et al., 2011). Thus, although this study focuses on explicit manifestations of these strategies (via self-report and experimental priming), they may reflect broader, partly implicit regulatory tendencies.

In politics, I argue, emotional acceptance implies noticing and embracing emotions and thoughts elicited by processing political information without trying to alter or escape them. This includes noticing and embracing discomfort associated with believing attitude-incongruent information. An anti-immigration voter, for instance, would notice and embrace the feelings of anxiety and self-threat elicited by earnestly listening to a factual claim that challenges an anti-immigration stance without taking steps to avoid or alter those feelings. Therefore, using emotional acceptance as an emotion regulation strategy when processing attitude-incongruent political information should enable the voter to believe attitude-incongruent political information and simultaneously downregulate negative emotions. Voters prone to using emotional acceptance should therefore engage less in motivated political reasoning when assessing political facts and therefore display smaller perceptual divides. This leads to the following hypothesis:

H1: Individual differences in emotional acceptance are negatively associated with perceptual divides in the evaluation of politically relevant facts.

People who are prone to using cognitive reappraisal, on the other hand, should engage more in motivated political reasoning when assessing political facts and should therefore display larger perceptual divides. This leads to the second hypothesis:

H2: Individual differences in cognitive reappraisal are positively associated with perceptual divides in the evaluation of politically relevant facts.

While hypothesis 1 and hypothesis 2 concern individual differences in people's general tendency to use emotion regulation strategies, people's use of emotion regulation strategies can change from situation to situation. Many psychotherapeutic interventions aim at altering clients' emotion regulation habits by promoting the use of acceptance or associated emotion regulation strategies (e.g., Acceptance and Commitment Therapy (Hayes et al., 2012), Mindfulness-Based Stress Reduction (Kabat-Zinn, 2003; Kabat-Zinn, 1982), and Metacognitive Therapy (Wells, 2011)). We should therefore expect situational features, which influence use of emotion regulation strategy, to influence motivated political reasoning and thus also to influence perceptions of politically relevant facts. This leads to hypothesis 3:

H3: Promoting use of emotional acceptance decreases perceptual divides in the evaluation of politically relevant facts.

For H1 to find support, we should expect individuals high in emotional acceptance to be more likely to believe attitude-incongruent facts and to display a lower difference between belief in attitude-congruent facts and attitude-incongruent facts (i.e., a smaller perceptual divide). For H2 to find support, individuals high in cognitive reappraisal should be less likely to believe attitude-incongruent facts and display a larger difference between belief in attitude-congruent facts and attitude-incongruent facts (i.e., a larger perceptual divide).

Finally, for H3 to find support, performing an emotional acceptance-inducing task should lead to greater belief in attitude-incongruent facts and a smaller difference between belief in attitude-congruent facts and attitude-incongruent facts.

Overview of the studies

I test the hypotheses in the context of perceptions of facts related to immigration politics in Denmark. Though Denmark is a relatively consensus-based multiparty system, immigration politics is a divisive and highly salient issue, which has dominated party politics over the last decade (cf. e.g., Hansen & Stubager, 2017). I contend that the immigration debate in Denmark tends to elicit strong emotions and that directionally motivated reasoning and perceptual divides, therefore, are very likely to be triggered in this context.

In all three studies, I recruited respondents from Dynata’s (formerly SSI) panel. Participants were invited to answer an online survey for economic compensation (see Table 1 for an overview of the studies). In study 1, I collected 1254 subjects in the fall of 2019. In studies 2 and 3, I collected 1314 and 1618 subjects in early and mid-2020, respectively. Sample sizes were decided based on power analyses. All samples are approximately representative of the Danish voting age population in terms of age, gender, education, and income (see Supplementary SI 1 for sample characteristics). In all studies, I use an instructional attention check (Oppenheimer et al., 2009) to weed out respondents who do not pay sufficient attention to the question texts (see Supplementary SI 2 for details). Only respondents who pass the attention check are counted as part of the full sample and included in the analyses. Studies 2 and 3 were preregistered (see Supplementary SI 3 for minor deviations from preregistrations not described in the main text).

