

SOME DISTRIBUTIONAL RESULTS FOR POISSON–VORONOI TESSELLATIONS

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Abstract

We consider the Voronoi tessellation based on a stationary Poisson process N in \mathbb{R}^d . We provide a complete and explicit description of the Palm distribution describing N as seen from a randomly chosen (typical) point on a k -face of the tessellation. In particular, we compute the joint distribution of the $d - k + 1$ neighbours of the k -face containing the typical point. Using this result as well as a fundamental general relationship between Palm probabilities, we then derive some properties of the typical k -face and its neighbours. Generalizing recent results of Muche (2005), we finally provide the joint distribution of the typical edge (typical 1-face) and its neighbours.

Keywords: Voronoi tessellation; Poisson process; random measure; Palm distribution; typical face; typical edge

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1. Introduction

Let N be a stationary Poisson process on \mathbb{R}^d with finite intensity $\lambda > 0$. The *Voronoi cell*, $C(x)$, of $x \in N$ is the set of all points $y \in \mathbb{R}^d$ whose distances from x are smaller than or equal to their distances from all other points of N . The subject of this paper is the *Poisson–Voronoi tessellation* $\{C(x) : x \in N\}$. Voronoi tessellations can be defined for more general point processes than N . They are fundamental models in stochastic geometry and constitute arguably one of the most popular types of mathematical model in applications; see [7], [10], [11], and the references therein. We will present a complete and explicit description of the *Palm probability measure* P_k^0 , $k \in \{0, \dots, d\}$, which describes the statistical behaviour of N as seen from a *typical point* chosen ‘uniformly’ on the k -faces of the tessellation. In particular, we obtain the joint distribution of the $d - k + 1$ neighbours of the k -dimensional face containing this typical point. For $k = 0$, our result essentially boils down to the distribution of the typical cell of the *Poisson Delaunay tessellation*, established in [5] and [6]. In fact, our present approach owes much to these seminal papers.

To formulate our main result we introduce some notation. We are working in \mathbb{R}^d with Euclidean norm $|\cdot|$, scalar product $\langle \cdot, \cdot \rangle$, and unit sphere $S^{d-1} := \{z \in \mathbb{R}^d : |z| = 1\}$. The i -dimensional Hausdorff measure is denoted by \mathcal{H}^i . The closed ball with centre $a \in \mathbb{R}^d$ and radius $r \geq 0$ is denoted by $B(a, r)$, while $B^0(a, r)$ denotes the corresponding open ball. The volume $\mathcal{H}^d(B(0, 1))$ of the unit ball in \mathbb{R}^d is denoted by κ_d . For $k \in \{0, \dots, d\}$, let \mathcal{F}_k be the system of all k -faces of the Voronoi cells $C(x)$, $x \in N$ (see Subsection 2.2 for related notation).

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We define the following stationary random measure on \mathbb{R}^d :

$$M_k := \sum_{F \in \mathcal{F}_k} \mathcal{H}^k(F \cap \cdot). \tag{1.1}$$

It is well known [5], [6], [10] that the *intensity*, $\mu_k := E[M_k([0, 1]^d)]$, of M_k is given by

$$\begin{aligned} \mu_k &= \lambda^{(d-k)/d} \frac{2^{d-k+1} \pi^{(d-k)/2}}{d(d-k+1)!} \\ &\times \frac{\Gamma(d-k+k/d) \Gamma((d^2-dk+k+1)/2) \Gamma(1+d/2)^{d-k+k/d}}{\Gamma((k+1)/2) \Gamma((d^2-dk+k)/2) \Gamma((d+1)/2)^{d-k}}. \end{aligned} \tag{1.2}$$

Consider the Palm probability measure, $P_k^0 \equiv P_{M_k}^0$, of M_k (defined in Subsection 2.1). In the (rather trivial) case in which $k = d$, P_k^0 equals the underlying (stationary) probability measure. Under P_k^0 the origin $0 \in \mathbb{R}^d$ can be interpreted as a typical point of M_k . Almost surely with respect to P_k^0 , there are exactly $d - k + 1$ different points, $X_{k,0}, \dots, X_{k,d-k} \in N$, such that

$$R_k := |X_{k,0}| = \dots = |X_{k,d-k}|$$

and

$$N \cap B^0(0, R_k) = \emptyset.$$

These are the neighbours of the k -face containing 0 . We may assume that $X_{k,0}, \dots, X_{k,d-k}$ are in *general position*, i.e. not contained in some affine space of dimension $d - k - 1$. Hence, there exists a unique $(d - k)$ -dimensional ball in the affine hull of these points containing the points on its boundary. We let Z_k denote the centre of this ball. The case in which $d = 3$ and $k = 1$ is illustrated in Figure 1. For $k \leq d - 1$, the random variable

$$R'_k := |X_{k,0} - Z_k|$$

is P_k^0 -almost surely positive and we define

$$U_{k,0} := \frac{X_{k,0} - Z_k}{R'_k}, \quad \dots, \quad U_{k,d-k} := \frac{X_{k,d-k} - Z_k}{R'_k}.$$

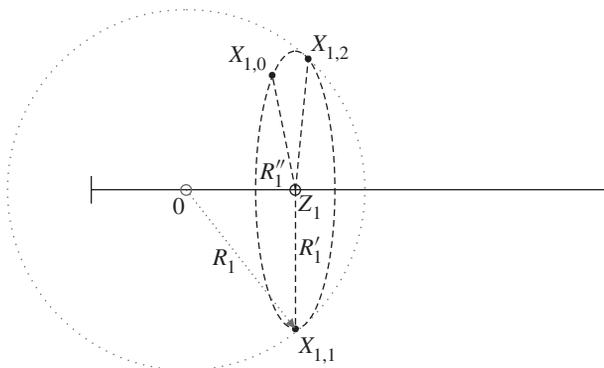


FIGURE 1: The situation under P_1^0 , for $d = 3$.

For $k = d$, we \mathbb{P}_k^0 -almost surely have $X_{k,0} = Z_k$ and we define $R'_k := 0$ and $U_{k,0} := 0$. Furthermore, for $k \geq 1$ the random variable

$$R'_k := |Z_k|$$

is \mathbb{P}_k^0 -almost surely positive and we define

$$U_k := \frac{Z_k}{R'_k}.$$

For $k = 0$, we have $Z_k = 0$ and we define $R'_k := 0$ and $U_k := 0$. From $|X_{k,0}| = \dots = |X_{k,d-k}|$ it easily follows that U_k is orthogonal to the affine hull of $U_{k,0}, \dots, U_{k,d-k}$. In particular,

$$R_k^2 = R_k'^2 + R_k''^2.$$

Under the Palm probability measure \mathbb{P}_k^0 , the random variables $R'_k, R''_k, U_{k,0}, \dots, U_{k,d-k}$, and U_k provide a natural description of the neighbours of the k -face containing 0:

$$N \cap B(0, R_k) = \{R''_k U_k + R'_k U_{k,i} : i = 0, \dots, d - k\} \quad \mathbb{P}_k^0\text{-almost surely.}$$

It is convenient to write

$$\Psi_k := \{U_{k,0}, \dots, U_{k,d-k}\}.$$

For $k \leq d - 1$, this is a finite point process on the unit sphere S^{d-1} .