In study 1, I investigate hypothesis 1 through a cross-sectional analysis with perceptual divides as the dependent variable. The evaluated facts are attributed to a non-expert, partisan source (leftwing or rightwing politicians). In study 2, I use a mix of cross-sectional and experimental analyses to investigate hypotheses 1, 2, and 3. This time, the facts are attributed to an expert, non-partisan source (Statistics Denmark). In this study, I focus entirely on attitude-incongruent facts. In study 3, I investigate hypotheses 1 and 2 with a mix of cross-sectional and experimental analyses. I use both attitude-

Table 1. Overview of studies

Study	Hypotheses tested	N	Type of sample	Facts evaluated	Source cue	Preregistration
1	1	1254	Nationally representative	Attitude-congruent and -incongruent	Non-expert partisan	-
2	1, 2, and 3	1314	Nationally representative	Only attitude-incongruent	Expert non-partisan	https://osf.io/3bn78/?view_only=9f2d14ce12264eb99af5cba672cd0d34
3	1 and 2	1618	Nationally representative	Attitude-congruent and -incongruent	Non-expert partisan and expert non-partisan (randomized)	https://osf.io/869za/?view_only=703b40ca868d41a7bb978d0d0d179f30

congruent and attitude-incongruent facts, and whether the facts are attributed to an expert or a non-expert source is determined through random assignment.

Study 1

Study 1 focuses on H1 and uses a cross-sectional design to test whether individual differences in emotional acceptance predict belief in attitude-incongruent facts and perceptual divides. Following such studies as Peterson and Iyengar (2020), I use a mix of correct and incorrect factual claims that have received significant public attention. I use two pro-immigration factual claims and two anti-immigration factual claims. The correct pro-immigration claim reads: “The right-wing’s ‘Start help’ [a renowned Danish social program], which halved unemployment benefits for immigrants, has increased poverty-related crimes and led many immigrant women to exit the labor force” whereas the incorrect pro-immigration claim reads: “Support for equality between men and women is just as high among non-western immigrants and descendants as among ethnic Danes.” The correct anti-immigration claim reads: “Non-western immigration costs the welfare state approximately 33 billion kroner a year. This is approximately three times as much as is used on police every year.” The incorrect anti-immigration claim reads: “Non-western immigrants are on average six times as likely to commit punishable crimes as ethnic Danes.”

Pro-immigration statements were introduced with the text: “Politicians on the political left have made the following claim.” Anti-immigration statements were introduced with the text: “Politicians on the political right have made the following claim.” Respondents were asked to evaluate to what extent they agree that the statement is correct on a five-point Likert scale (cf. e.g., Nyhan & Reifler, 2010).

I create an index for belief in attitude-incongruent facts ($M = .36$, $SD = .25$) and for attitude-congruent facts ($M = .65$, $SD = .19$), both ranging from 0 to 1. Attitude-congruency of each statement is determined by dividing respondents into anti- and pro-immigration respondents (dependent on whether their mean score on the immigration attitudes index was higher than .5). I create a perceptual divide measure by subtracting belief in attitude-incongruent facts from belief in attitude-congruent facts ($M = .65$, $SD = .19$), ranging from 0 to 1 with high values indicating a large perceptual divide.

I create an immigration attitudes index ($\alpha = .81$, $M = .62$, $SD = .27$) ranging from 0 to 1, with high values indicating anti-immigration attitudes. I used a 10-point immigration attitude self-placement scale and two five-point Likert scale items concerning punishments for criminal immigrants and whether immigration constitutes a threat.

To measure individual differences in trait emotional acceptance, I use the acceptance subcomponent of the Cognitive Affective Mindfulness Scale—Revised (CAMS-R) (Feldman et al., 2007). The CAMS-R is a validated measure that correlates with mindfulness and well-being and negatively predicts anxiety, distress, and depression (Feldman et al., 2007). The acceptance subscale includes three items: “I am able to accept the thoughts and feelings I have,” “I can accept things I cannot change,” and “I can tolerate emotional pain.” Scores are recorded on a scale from 1 (almost never) to 5 (almost always). I create an index ranging from 0 to 1 ($\alpha = .80$, $M = .66$, $SD = .24$) with high values indicating high emotional acceptance.

I control for standard sociodemographic factors (education, age, gender, income, and ethnicity) and a series of psychological covariates, which have been shown in previous studies to be important predictors of motivated political reasoning: political sophistication, attitude strength (Taber & Lodge, 2006), need for cognition, and need for affect (Arceneaux & Vander Wielen, 2013, 2017). I measure political sophistication with four knowledge questions from the Danish national election study conducted in 2019 (Hansen, 2021), e.g., “Which parties were in government up until the election in 2019?” I measure attitude strength with two questions about how strong participants’ attitudes on immigration politics are, such as “How much do you personally care about the issue of immigration?” I measure need for cognition with three items from Cacioppo et al. (1984), such as “I like to have the responsibility of handling a situation that requires a lot of thinking.” To measure need for affect, I use three questions