The following theorem gives a complete and explicit description of the Palm probability measure, \mathbb{P}_k^0 , of M_k . We fix a $(d - k)$ -dimensional linear subspace $L \subset \mathbb{R}^d$ and denote by \mathbb{S}_L and \mathbb{S}_{L^\perp} the uniform distributions on the unit spheres in L and, respectively, in the orthogonal complement, L^\perp , of L . In the degenerate case, in which $L = \{0\}$, we let \mathbb{S}_L be the Dirac measure at 0.

Theorem 1.1. *Assume that N is a stationary Poisson process of intensity $\lambda > 0$ and let $k \in \{0, \dots, d\}$. Then the following assertions hold under \mathbb{P}_k^0 .*

- (i) *The random variables $(\{x \in N : |x| > R_k\}, R_k), R_k'^2/R_k^2$, and (Ψ_k, U_k) are independent.*
- (ii) *R_k^d is gamma distributed with shape parameter $d - k + k/d$ and scale parameter $\lambda \kappa_d$.*
- (iii) *The conditional distribution of $\{x \in N : |x| > R_k\}$ given $R_k = r$ can be chosen to be the distribution of a homogeneous Poisson process on the complement of the ball $B(0, r)$, with intensity λ .*
- (iv) *For $k \in \{1, \dots, d - 1\}$, $R_k'^2/R_k^2$ has a beta distribution with parameters $d(d - k)/2$ and $k/2$.*
- (v) *The random pair (Ψ_k, U_k) has distribution*

$$\begin{aligned} \mathbb{Q}_k := c_k^{-1} \int \dots \int 1_{\{(\vartheta u_0, \dots, \vartheta u_{d-k}, \vartheta u) \in \cdot\}} \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \\ \times \mathbb{S}_L(du_0) \dots \mathbb{S}_L(du_{d-k}) \mathbb{S}_{L^\perp}(du) \nu(d\vartheta), \end{aligned} \tag{1.3}$$

where $c_d := 1$, $c_{d-1} := 2^{d-1}$, and, for $k < d - 1$,

$$c_k := \frac{1}{((d - k)!)^{k+1}} \left[\frac{\Gamma((d - k)/2)}{\Gamma((d + 1)/2)} \right]^{d-k} \frac{\Gamma((d^2 - kd + k + 1)/2)}{\Gamma((d^2 - kd)/2)} \times \frac{\Gamma((k + 2)/2) \cdots \Gamma(d/2)}{\Gamma(\frac{1}{2}) \cdots \Gamma((d - k - 1)/2)};$$

$\Delta_{d-k}(u_0, \dots, u_{d-k})$ is the $(d - k)$ -dimensional volume of the simplex spanned by the vectors u_0, \dots, u_{d-k} (and is identically equal to 1 for $k = d$); and ν is the uniform distribution on the rotation group SO_d .

The proof of this theorem is given in Section 3 and is a straightforward application of well-established methods (see [5], [6], [7], and [10]). Despite the lack of a novel method, it is pleasing that the whole proof is the result of just one calculation. In the special case in which $k = 0$ (mentioned above), the integration with respect to ν can be dropped. For $k = d$ (i.e. $P_k^0 = P$), the result is not only well known but also almost trivial.

Based on Theorem 1.1 and a general relationship between Palm probabilities (see Proposition 2.1), in Section 4 we will discuss some distributional results for the typical k -face. In particular, we derive the joint distribution of the directions to the neighbours of the typical k -face. In Section 5 we will then provide complete and explicit formulae for the joint distribution of the typical edge and its neighbours. In particular, we obtain the distributional results derived and cited in [8] in a more transparent and rigorous way.

2. Poisson–Voronoi tessellations

2.1. Palm measures

Let \mathcal{N}' denote the set of all locally finite simple counting measures on \mathbb{R}^d , i.e. the set of all measures ω on \mathbb{R}^d that are integer valued on bounded sets and have $\omega(\{x\}) \leq 1$ for all $x \in \mathbb{R}^d$. Any $\omega \in \mathcal{N}'$ is identified with its support $\{x \in \mathbb{R}^d : \omega(\{x\}) > 0\}$, a locally finite subset of \mathbb{R}^d (i.e. by a point of ω we will mean a point in the support of ω). Actually, ω is just the sum of the Dirac measures δ_x over all x in the support of ω . We let \mathcal{N}' denote the σ -field generated by the mappings $\omega \mapsto \omega(B)$, $B \in \mathcal{B}^d$, where \mathcal{B}^d denotes the Borel σ -field on \mathbb{R}^d . Following [7] we say that the points of $\omega \in \mathcal{N}'$ are in *general quadratic position* if both no $d + 2$ points of ω lie on the boundary of some ball and any $k \in \{2, \dots, d + 1\}$ points $x_1, \dots, x_k \in \omega$ are in *general position*, i.e. do not lie in a $(k - 2)$ -dimensional affine subspace of \mathbb{R}^d . We denote by \mathcal{N} the measurable set of all $\omega \in \mathcal{N}'$ whose convex hull coincides with \mathbb{R}^d and whose points are in general quadratic position. Note that \mathcal{N} is shift invariant, in the sense that $\omega \in \mathcal{N}'$ belongs to \mathcal{N} if and only if $\omega + x := \{y + x : y \in \omega\} \in \mathcal{N}$ for all $x \in \mathbb{R}^d$. We write $\mathcal{N} := \mathcal{N}' \cap \mathcal{N}$ and denote by N the identity on \mathcal{N} .

We consider a probability measure P on $(\mathcal{N}, \mathcal{N})$ such that N is a homogeneous Poisson process of intensity $\lambda > 0$. This is justified by the well-known fact (see [7] and [10]) that almost all realizations of a stationary Poisson process (with positive intensity) are in \mathcal{N} . In this paper all random elements will be defined on the probability space $(\mathcal{N}, \mathcal{N}, P)$. Note that P is *stationary* in the sense that $P(N + x \in \cdot) = P$ for all $x \in \mathbb{R}^d$.

A *random measure* M on \mathbb{R}^d (see, e.g. [2]) is a random variable taking its values in the space, \mathcal{M} , of all locally bounded measures α on \mathbb{R}^d equipped with the σ -field \mathcal{M} generated by the mappings $\alpha \mapsto \alpha(B)$, $B \in \mathcal{B}^d$. Note that \mathcal{N}' is a measurable subset of \mathcal{M} . A *point process* on \mathbb{R}^d is a random measure M satisfying $P(M \in \mathcal{N}') = 1$. For a random measure M ,

it is convenient to write $M(\omega, B)$ instead of $M(\omega)(B)$. A random measure M on \mathbb{R}^d is said to be *stationary* if

$$M(\omega, B + x) = M(\omega - x, B), \quad \omega \in \mathcal{N}, x \in \mathbb{R}^d, B \in \mathcal{B}^d.$$

If M is a stationary random measure then the distribution of $M(\cdot + x)$ is the same for any $x \in \mathbb{R}^d$. The measure

$$P_M(A) := \iint 1_{\{\omega-x \in A: x \in [0,1]^d\}} M(\omega, dx) P(d\omega), \quad A \in \mathcal{N},$$

is called the *Palm measure* of M (with respect to P); see [3]. This measure is σ -finite and satisfies the *refined Campbell theorem*, i.e.

$$E \left[\int f(N - x, x) M(dx) \right] = E_M \left[\int f(N, x) dx \right]$$

for all measurable functions $f: \mathcal{N} \times \mathbb{R}^d \rightarrow [0, \infty)$, where E_M denotes expectation with respect to P_M and dx indicates integration with respect to Lebesgue measure. If the intensity, $\lambda_M := E[M([0, 1]^d)]$, of M is positive and finite, then we can define the Palm probability measure, $P_M^0 := \lambda_M^{-1} P_M$, of M . The expectation of a random variable ξ with respect to P_M^0 is denoted by $E_M^0[\xi]$.