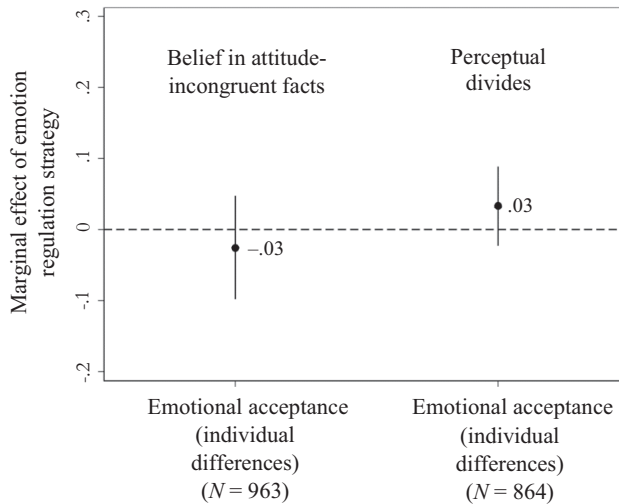


Figure 1. Study 1 results. Effect of emotional acceptance on belief in attitude-incongruent facts and perceptual divides in study 1. Estimates are unstandardized OLS regression coefficients, and vertical lines indicate 95% confidence intervals. Immigration attitude, political sophistication, attitude strength, age, gender, income, ethnicity, and education are included as covariates.

from each of the two dimensions in Need For Affect (approach and avoidance) in Maio & Esses (2001), such as “I am a very emotional person.” See [Supplementary SI 4](#) for full question wordings and [Supplementary SI 5](#) for descriptive statistics for measures where these are not provided in the main text. Supplementary see [SI 6](#) for pairwise correlations between continuous predictors.

Results. As expected, respondents believe attitude-congruent facts much more than attitude-incongruent facts. On average, respondents report 31 percentage points more belief in attitude-congruent than attitude-incongruent facts ($p < .001$). This shows that perceptual divides identified by previous studies in other contexts (Bisgaard, 2015; Jerit & Barabas, 2012; Miller et al., 2016; Nyhan & Reifler, 2010; Peterson & Iyengar, 2020; Robison, 2018; Tworzecki & Markowski, 2014; Vegetti & Mancosu, 2020) replicate in Denmark.

[Figure 1](#) displays the association between emotional acceptance and belief in attitude-incongruent facts and perceptual divides (Supplementary see [SI 7](#) for full regression tables for all three studies). Emotional acceptance is not associated with either belief in attitude-incongruent facts ($b = -.03$, $p = .49$) or perceptual divides ($b = .03$, $p = .24$) at levels approaching statistical significance. These results contradict the theoretical expectation expressed in H1, as emotional acceptance neither predicts belief in attitude-incongruent facts nor predicts perceptual divides. This suggests that emotional acceptance does not reduce partisan perceptions of political facts, at least in the context of facts related to immigration politics in Denmark.

The factual claims used in study 1 are accompanied by a non-expert source cue. Given the relatively low credibility of the source cues (a partisan source), one could argue that it is relatively rational (or reasonable) to be skeptical of the attitude-incongruent factual claims. While perceptual divides, such as those studied here, have previously been attributed to motivated political reasoning (Flynn et al., 2017; Hartman & Newmark, 2012; Jerit & Barabas, 2012; Miller et al., 2016; Nyhan & Reifler, 2010; Peterson & Iyengar, 2020; Schaffner & Roche, 2016), others have argued that they may be attributable to more rational processes (e.g., Druckman & McGrath, 2019; Gerber & Green, 1999). This implies a risk that the ineffectiveness of emotional acceptance may be due to the perceptual divides being driven by rational processes—and thus not by motivated reasoning. One way to accommodate this issue is to use a highly credible expert source and only use true factual claims. Perceptual divides, which persist even under such conditions, are more likely to be caused by motivated political reasoning, as there are fewer “reasonable” grounds for disbelieving the facts.

Study 2

In study 2, I augment the research design used in study 1 by attributing all factual statements to an expert source and by only using true factual claims. Furthermore, to ensure that the factual claims arouse emotions, I accompany each factual statement with a graph (Supplementary see SI 4 for details) illustrating the claim to make it more tangible (Nyhan & Reifler, 2018). Following Khanna and Sood (2018), in this study, I only expose respondents to attitude-incongruent factual claims. I do so to focus narrowly on the specific context where emotional acceptance is theorized to influence reasoning.

In this study, I also test H2 by investigating whether individual differences in cognitive reappraisal are associated with low belief in attitude-incongruent facts. Moreover, I test H3 by investigating whether situational cues prompting the use of emotional acceptance increase belief in attitude-incongruent facts. The test of H3 occurs through an experiment where participants in the treatment group are prompted to use emotional acceptance to regulate emotions while evaluating a set of attitude-incongruent political facts (see Supplementary SI 6 for a balance test).