2.2. Voronoi tessellations

Closely following [7] and Section 6.2 of [10], we now give a detailed description of Voronoi tessellations. We refer the reader with measurability questions to the latter monograph.

Let $\omega \in \mathcal{N}$. The Voronoi cell, $C(\omega, x)$, of $x \in \omega$ consists of all points $y \in \mathbb{R}^d$ satisfying $|y - x| \leq \min\{|y - z|: z \in \omega\}$. By $\mathcal{S}_d(\omega) := \{C(\omega, x): x \in \omega\}$ we then denote the Voronoi tessellation based on ω . The elements of $\mathcal{S}_d(\omega)$ are bounded and convex polytopes, and any bounded set is intersected by only finitely many cells in $\mathcal{S}_d(\omega)$ (see Satz 6.2.1 of [10]). The subject of this paper is $\mathcal{S}_d(N)$, the random tessellation based on the Poisson process N .

For $k \in \{0, \dots, d - 1\}$, a *k-face* (of $\mathcal{S}_d(\omega)$) is a nonempty intersection of $d - k + 1$ Voronoi cells (see also Satz 6.2.3 of [10]). The system of all such *k-faces* is denoted by $\mathcal{S}_k(\omega)$. A cell is referred to as a *d-face*. For $k \in \{0, \dots, d\}$ and $x \in \mathbb{R}^d$, we write $F_k(\omega, x) \equiv F \in \mathcal{S}_k(\omega)$ provided that x is in the relative interior of F . For all other x , we define $F_k(x) := \{x\}$ and note the following covariance property:

$$F_k(\omega, x) = F_k(\omega - x, 0) + x, \quad \omega \in \mathcal{N}, x \in \mathbb{R}^d.$$

The *k-faces* are generated in the following way. For points $x_0, \dots, x_{d-k} \in \mathbb{R}^d$ in general position, we let $z(x_0, \dots, x_{d-k})$ denote the centre of the uniquely determined $(d - k)$ -dimensional ball with x_0, \dots, x_{d-k} on its boundary. Furthermore, we let $F(x_0, \dots, x_{d-k})$ denote the *k-dimensional affine space* that is orthogonal to the above ball and contains $z(x_0, \dots, x_{d-k})$. Let $F \in \mathcal{S}_k(\omega)$. Then there exist (lexicographically ordered) points $x_0, \dots, x_{d-k} \in \omega$ such that

$$F = \{x \in F(x_0, \dots, x_{d-k}): B^0(x, |x - x_0|) \cap \omega = \emptyset\}. \tag{2.1}$$

Conversely, given different points $x_0, \dots, x_{d-k} \in \omega$ such that the thus-defined set F has a nonempty relative interior, we have $F \in \mathcal{F}_k(\omega)$.

Let $\omega \in N$, let x be a point in the relative interior of some $F \in \mathcal{F}_k(\omega)$, and choose $x_0, \dots, x_{d-k} \in \omega$ as in (2.1). By definition of N the set $\{x_0, \dots, x_{d-k}\} \subset \omega$ is uniquely determined by (2.1), and we define

$$\begin{aligned} R_k(\omega, x) &:= |x - x_0| = \dots = |x - x_{d-k}|, \\ X_{k,i}(\omega, x) &:= x_i, \quad i = 0, \dots, d - k, \\ Z_k(\omega, x) &:= z(x_0, \dots, x_{d-k}). \end{aligned}$$

For points $x \in \mathbb{R}^d$ that are not in the relative interior of some k -face, we set $R_k(\omega, x) := 0$ and $X_{k,0}(\omega, x) = \dots = X_{k,d-k}(\omega, x) = Z_k(\omega, x) := x$. For $k = 0$, we always have $Z_k(\omega, x) = x$. The mappings introduced above possess natural covariance and invariance properties with respect to shifts; e.g.

$$X_{k,i}(\omega, x) = X_{k,i}(\omega - x, 0) + x, \quad \omega \in N, x \in \mathbb{R}^d, \tag{2.2}$$

$$Z_k(\omega, x) = Z_k(\omega - x, 0) + x, \quad \omega \in N, x \in \mathbb{R}^d,$$

$$R_k(\omega, x) = R_k(\omega - x, 0), \quad \omega \in N, x \in \mathbb{R}^d. \tag{2.3}$$

2.3. Random measures associated with the Voronoi tessellation

Recall from (1.1) that, for any $k \in \{0, \dots, d\}$,

$$M_k = \sum_{F \in \mathcal{F}_k(N)} \mathcal{H}^k(F \cap \cdot)$$

is a stationary random measure, with Palm probability measure $P_k^0 \equiv P_{M_k}^0$.

We now construct another series of random measures using suitable centres of the k -faces. Let $k \in \{0, \dots, d\}$ and denote by \mathcal{P}_k the system of all k -dimensional, nonempty, bounded polytopes $F \subset \mathbb{R}^d$. We define a measurable mapping $\pi_k : N \times \mathcal{P}_k \rightarrow \mathbb{R}^d$ in the following way. For $\omega \in N$ and $F \in \mathcal{F}_k(\omega)$, we choose an arbitrary point y in the relative interior of F and define $\pi_k(\omega, F) := Z_k(\omega, y)$. In all other cases, we define $\pi_k(\omega, F)$ to be the centre of the smallest ball circumscribing $F \in \mathcal{P}_k$. From (2.2) we have the following covariance property:

$$\pi_k(\omega, F) = \pi_k(\omega - y, F - y) + y, \quad \omega \in N, F \in \mathcal{P}_k. \tag{2.4}$$

As the factorial moment measures of a Poisson process are absolutely continuous with respect to Lebesgue measure, it can easily be shown that the points $\pi_k(N, F)$, $F \in \mathcal{F}_k(N)$, are almost surely mutually different. Moreover, for P-almost every ω ,

$$\omega_k := \sum_{F \in \mathcal{F}_k(\omega)} \delta_{\pi_k(\omega, F)}$$

is an element of N' . For $\omega_k \in N'$ and $x \in \omega_k$, there exists a unique k -face, $C_k(\omega, x) \in \mathcal{F}_k(\omega)$, defined by the equivalence

$$C_k(\omega, x) = F \iff \pi_k(\omega, F) = x. \tag{2.5}$$

In all other cases, we define $C_k(\omega, x) := \{x\}$. For points $x \in \mathbb{R}^d$ in the relative interior of some $F \in \mathcal{F}_k(\omega)$, we define $\pi_k(\omega, x) := \pi_k(\omega, F_k(\omega, x))$. Otherwise, we let $\pi_k(\omega, x) := x$. From (2.4) we have

$$C_k(\omega, x) = C_k(\omega - x, 0) + x, \quad \omega \in N, x \in \mathbb{R}^d, \tag{2.6}$$

$$\pi_k(\omega, x) = \pi_k(\omega - x, 0) + x, \quad \omega \in N, x \in \mathbb{R}^d. \tag{2.7}$$

We often write $F_k(x) \equiv F_k(N, x)$, $C_k(x) \equiv C_k(N, x)$, $\pi_k(x) \equiv \pi_k(N, x)$, and $\pi_k(F) \equiv \pi_k(N, F)$.

We now define N_k , the point process of centres of k -faces, by $N_k(\omega) := \omega_k$ for $\omega_k \in N'$ and $N_k(\omega) := 0$ otherwise. By (2.7), N_k is stationary. Note that $N_d = N$. The distribution of the cell $C_k(0)$ under the Palm probability measure $\mathbb{P}_{N_k}^0$ is that of the *typical k -face* of the Voronoi tessellation. Actually, we interpret $\mathbb{P}_{N_k}^0$ itself as the distribution of N as seen from a typical k -face. As expected, the Palm measure of N_k is closely related to the Palm measure of M_k .

Proposition 2.1. *For all measurable functions $g : N \rightarrow [0, \infty)$, we have*

$$\mathbb{E}_{M_k}[g(N - \pi_k(0))] = \mathbb{E}_{N_k}[\mathcal{H}^k(C_k(0))g(N)], \tag{2.8}$$

$$\mathbb{E}_{N_k}[g(N)] = \mathbb{E}_{M_k}[\mathcal{H}^k(F_k(0))^{-1}g(N - \pi_k(0))]. \tag{2.9}$$

Proof. Both formulae are basically known. The first can be most easily proved with the help of the exchange formula of [9]. To prove the second we can apply the first with g replaced with $g\mathcal{H}^k(C_k(0))^{-1}$. Then (2.9) follows from

$$C_k(N - \pi_k(0), 0) = F_k(0) - \pi_k(0) \quad \mathbb{P}_{M_k}\text{-almost surely.}$$

By (2.8), the intensity, λ_k , of N_k is related to the intensity, μ_k , of M_k by the intuitively obvious formula

$$\mu_k = \lambda_k \mathbb{E}_{N_k}^0[\mathcal{H}^k(C_k(0))]. \tag{2.10}$$

Equation (2.9) tells us that the Palm measure \mathbb{P}_{N_k} is an *area-debiased* version of the Palm measure \mathbb{P}_{M_k} .

The next proposition is basically a well-known fact about general stationary face-to-face tessellations. For more details, we refer the reader to Theorem 5.1 of [6], which treats the case of shift-invariant functions.

Proposition 2.2. *For any measurable function $g : N \times N \rightarrow [0, \infty)$ and any $j, 0 \leq j < k \leq d$, we have*

$$\mathbb{E}_{N_j} \left[\sum_{F \in \mathcal{F}_k(N), C_j(0) \subset F} g(N, N - \pi_k(F)) \right] = \mathbb{E}_{N_k} \left[\sum_{G \in \mathcal{F}_j(N), G \subset C_k(0)} g(N - \pi_j(G), N) \right].$$

As $M_0 = N_0$ has a finite intensity, it follows from the version of this proposition with $j = 0$ and $g := 1$ that the intensity of N_k is finite for all $k \in \{0, \dots, d\}$. The values of this intensity

in the cases $d = 2$ and $d = 3$ can be found in, e.g. [10, p. 262]. The following special case will be important in Section 5.

Corollary 2.1. *For any measurable function $f : N \rightarrow [0, \infty)$, we have*

$$E_{N_0} \left[\sum_{F \in \mathcal{F}_1(N), 0 \in F} f(N - \pi_1(F)) \right] = 2 E_{N_1} [f(N)].$$

3. Proof of Theorem 1.1

We first need to introduce further notation. Let $k \in \{0, \dots, d\}$, let $\omega \in N$, and let x be a point in the relative interior of some $F \in \mathcal{F}_k(\omega)$. For $k \leq d - 1$, the number

$$R'_k(\omega, x) := |X_{k,0}(\omega, x) - Z_k(\omega, x)|$$

is positive, and we can thus define the unit vectors

$$U_{k,i}(\omega, x) := \frac{X_{k,i}(\omega, x) - Z_k(\omega, x)}{R'_k(\omega, x)}, \quad i = 0, \dots, d - k.$$

For $k = d$, we define $R'_d(\omega, x) := 0$ and $U_{d,0}(\omega, x) := 0$. Furthermore, for $k \geq 1$ we define

$$R''_k(\omega, x) := |x - Z_k(\omega, x)|$$

and, given that $R''_k(\omega, x) > 0$, the unit vector

$$U_k(\omega, x) := \frac{Z_k(\omega, x) - x}{R''_k(\omega, x)}.$$

In the exceptional case in which $R''_k(\omega, x) = 0$, we choose $U_k(\omega, x)$ to equal some fixed unit vector. For $k = 0$, we define $R''_0(\omega, x) := 0$ and $U_0(\omega, x) := 0$. For points $x \in \mathbb{R}^d$ that are not in the relative interior of some k -face, we let $R'_k(\omega, x) = R''_k(\omega, x) \equiv 0$ and choose $U_{k,0}(\omega, x), \dots, U_{k,d-k}(\omega, x), U_k(\omega, x)$ to be fixed unit vectors. Exceptions are the cases in which $k = 0$ and $k = d$, where we define $U_0(\omega, x) := 0$ and $U_{d,0}(\omega, x) := 0$, respectively. The mappings R'_k, R''_k , and $U_{k,i}$ have the same invariance property as R_k in (2.3).

We can now rewrite the random variables occurring in Theorem 1.1 as follows: $R_k \equiv R_k(N, 0)$, $R'_k \equiv R'_k(N, 0)$, $R''_k \equiv R''_k(N, 0)$, $U_{k,i} \equiv U_{k,i}(N, 0)$ for $i = 0, \dots, d - k$, and $U_k \equiv U_k(N, 0)$. For ease of exposition of the proof of the theorem, we now assume that $1 \leq k \leq d - 1$. The (well-known) case in which $k = d$ is easy to treat while the (somewhat easier) case in which $k = 0$ can be proved similarly. Our proof is similar to those of Theorem 7.2 of [6] and Satz 6.2.4 of [10]. Let $B := B(0, 1)$. For all measurable functions $h : N \rightarrow [0, \infty)$, we have

$$\begin{aligned} \kappa_d E_{M_k} [h(N)] &= E \left[\int_B h(N - y) M_k(dy) \right] \\ &= \frac{1}{(d - k + 1)!} E \left[\sum_{x_0, \dots, x_{d-k} \in N}^* \int h(N - y) 1_{\{y \in F(x_0, \dots, x_{d-k}) \cap B\}} \right. \\ &\quad \left. \times 1_{\{B^0(y, |y-x_0|) \cap N = \emptyset\}} \mathcal{H}^k(dy) \right], \end{aligned}$$

where ‘*’ indicates that the summation is over pairwise-different points of N . Using an iterated version of Mecke’s fundamental formula for Poisson processes (see Satz 3.1 of [3]) and letting

$$c := \lambda^{d-k+1} (\kappa_d(d-k+1)!)^{-1},$$

we obtain

$$E_{M_k}[h(N)] = c E \left[\int \cdots \int h((N \cup \{x_0, \dots, x_{d-k}\}) - y) \mathbf{1}_{\{y \in F(x_0, \dots, x_{d-k}) \cap B\}} \times \mathbf{1}_{\{N \cap B^0(y, |y-x_0|) = \emptyset\}} \mathcal{H}^k(dy) dx_0 \cdots dx_{d-k} \right].$$