The procedure is as follows: After reporting immigration attitudes, emotion regulation strategies, and covariates, participants read and evaluate two attitude-incongruent factual statements. Participants in the treatment group then listen to an audio clip that guides them through a short exercise in accepting emotions, and they are invited to use emotional acceptance when subsequently evaluating two other attitude-incongruent facts. Participants in the control group engage in a guided mind-wandering task before evaluating the two other attitude-incongruent facts. All participants then answer a manipulation check before reading a thorough debriefing.

Respondents only evaluate attitude-incongruent facts as determined by their immigration attitude position. The facts selected for study 2 are reported in Table 2. They are designed to elicit strong emotional reactions, to be very congenial to one side of the immigration debate, and to be correct. In all cases, the facts are attributed to Statistics Denmark (a highly authoritative expert source), which is also their true source, and they are accompanied by a graph illustrating the claim (see Supplementary SI). I create an index for belief in the attitude-incongruent facts used before the intervention ($M = .45$, $SD = .29$), one for the facts used after the intervention ($M = .46$, $SD = .3$), and one for all items combined ($M = .45$, $SD = .25$). All indices range from 0 to 1.

Designing an experimental manipulation strong enough to impact emotion regulation strategy in an online survey is a challenge, but has been accomplished successfully before (Ford et al., 2018;

Table 2. Facts evaluated in studies 2 and 3

Facts evaluated before intervention	
Anti-immigration	It is four times as frequent among ethnic newcomers (immigrants and decedents with non-western background) as among ethnic Danes, to favor raising boys with more freedom than girls.
Anti-immigration	Approximately one in five ethnic newcomers do not think it is good to have a representative democracy in Denmark.
Pro-immigration	Only 2 percent of young, ethnic newcomers (immigrants and descendants with non-western background) are not in favor of gender equality.
Pro-immigration	Second generation immigrants support freedom of speech as much as ethnic Danes.
Facts evaluated after intervention	
Anti-immigration	People with non-western background are between four and five times as likely to have violated the criminal code as ethnic Danes.
Anti-immigration	Non-western immigration costs the welfare state around 33 billion kroner a year—approximately three times as much as total expenditure on policing.
Pro-immigration	Total costs associated with non-western immigration correspond to a small fraction of yearly public expenditure (about three percent).
Pro-immigration	The vast majority, approximately 97 percent, of non-western immigrants and descendants have never broken the law.

Hafenbrack et al., 2014; see also Halperin et al., 2013). Participants in both the treatment and the control group listen to a 3 minutes 30 seconds long audio file. In the treatment group, the audio file guides respondents to scan their body for an unpleasant sensation, to focus their attention on that sensation, and to let go of the struggle with the sensation (to accept it). The manipulation is based on a transcript of a guided emotional acceptance exercise from Harris (2011). In the control condition, the audio file guides participants through a short mind-wandering task, designed to mimic a “waking, baseline mental state” (Hafenbrack et al., 2014, p. 371). For transcripts of both audio files, see [Supplementary SI 8](#).

As a manipulation check, I use three items from the Toronto Mindfulness Scale (TMS), which measure “decentering”—a concept closely associated with emotional acceptance (Lau et al., 2006). The TMS is a state measure rather than a dispositional measure, which makes it useful as a subjective manipulation check in this context. An example of an item is “I approached each feeling by trying to accept it, no matter whether it was pleasant or unpleasant.” The index ranges from 0 to 1 ($\alpha = .7$, $M = .63$, $SD = .21$). I pilot-tested the acceptance manipulation on a sample of 213 respondents on Amazon M-Turk, and I found a 10 percentage point increase in TMS of the treatment ($p < .01$).

To measure cognitive reappraisal, I use the emotion regulation questionnaire (ERQ) developed by Gross & John (2003, p. 403). The ERQ is a validated measure (Aldao et al., 2010) with two subscales; one measuring the use of cognitive reappraisal, and one measuring suppression. I only use the four items measuring cognitive reappraisals: (1) “When I want to feel more positive emotion (such as joy or amusement), I change what I’m thinking about”; (2) “when I want to feel less negative emotion (such as sadness or anger), I change what I’m thinking about”; (3) “when I want to feel more positive emotion, I change the way I’m thinking about the situation”; and (4) “when I want to feel less negative emotion, I change the way I’m thinking about the situation.” Scores are reported on a Likert-type five-point scale. I create an index ranging from 0 to 1 ($\alpha = .85$, $M = .54$, $SD = .25$).

Individual differences in emotional acceptance are operationalized the same way as in study 1 ($\alpha = .70$, $M = .67$, $SD = .21$). I again control for political sophistication (measured the same way as in study 1) and attitude strength as well as income, education, gender, and age (see [Supplementary SI 4](#) for details).