For $\omega \in N, t \geq 0$, and $y \in \mathbb{R}^d$, we define

$$\omega^t := \omega \cap (\mathbb{R}^d \setminus B(0, t)) \quad \text{and} \quad \Psi(\omega, y) := \{U_0(\omega, y), \dots, U_{d-k}(\omega, y)\},$$

where on the right-hand side of the latter definition (and below) we omit the indices referring to the dimension of the faces. We apply the previous result to

$$h(\omega) := g_1(\omega^{R(\omega, 0)}) g_2(R(\omega, 0), R''(\omega, 0)) w_1(\Psi(\omega, 0)) w_2(U(\omega, 0)),$$

for suitable measurable, nonnegative functions g_1, g_2, w_1 , and w_2 . Using covariance and invariance properties as in (2.2) and (2.3), we obtain

$$E_{M_k}[h(N)] = c E \left[\int \cdots \int \mathbf{1}_{\{y \in F(x_0, \dots, x_{d-k}) \cap B\}} \mathbf{1}_{\{N \cap B^0(y, |y-x_0|) = \emptyset\}} \times g_1((N - y)^{R(N \cup \{x_0, \dots, x_{d-k}\}, y)}) \times g_2(R(N \cup \{x_0, \dots, x_{d-k}\}, y), R''(N \cup \{x_0, \dots, x_{d-k}\}, y)) \times w_1(\Psi(N \cup \{x_0, \dots, x_{d-k}\}, y)) \times w_2(U(N \cup \{x_0, \dots, x_{d-k}\}, y)) \mathcal{H}^k(dy) dx_0 \cdots dx_{d-k} \right].$$

Assume that x_0, \dots, x_{d-k} are in general position. Taking $y \in F(x_0, \dots, x_{d-k})$ and assuming that

$$N \cap B(y, |y - x_0|) = \emptyset,$$

we have

$$\begin{aligned} R(N \cup \{x_0, \dots, x_{d-k}\}, y) &= |x_0 - y|, \\ R''(N \cup \{x_0, \dots, x_{d-k}\}, y) &= |z(x_0, \dots, x_{d-k}) - y|, \\ \Psi(N \cup \{x_0, \dots, x_{d-k}\}, y) &= \tilde{\Psi}(x_0, \dots, x_{d-k}), \\ U(N \cup \{x_0, \dots, x_{d-k}\}, y) &= \tilde{U}(x_0, \dots, x_{d-k}, y), \end{aligned}$$

where

$$\begin{aligned} \tilde{\Psi}(x_0, \dots, x_{d-k}) &:= \left\{ \frac{x_0 - z(x_0, \dots, x_{d-k})}{|x_0 - z(x_0, \dots, x_{d-k})|}, \dots, \frac{x_{d-k} - z(x_0, \dots, x_{d-k})}{|x_{d-k} - z(x_0, \dots, x_{d-k})|} \right\}, \\ \tilde{U}(x_0, \dots, x_{d-k}, y) &:= \frac{z(x_0, \dots, x_{d-k}) - y}{|z(x_0, \dots, x_{d-k}) - y|}. \end{aligned}$$

(Here we take $\tilde{U}(x_0, \dots, x_{d-k}, y)$ to equal some fixed unit vector if $z(x_0, \dots, x_{d-k}) = y$. If x_0, \dots, x_{d-k} are not in general position, then all of these functions can be defined arbitrarily.) It follows that

$$\begin{aligned} E_{M_k}[h(N)] = c E & \left[\int \cdots \int 1_{\{y \in F(x_0, \dots, x_{d-k}) \cap B\}} 1_{\{N \cap B^0(y, |y-x_0|) = \emptyset\}} \right. \\ & \times g_1((N-y)^{|x_0-y|}) g_2(|x_0-y|, |z(x_0, \dots, x_{d-k})-y|) \\ & \times w_1(\tilde{\Psi}(x_0, \dots, x_{d-k})) \\ & \left. \times w_2(\tilde{U}(x_0, \dots, x_{d-k}, y)) \mathcal{H}^k(dy) dx_0 \cdots dx_{d-k} \right]. \end{aligned}$$

Since N is stationary and has independent increments, we obtain

$$\begin{aligned} E_{M_k}[h(N)] = c \int \cdots \int & 1_{\{y \in F(x_0, \dots, x_{d-k}) \cap B\}} \exp[-\lambda \kappa_d |y-x_0|^d] \\ & \times E[g_1(N^{|x_0-y|})] g_2(|x_0-y|, |z(x_0, \dots, x_{d-k})-y|) \\ & \times w_1(\tilde{\Psi}(x_0, \dots, x_{d-k})) \\ & \times w_2(\tilde{U}(x_0, \dots, x_{d-k}, y)) \mathcal{H}^k(dy) dx_0 \cdots dx_{d-k}. \end{aligned}$$

Following [10, pp. 260–261], we next apply the Blaschke–Petkantschin formula (see, e.g. Satz 7.2.1 of [10]) and Satz 7.2.2 of [10]. Recalling the definitions of the functions $\tilde{\Psi}$ and \tilde{U} , for a suitable constant $c' > 0$ we obtain

$$\begin{aligned} E_{M_k}[h(N)] = c' \int \cdots \int & 1_{\{y \in (z+\vartheta L^\perp) \cap B, z \in \vartheta(y_0+L), r \geq 0, y_0 \in L^\perp\}} \\ & \times E \left[g_1 \left(N \sqrt{|y-z|^2+r^2} \right) \right] \\ & \times g_2 \left(\sqrt{|y-z|^2+r^2}, |z-y| \right) w_1(\{u_0, \dots, u_{d-k}\}) \\ & \times w_2 \left(\frac{z-y}{|z-y|} \right) \exp \left[-\lambda \kappa_d \left(\sqrt{|y-z|^2+r^2} \right)^d \right] r^{d(d-k)-1} \\ & \times \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \\ & \times \mathcal{H}^k(dy) \mathbb{S}_{\vartheta L}(du_0) \cdots \mathbb{S}_{\vartheta L}(du_{d-k}) dr \mathcal{H}^{d-k}(dz) \mathcal{H}^k(dy_0) \nu(d\vartheta), \end{aligned}$$

where $L \subset \mathbb{R}^d$ is some fixed $(d-k)$ -dimensional linear subspace and $\mathbb{S}_{\vartheta L}$ is the uniform distribution on the unit sphere in the linear subspace ϑL . We take

$$g_2(a, b) := g_3(a) g_4((a^2 - b^2)/a^2), \quad 0 \leq b < a,$$

for measurable functions $g_3, g_4 : [0, \infty) \rightarrow [0, \infty)$. Noting that $\Delta_{d-k}(\cdot), |\cdot|$, and $\mathcal{H}^i(\cdot)$ are invariant under rotations, and replacing $(u_0, \dots, u_{d-k}, y, z)$ with $(\vartheta u_0, \dots, \vartheta u_{d-k}, \vartheta y, \vartheta z)$, we arrive at