Results. As expected, participants with moderate immigration attitudes believe attitude-incongruent political facts much more than participants with strong and extreme attitudes. Attitude extremity decreases belief in attitude-incongruent facts by 27 percentage points ($p < .001$). This is similar to the results in study 1.

[Figure 2](#) shows the associations between emotion regulation strategy and belief in attitude-incongruent facts. Individual differences in emotional acceptance predict approximately 12 percentage points ($p < .01$) increased belief in the attitude-incongruent facts evaluated before the treatment. This supports H1 as emotional acceptance is associated with high belief in attitude-incongruent facts, and, thus contradicts the results found in study 1. Individual differences in cognitive reappraisal, in contrast, are unrelated to belief in attitude-incongruent facts ($p = .42$). This contradicts the expectation in H2, according to which the two should be negatively associated.

Moving on to the effect of situational emotional acceptance, the results show that the experimental treatment has no effect on belief in attitude-incongruent facts ($p = .40$), although it leads to a 6 percentage points increase in the manipulation check ($p < .001$). This indicates that situational cues increasing emotional acceptance do not increase belief in attitude-incongruent facts, which contradicts H3.

The results in studies 1 and 2 are inconsistent, as individual differences in emotional acceptance were associated with belief in attitude-incongruent facts in study 2 but not in study 1. This raises the question of whether this inconsistency is due to sampling error or to differences in the design of the two studies, such as the use of a non-expert source cue in study 1 and an expert source cue in Study 2. To resolve this issue, I conducted a third study, which provides a final test of the influence of individual differences in emotional acceptance on perceptual divides. Study 3 also provides an additional test of H2 (the effect of individual differences in cognitive reappraisal), which did not find support in Study 2.

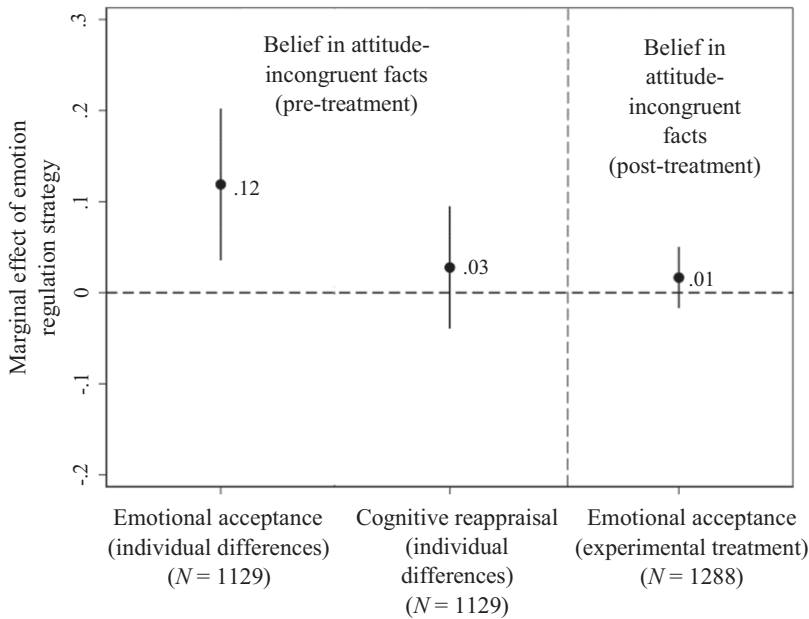


Figure 2. Study 2 results. Effect of emotional acceptance and cognitive reappraisal on belief in attitude-incongruent facts in study 2. Estimates are unstandardized OLS regression coefficients, and vertical lines indicate 95% confidence intervals. Immigration attitude, political sophistication, attitude strength, age, gender, income, and education are included as covariates. All models were preregistered.

Study 3

Study 3 provides two core methodological advantages over Study 2. First, I experimentally manipulate whether the facts are attributed to an expert source or not. This enables testing whether individual differences in emotional acceptance are only associated with belief in attitude-incongruent facts when they are attributed to an expert source. Second, participants evaluate both congruent and incongruent facts (i.e., all eight facts displayed in Table 2), thus enabling a comparison of belief in attitude-congruent and attitude-incongruent facts.

Participants are randomly assigned to one of two conditions: one where all facts are attributed to a non-expert partisan source, as in study 1, or one where all facts are attributed to the same expert source as in study 2. I use the same factual statements as used in study 2 and the same graphical illustrations. Immigration attitudes and attitude strength are measured as in study 1, and remaining individual differences variables are measured as in study 2 (see Supplementary SI 4, 5, and 6 for details, descriptive statistics, and balance test).

Results. As in studies 1 and 2, participants report 24 percentage points ($p < .001$) higher belief in attitude-congruent than attitude-incongruent facts and thus display clear perceptual divides. Figure 3 shows the effect of individual differences in emotional acceptance and cognitive reappraisal on belief in attitude-incongruent facts and perceptual divides, dependent on whether the facts were attributed to an expert source or a non-expert source. Individual differences in emotional acceptance have no significant interaction effect with expert source on belief in attitude-incongruent facts ($p = .15$), on belief in attitude-congruent facts ($p = .45$), or on perceptual divides ($p = .64$). Furthermore, emotional acceptance has no significant effect on attitude-incongruent facts, regardless of whether the source is expert or non-expert ($p > .05$ in both cases). Emotional acceptance does have a significant positive effect on belief in attitude-congruent facts ($b = .85$, $p = .029$) when the source is expert but not when the source is non-expert ($p = .19$). This adds up to a non-significant association between emotional acceptance and perceptual divides (belief in attitude-congruent facts minus belief in attitude-incongruent facts), regardless of

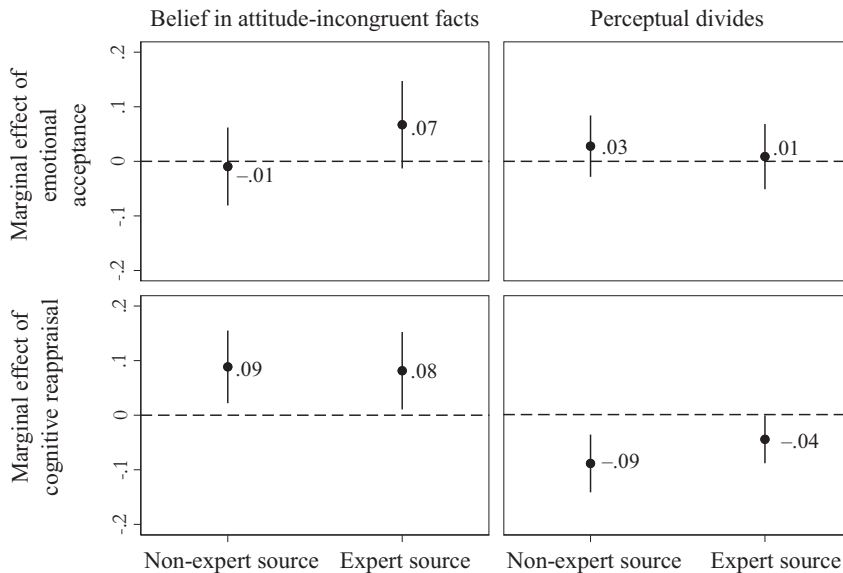


Figure 3. Study 3 results. Effect of emotional acceptance and cognitive reappraisal on belief in attitude-incongruent facts and perceptual divides in study 3 ($N = 1407$). Estimates are unstandardized OLS regression coefficients, and vertical lines indicate 95% confidence intervals. Immigration attitude, political sophistication, attitude strength, age, gender, income, and education are included as covariates. Analyses of effects of emotional acceptance were preregistered ($N = 1477$); analyses of effects of cognitive reappraisal were not ($N = 1407$).

whether the source is expert or non-expert ($p > .05$ in both cases). These results contradict the preregistered expectations and H1. The results indicate that the divergence between the findings in studies 1 and 2 did not stem from the difference in source attribution but from something else, possibly sampling error. The results indicate that emotional acceptance does not reduce perceptual divides over politically relevant facts as theorized.

Cognitive reappraisal, in contrast, is significantly positively associated with belief in attitude-incongruent facts and negatively associated with perceptual divides. While I find no sign of an interaction effect with expert source, the effect-sizes at different source conditions range between 4 and 9 percentage points ($p < .05$ in all cases) and between 8 and 7 percentage points ($p < .001$ in all cases) when pooling observations with expert and non-expert source cues (these analyses were not preregistered). These results are exactly the opposite of the expectation in H2 as I find cognitive reappraisal to be negatively associated with perceptual divides and positively associated with belief in attitude-incongruent political facts. The results are inconsistent with the findings in study 2 where cognitive reappraisal was not significantly associated with belief in attitude-incongruent political facts.

Across all three studies, I thus find that neither individual differences in emotional acceptance nor in cognitive reappraisal have a consistent association with belief in attitude-incongruent facts or with perceptual divides. To the extent that cognitive reappraisals do predict perceptual divides, the association is in the exact opposite direction of the theoretical expectation. Furthermore, an experimental manipulation, which successfully increases emotional acceptance, does not increase belief in attitude-incongruent facts. I thus do not find support for any of the theoretical expectations derived from the emotion regulation account of motivated political reasoning.