$$\begin{aligned} E_{M_k}[h(N)] &= c' \int \cdots \int 1_{\{y \in (z+L^\perp) \cap B, z \in y_0+L, r \geq 0, y_0 \in L^\perp\}} \\ &\quad \times E \left[g_1 \left(N \sqrt{|y-z|^2+r^2} \right) \right] g_3 \left(\sqrt{|y-z|^2+r^2} \right) \\ &\quad \times g_4 \left(\frac{r^2}{|y-z|^2+r^2} \right) w_1(\{\vartheta u_0, \dots, \vartheta u_{d-k}\}) \\ &\quad \times w_2 \left(\frac{\vartheta(z-y)}{|z-y|} \right) \exp \left[-\lambda \kappa_d \left(\sqrt{|y-z|^2+r^2} \right)^d \right] r^{d(d-k)-1} \\ &\quad \times \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \\ &\quad \times \mathcal{H}^k(dy) \mathbb{S}_L(du_0) \cdots \mathbb{S}_L(du_{d-k}) dr \mathcal{H}^{d-k}(dz) \mathcal{H}^k(dy_0) \nu(d\vartheta) \\ &= c' \int \cdots \int 1_{\{x \in L^\perp, z-x \in B, z \in y_0+L, r \geq 0, y_0 \in L^\perp\}} E \left[g_1 \left(N \sqrt{|x|^2+r^2} \right) \right] \\ &\quad \times g_3 \left(\sqrt{|x|^2+r^2} \right) g_4 \left(\frac{r^2}{|x|^2+r^2} \right) w_1(\{\vartheta u_0, \dots, \vartheta u_{d-k}\}) \\ &\quad \times w_2 \left(\frac{\vartheta x}{|x|} \right) \exp \left[-\lambda \kappa_d \left(\sqrt{|x|^2+r^2} \right)^d \right] r^{d(d-k)-1} \\ &\quad \times \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \mathcal{H}^k(dx) \mathbb{S}_L(du_0) \cdots \mathbb{S}_L(du_{d-k}) \\ &\quad \times dr \mathcal{H}^{d-k}(dz) \mathcal{H}^k(dy_0) \nu(d\vartheta), \end{aligned}$$

where we have used the change of variable $y = z - x$ to obtain the second equality. For any fixed $x \in \mathbb{R}^d$, it can easily be shown that

$$\iint 1_{\{z-x \in B, z \in y_0+L, y_0 \in L^\perp\}} \mathcal{H}^{d-k}(dz) \mathcal{H}^k(dy_0) = \kappa_d.$$

Hence,

$$\begin{aligned} E_{M_k}[h(N)] &= \kappa_d c' \int \cdots \int 1_{\{x \in L^\perp, r \geq 0\}} E \left[g_1 \left(N \sqrt{|x|^2+r^2} \right) \right] g_3 \left(\sqrt{|x|^2+r^2} \right) \\ &\quad \times g_4 \left(\frac{r^2}{|x|^2+r^2} \right) w_1(\{\vartheta u_0, \dots, \vartheta u_{d-k}\}) \\ &\quad \times w_2 \left(\frac{\vartheta x}{|x|} \right) \exp \left[-\lambda \kappa_d \left(\sqrt{|x|^2+r^2} \right)^d \right] \\ &\quad \times r^{d(d-k)-1} \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \\ &\quad \times \mathcal{H}^k(dx) \mathbb{S}_L(du_0) \cdots \mathbb{S}_L(du_{d-k}) dr \nu(d\vartheta). \end{aligned}$$

Using polar coordinates in L^\perp gives

$$\begin{aligned} E_{M_k}[h(N)] &= c'' \int_0^\infty \int_0^\infty E\left[g_1\left(N\sqrt{s^2+r^2}\right)\right] g_3(\sqrt{s^2+r^2}) g_4\left(\frac{r^2}{r^2+s^2}\right) \\ &\quad \times \exp[-\lambda\kappa_d(\sqrt{s^2+r^2})^d] r^{d(d-k)-1} s^{k-1} dr ds \\ &\quad \times \int \cdots \int w_1(\{\vartheta u_0, \dots, \vartheta u_{d-k}\}) w_2(\vartheta u) \\ &\quad \times \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \mathbb{S}_L(du_0) \cdots \mathbb{S}_L(du_{d-k}) \mathbb{S}_{L^\perp}(du) \nu(d\vartheta), \end{aligned}$$

for a suitable constant c'' . It remains to treat the first double integral. Here, using the substitution $(s, r) = (v\sqrt{1-t}, v\sqrt{t})$, $t \in [0, 1]$, $v \in [0, \infty)$, with Jacobian $v/2\sqrt{t(1-t)}$, we find that the integral is proportional to

$$\int_0^\infty E[g_1(N^v)] g_3(v) v^{d(d-k)+k-1} e^{-\lambda\kappa_d v^d} dv \int_0^1 t^{d(d-k)/2-1} (1-t)^{k/2-1} g_4(t) dt.$$

The substitution $z = v^d$ completes the proof of the theorem.

4. Typical faces of a Poisson–Voronoi tessellation

In this section we deal with the distribution of the area of the typical k -face and its neighbours. Our basic general tool is Proposition 2.1. We first need to establish some results that hold under \mathbb{P}_k^0 .

Proposition 4.1. *Let $k \in \{1, \dots, d\}$. Then $(\mathcal{H}^k(F_k(0)), R'_k, R''_k)$ and Ψ_k are independent under the Palm probability measure \mathbb{P}_k^0 .*

Proof. For $\omega \in \mathcal{N}$, $r, r'' \geq 0$, and $u, v \in S^{d-1}$, we define

$$\rho(\omega, r, r'', u, v) := \sup\{t \geq 0: B^0(tv, \sqrt{r^2+t^2-2r''t\langle u, v \rangle}) \cap \omega^r = \emptyset\}.$$

Recall that $\omega^r := \omega \cap (\mathbb{R}^d \setminus B(0, r))$. An interpretation of this function will be given below.

In the following we will always assume that 0 is in the relative interior of some k -face. By definition, the linear hull of the k -face $F_k(0)$ is then given by

$$G_k := \text{Span}(U_{k,0}, \dots, U_{k,d-k})^\perp.$$

Take $v \in G_k \cap S^{d-1}$ and recall that $X_{k,0}$ is one of the neighbours of $F_k(0)$. Then

$$\begin{aligned} |X_{k,0} - tv|^2 &= |R'_k U_{k,0} + R''_k U_k - tv|^2 \\ &= R_k'^2 + |R''_k U_k - tv|^2 \\ &= R_k'^2 + R_k''^2 + t^2 - 2R_k'' t \langle U_k, v \rangle \\ &= R_k^2 + t^2 - 2R_k'' t \langle U_k, v \rangle. \end{aligned}$$

Hence, we have the equivalence

$$tv \in F_k(0) \iff \rho(N, R_k, R_k'', U_k, v) \geq t,$$

according to which $\rho(N, R_k, R''_k, U_k, \cdot)$ can be interpreted as the *radial function* of $F_k(0)$. It follows that

$$\mathcal{H}^k(F_k(0)) = A_{G_k}(N, R_k, R''_k, U_k),$$

where, for a k -dimensional linear subspace S ,

$$A_S(\omega, r, r'', u) := \frac{1}{k} \int_{S^{d-1} \cap S} \rho(\omega, r, r'', u, v)^k \mathcal{H}^{k-1}(dv).$$