Discussion

The results suggest that the emotion regulation strategy does not influence voters' perceptions of political facts.

This, I argue, indicates that voters do not engage in motivated political reasoning to regulate their emotions in the particular context investigated in this research. In the following, I discuss three potential objections against this conclusion as well as the implications of my findings for motivated political reasoning theory.

A first potential objection is that motivated political reasoning might not be the cause of the large perceptual divides I find between pro- and anti-immigration voters (and between voters with moderate and extreme immigration attitudes). If these perceptual divides instead stem from partisan cheerleading (see e.g., Bullock et al., 2015), it is not clear that the emotion regulation account of motivated political reasoning can be tested by investigating whether emotion regulation strategy influences perceptual divides. This objection holds true if perceptual divides are driven entirely by partisan cheerleading, but not if a substantial part of perceptual divides stems from motivated reasoning. We can be fairly sure that perceptual divides are at least partially caused by motivated political reasoning because a host of studies have found perceptual divides to persist even when research participants are rewarded financially for giving correct answers to factual political questions (e.g., Bullock et al., 2015; Hill, 2017; Khanna & Sood, 2018), and because recent research has found support for the motivated reasoning explanation of perceptual divides rather than competing explanations (Berinsky, 2018; Peterson & Iyengar, 2020).

A second potential objection is that the present research used explicit measures and an explicit experimental manipulation, while motivated political reasoning could be an entirely implicit (i.e., unconscious) emotion regulation process. Therefore, some might be concerned that the measures and the experimental treatment used here do not tap the intended process. This is unlikely, I argue, because previous research has shown that explicit self-report scales (and interventions) *do* measure (and influence) implicit emotion regulation processes (Desbordes et al., 2012; Drabant et al., 2009; Guendelman et al., 2017), and because emotion regulation strategies involve both implicit and explicit processes that are closely connected and influence each other (Braunstein et al., 2017; Koole et al., 2015).

A third potential objection is that individuals might use other strategies to minimize negative affect beyond simply rejecting or disbelieving attitude-incongruent information. For example, individuals could reduce the emotional impact of such information by down-weighting its importance (Festinger, 1962), such that emotion regulation strategies do not manifest themselves through outright dismissal of facts, but rather through more subtle cognitive adjustments that preserve preexisting attitudes while maintaining some degree of factual acceptance (Bisgaard, 2015). Such processes blur the line between directionally motivated reasoning and accuracy-based goals, as individuals may be motivated to protect their identity by devaluing the information without necessarily rejecting its truthfulness.

Future research should take this into account by also investigating the relationship between emotion regulation strategies and the extent to which individuals alter their attributions of importance of attitudinally congruent or incongruent political information, when presented with such information.

Scope conditions. It is important to note that the results should not be interpreted as a broad refutation of the emotional regulation account of motivated political reasoning, but rather as a limited falsification within the specific context that this study investigates. While the results suggest that emotion regulation processes do not underlie motivated reasoning about political facts related to immigration politics in Denmark, it is important to recognize that this conclusion applies only to a particular type of motivated reasoning (about political facts), a relatively narrow sample (Danish voters), and a single political issue (politics of immigration). Emotion regulation strategies may play an important role in motivated political reasoning in other political contexts with different populations or on different political issues. Future research is needed to determine whether the findings generalize beyond this particular setting.

Furthermore, the theoretical argument of the present research relies on the assumption that directional reasoning motivations are to some extent present when the research participants evaluate the political facts. If research participants are solely accuracy motivated, there is no reason to expect emotion regulation strategies to influence their evaluations of the factual claims. This theoretical assumption seems to be consistent with the observed data, as left- and right-leaning participants disagree to a considerable extent in their factual evaluations, similar to what other studies on motivated reasoning and evaluations of politically relevant facts have found (e.g. Bisgaard, 2019; Flynn & Krupnikov, 2019;

Miller et al., 2016; Peterson & Iyengar, 2020). Nonetheless, the present studies did not directly assess or manipulate participants' motivational goals, and the relative strength of directional versus accuracy motivations may moderate the impact of emotion regulation strategies. Future research could investigate this possibility by experimentally manipulating the salience of directional or accuracy goals and testing their interaction with emotion regulation strategies (Bolsen et al., 2014; Bullock et al., 2015; Khanna & Sood, 2018; Peterson & Iyengar, 2020).

Implications. The results in this study raise important questions about a widespread theoretical understanding of motivated political reasoning. While the idea of categorizing motivated political reasoning as emotion regulation is relatively new (Kiil, 2024; Thibodeau et al., 2015; Westen et al., 2006; Westen & Blagov, 2007), the underlying idea that people engage in biased reasoning to avoid or reduce psychological discomfort can be traced back to Festinger (1962), and much research since has assumed that the “fuel” driving motivated political reasoning is a desire to avoid such psychological discomfort (Baekgaard et al., 2017, p. 1119; Bisgaard, 2019, p. 3; Christensen et al., 2018; Groenendyk, 2013, 2018; Kuklinski et al., 2000, p. 794; Kunda, 1990; Peterson & Iyengar, 2020, p. 135; Ploger et al., 2021). Nonetheless, this explanation of motivated political reasoning has very rarely been subject to stringent empirical tests, and the present study is the first to investigate whether emotion regulation habits influence outcomes caused by motivated political reasoning. The fact that the emotion regulation account does not find support in the specific context of the present study suggests that we should test further the widespread assumption that one of the central mechanisms behind motivated political reasoning is a desire to avoid experiencing psychological discomfort.

To the extent that motivated political reasoning is not driven by efforts to regulate emotions—as the results in the specific context of the present study suggest—what drives it then? Perhaps the most prominent alternative to the emotion regulation explanation of motivated political reasoning is Lodge and Taber (2013)'s JQP model. In JQP, motivated political reasoning occurs because of how the architecture of human memory is designed (Lodge & Taber, 2013, p. 227). According to this model, the fundamental driver of motivated reasoning in politics is an automatic and unconscious process called *affective contagion* (Lodge & Taber, 2013, p. 135); a memory process that is “[...] *hard wired* into the very architecture of the brain [...]” (Lodge & Taber, 2013, p. 227 emphasis in original). This makes motivated political reasoning an unavoidable part of the human condition, and debiasing attempts very unlikely to be successful (Lodge & Taber, 2013, p. 229). The present study does not test the JQP model, but the results are not inconsistent with it, since JQP implies that conscious attempts to regulate affective impulses should generally be unsuccessful (Lodge & Taber, 2013, p. 229).

It is also worth noting that the two mechanisms discussed in this article—emotion regulation and affective contagion (as proposed by the JQP model)—are not necessarily mutually exclusive. They might operate in parallel, alternate depending on the context, or vary across individuals. Arceneaux and Vander Wielen (2017) argue that people vary in their disposition toward cognitive reflection, which may influence the extent to which they engage in conscious emotion regulation versus relying on automatic affective processes.

Conclusion

Contemporary theories in political psychology claim that political reasoning is strongly influenced by affective processes, which lead voters to defend their preexisting political attitudes when confronted with political information. This, in turn, has problematic downstream consequences such as perceptual divides over politically relevant facts, which may undermine accountability. An influential account of motivated political reasoning claims that voters engage in motivated political reasoning to control which emotions they experience when they process political information. This account is relied upon by many, but has rarely been tested. In this paper, I have provided the first empirical test of whether the strategies voters use to regulate their affective impulses influence outcomes caused by motivated political

reasoning, focusing specifically on how emotional acceptance and cognitive reappraisal influence perceptual divides over facts related to immigration politics in Denmark.

I did not find empirical support for the theoretical expectations derived from the emotion regulation account of motivated political reasoning. Neither individual differences in emotional acceptance nor in cognitive reappraisal were consistently associated with belief in attitude-incongruent facts or with perceptual divides. Furthermore, an experimental treatment, which successfully increased emotional acceptance, had no discernible effect whatsoever.

These results indicate that in the specific context of perceptions of facts related to immigration politics, Danish voters do not engage in motivated political reasoning to regulate their emotions. The results in this specific context are thus at odds with previous research, which has argued or assumed that emotion regulation underlies motivated political reasoning (e.g., Thibodeau et al., 2015; Westen et al., 2006). The findings are not, however, inconsistent with Lodge and Taber's (2013) JQP model, where motivated political reasoning stems from the basic architecture of human memory, not from conscious or unconscious attempts to regulate emotions.

Thus, while emotion regulation has previously been found to have benevolent effects on such important political variables as tolerance, support for conflict resolution, and coping in the face of emotion-eliciting political information, it does not seem to provide a panacea for all democratically undesirable consequences of motivated political reasoning, such as perceptual divides over politically relevant facts.

Supplementary material. The supplementary material for this article can be found at <http://doi.org/10.1017/pls.2025.10001>.

Data availability statement. This article was awarded Open Data, Open Materials, and Preregistration badges for open science practices. The data, coding files, and pre-registration are publicly available at: https://osf.io/mcrg4/files/osfstorage?view_only=None, https://osf.io/3bn78/?view_only=9f2d14ce12264eb99af5cba672cd0d34 and https://osf.io/869za/?view_only=703b40ca868d41a7bb978d0d0d179f30

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