We now apply Theorem 1.1, to obtain, for measurable, nonnegative functions f_1, f_2 , and f_3 with suitable domains,

$$\begin{aligned} & E_{M_k}^0[f_1(\mathcal{H}^k(F_k(0)))f_2(R_k, R''_k)f_3(\Psi_k)] \\ &= c_k^{-1} \int \cdots \int f_1(A_{\vartheta L^\perp}(\omega, r, r'', \vartheta u))f_2(r, r'')f_3(\{\vartheta u_0, \dots, \vartheta u_{d-k}\}) \\ &\quad \times \Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \\ &\quad \times \mathbb{S}_L(du_0) \cdots \mathbb{S}_L(du_{d-k})\mathbb{S}_{L^\perp}(du)v(d\vartheta)\Pi^r(d\omega)P_k^0((R, R'') \in d(r, r'')), \end{aligned} \tag{4.1}$$

where Π^r is the distribution of N^r under P . For any fixed $r, r'' \geq 0$, we have the invariance property $\rho(\vartheta\omega, r, r'', \vartheta u, \vartheta v) = \rho(\omega, r, r'', u, v)$ for $\omega \in N, u, v \in S^{d-1}$, and $\vartheta \in SO_d$. As Π^r is invariant under rotations, it follows that

$$\begin{aligned} & \iint f_1(A_{\vartheta L^\perp}(\omega, r, r'', \vartheta u))\mathbb{S}_{L^\perp}(du)\Pi^r(d\omega) \\ &= \iint f_1(A_{\vartheta L^\perp}(\vartheta\omega, r, r'', \vartheta u))\mathbb{S}_{L^\perp}(du)\Pi^r(d\omega) \\ &= \iint f_1(A_{L^\perp}(\omega, r, r'', u))\mathbb{S}_{L^\perp}(du)\Pi^r(d\omega) \\ &=: f_1^*(r, r''). \end{aligned}$$

Inserting this result into (4.1), we obtain

$$\begin{aligned} & E_{M_k}^0[f_1(\mathcal{H}^k(F_k(0)))f_2(R_k, R''_k)f_3(\Psi_k)] \\ &= c_k^{-1} \int \cdots \int f_1^*(r, r'')f_2(r, r'')f_3(\{\vartheta u_0, \dots, \vartheta u_{d-k}\})\Delta_{d-k}(u_0, \dots, u_{d-k})^{k+1} \\ &\quad \times \mathbb{S}_L(du_0) \cdots \mathbb{S}_L(du_{d-k})v(d\vartheta)P_k^0((R, R'') \in d(r, r'')) \\ &= E_{M_k}^0[f_1^*(R_k, R''_k)f_2(R_k, R''_k)f_3(\Psi_k)] \\ &= E_{M_k}^0[f_1^*(R_k, R''_k)f_2(R_k, R''_k)]E_{M_k}^0[f_3(\Psi_k)], \end{aligned}$$

where we have again used Theorem 1.1. Using this and specializing to the case in which $f_3 := 1$ yields the assertion.

The preceding proof yields more detailed, though not very explicit, information on the distribution of $(\mathcal{H}^k(F_k(0)), R_k, R''_k)$ under P_k^0 .

Corollary 4.1. *For all $k \in \{1, \dots, d\}$, we P_k^0 -almost surely have*

$$P_k^0(\mathcal{H}^k(F_k(0)) \in \cdot \mid R_k, R'_k) = \iint 1_{\{A_{L^\perp}(\omega, R_k, R'_k, u) \in \cdot\}} \mathbb{S}_{L^\perp}(du) \Pi^{R_k}(d\omega).$$

For $x \in N_k$, we let $\rho_k(x) \equiv p_k(N, x)$ denote the distance of x from N and $V_{k,0}(x), \dots, V_{k,d-k}(x)$ the unit vectors such that $\rho_k(x)V_{k,0}(x), \dots, \rho_k(x)V_{k,d-k}(x)$ are the neighbours of the k -face $C_k(x)$ (see (2.5)). We let $\Phi_k(x) \equiv \Phi_k(N, x) := \{V_{k,0}(x), \dots, V_{k,d-k}(x)\}$. For $x \notin N_k$, we give $\rho_k(x)$ and $\Phi_k(x)$ some fixed values.

Theorem 4.1. *For any $k \in \{1, \dots, d\}$, the following assertions hold under $P_{N_k}^0$.*

- (i) *The random variables $(\mathcal{H}^k(C_k(0)), \rho_k(0))$ and $\Phi_k(0)$ are independent.*
- (ii) *The distribution of $\Phi_k(0)$ is $\mathbb{Q}_k(\cdot \times S^{d-1})$, where \mathbb{Q}_k is as given in (1.3).*

Proof. We will use Proposition 4.1 together with (2.9). Assume that 0 is in the relative interior of a k -face. Then $\Phi_k(N - \pi_k(0), 0) = \Psi_k$, $\rho_k(N - \pi_k(0), 0) = R'_k$, and $C_k(N - \pi_k(0), 0) = F_k(0) - \pi_k(0)$. Applying (2.9) to measurable functions h_1 and h_2 (with suitable domains), we obtain

$$\begin{aligned} \lambda_k E_{N_k}^0 [h_1(\mathcal{H}^k(C_k(0)), \rho_k(0))h_2(\Phi_k(0))] &= \mu_k E_{M_k}^0 [\mathcal{H}^k(F_k(N, 0))^{-1}h_1(\mathcal{H}^k(F_k(0)), R'_k)h_2(\Psi_k)] \\ &= \mu_k E_{M_k}^0 [\mathcal{H}^k(F_k(N, 0))^{-1}h_1(\mathcal{H}^k(F_k(0)), R'_k)] E_{M_k}^0 [h_2(\Psi_k)] \\ &= \lambda_k E_{N_k}^0 [h_1(\mathcal{H}^k(C_k(0)), \rho_k(0))] E_{M_k}^0 [h_2(\Psi_k)], \end{aligned}$$

where we have used Proposition 4.1 to obtain the second equality. Specializing to the case in which $h_1 := 1$ yields (i) and, in view of Theorem 1.1(v), also (ii).

The above proof and Corollary 4.1 yield the following result.

Corollary 4.2. *For all $k \in \{1, \dots, d\}$, we have*

$$\begin{aligned} P_{N_k}^0((\mathcal{H}^k(C_k(0)), \rho_k(0)) \in \cdot) &= \frac{\mu_k}{\lambda_k} \iiint A_{L^\perp}(\omega, r, r'', u)^{-1} 1_{\{(A_{L^\perp}(\omega, r, r'', u), \sqrt{r^2 - r''^2}) \in \cdot\}} \\ &\quad \times \mathbb{S}_{L^\perp}(du) \Pi^r(d\omega) P_k^0((R_k, R'_k) \in d(r, r'')), \end{aligned}$$

where the distribution $P_k^0((R_k, R'_k) \in \cdot)$ is as given in Theorem 1.1.

Finally in this section, we establish a more detailed version of Proposition 4.1 that holds for $k = 1$. For a point x in the relative interior of some edge $F \in \mathfrak{F}_1(N)$, we define $U_1^*(x)$ as the unique vector in $\{-U_1(N, x), U_1(N, x)\}$ such that $-U_1^*(x)$ is lexicographically smaller than $U_1^*(x)$. Furthermore, we let $I_1(x)$ denote the $\{-1, 1\}$ -valued random variable satisfying

$$U_1(N, x) = I_1(x)U_1^*(x)$$

and $T'(x)$ and $T''(x)$ the nonnegative random variables satisfying

$$F_1(x) = [x - T'(x)U_1^*(x), x + T''(x)U_1^*(x)].$$

Note that $T'(x) + T''(x) = \mathcal{H}^1(F_1(x))$ is the length of the edge $F_1(x)$. For all points x that are not in the relative interior of some edge, we give $U_1^*(x)$, $I_1(x)$, $T'(x)$, and $T''(x)$ some fixed

values. Similarly to above, we use the abbreviations $T' \equiv T'(0)$, $T'' \equiv T''(0)$, $I_1 \equiv I_1(0)$, and $U_1^* \equiv U_1^*(0)$.

Together with Theorem 1.1, the following proposition provides a complete description of the joint distribution of the edge $[-T'U_1^*, T''U_1^*]$ and its neighbours under P_1^0 . We denote by $V(a, s, t)$ the d -dimensional volume of the union of two balls with radii s and t whose centres are separated by a distance a .

Proposition 4.2. *Under the Palm probability measure P_1^0 , the random variables Ψ_1 and $(T', T'', I_1, R_1, R'_1)$ are independent and the random variables T' and T'' are conditionally independent given (I_1, R_1, R'_1) . Furthermore, for all $t_1, t_2 > 0$, we almost surely have*

$$P_1^0(T' > t_1 \mid I_1, R_1, R'_1) = e^{\lambda \kappa_d R_1^d} \exp\left[-\lambda V\left(t_1, R_1, \sqrt{R_1^2 + 2I_1 R_1'' t_1 + t_1^2}\right)\right],$$

$$P_1^0(T'' > t_2 \mid I_1, R_1, R'_1) = e^{\lambda \kappa_d R_1^d} \exp\left[-\lambda V\left(t_2, R_1, \sqrt{R_1^2 - 2I_1 R_1'' t_2 + t_2^2}\right)\right].$$

Proof. P_1^0 -almost surely, the point 0 is in the relative interior of some $F \in \mathcal{F}_1(N)$ and

$$\{T' > t_1\} = \left\{N \cap B\left(-t_1 U_1^*, \sqrt{R_1^2 + (t_1 + I_1 R_1'')^2}\right) \cap (\mathbb{R}^d \setminus B(0, R_1)) = \emptyset\right\},$$

$$\{T'' > t_2\} = \left\{N \cap B\left(t_2 U_1^*, \sqrt{R_1^2 + (t_2 - I_1 R_1'')^2}\right) \cap (\mathbb{R}^d \setminus B(0, R_1)) = \emptyset\right\}.$$

It is easy to check that

$$B\left(-t_1 U_1^*, \sqrt{R_1^2 + (t_1 + I_1 R_1'')^2}\right) \cap B\left(t_2 U_1^*, \sqrt{R_1^2 + (t_2 - I_1 R_1'')^2}\right) \cap (\mathbb{R}^d \setminus B(0, R_1)) = \emptyset.$$

Assertions (i) and (iii) of Theorem 1.1 imply that

$$P(N \setminus B(0, R_1) \in \cdot \mid R_1, R'_1, \Psi_1, I_1) = P(N \setminus B(0, R_1) \in \cdot \mid R_1) \quad P_1^0\text{-almost surely.}$$

Hence, T' and T'' are conditionally independent given (Ψ_1, I_1, R_1, R'_1) . The theorem also implies that

$$P_1^0(T' > t_1 \mid R_1, R'_1, \Psi_1, I_1)$$

$$= \exp\left[-\lambda \mathcal{H}^d\left(B\left(-t_1 U_1^*, \sqrt{R_1^2 + 2I_1 R_1'' t_1 + t_1^2}\right) \setminus B(0, R_1)\right)\right]$$

$$= e^{\lambda \kappa_d R_1^d} \exp\left[-\lambda V\left(t_1, R_1, \sqrt{R_1^2 + 2I_1 R_1'' t_1 + t_1^2}\right)\right],$$

where we have used the fact that $R_1^2 + R_1''^2 = R_1^2$. A similar formula holds for the conditional distribution of T'' . In particular, we obtain the asserted formulae. Given Ψ_1 , there are almost surely only two possible values for U_1 . Hence, it follows easily from Theorem 1.1(v) that Ψ_1 and I_1 are independent. Using the fact that the above conditional probabilities depend only on (I_1, R_1, R_1'') , we obtain the remaining assertions using Theorem 1.1(i).

5. The typical edge and its neighbours

The topic of this, final, section is the Palm probability measure $P_{N_1}^0$. It turns out that under this measure Theorem 4.1 (and Corollary 4.2) can be considerably improved.

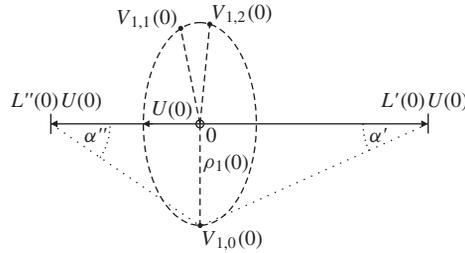


FIGURE 2: The situation under $\mathbb{P}_{N_1}^0$, for $d = 3$.

First we introduce a convenient description of the typical edge $C_1(0)$ (see (2.5) for notation). For $x \in N_1$, we let $U(x)$ denote the unit vector orthogonal to the affine hull of $\Phi_1(x)$ such that $-U(x)$ is lexicographically smaller than $U(x)$. Let $L'(x)$ and $L''(x)$, $L'(x) \leq L''(x)$, denote random variables such that

$$C_1(x) = [L'(x)U(x), L''(x)U(x)].$$

For $x \notin N_1$, we give $U(x)$ some fixed value and define $L'(x) = L''(x) := 0$. The case in which $d = 3$ and $k = 1$ is illustrated in Figure 2 (using the notation introduced just before Theorem 4.1).

By (2.10), we have $\mu_1 = \lambda_1 E_{N_1}^0 [L''(0) - L'(0)]$. Since $2\lambda_1 = (d + 1)\lambda_0$ (to see this, set $f := 1$ in Corollary 2.1) and $\lambda_0 = \mu_0$, we obtain

$$E_{N_1}^0 [L''(0) - L'(0)] = \frac{\mu_1}{\lambda_1} = \frac{2\mu_1}{(d + 1)\lambda_0} = \frac{2\mu_1}{(d + 1)\mu_0}.$$

Hence, (1.2) yields a formula for the mean length, $E_{N_1}^0 [L''(0) - L'(0)]$, of the typical edge; see also [8, p. 285].

We now introduce the notation

$$V^*(t, r, w) := \frac{d}{dt} V(t, r, \sqrt{t^2 + r^2 - 2trw}), \quad t, r \geq 0, |w| \leq 1. \tag{5.1}$$

More details on this function can be found in [8] and below. We define

$$J(u_0, u_1, \dots, u_d) := 1 - 2 \times \mathbb{1}_{\{z(u_1, \dots, u_d), u_1 - u_0 \leq 0\}}, \quad u_0, \dots, u_d \in S^{d-1},$$

where $z(u_1, \dots, u_d) := 0$ for points u_1, \dots, u_d that are not in general position. To interpret this function we consider the Voronoi tessellation generated by points $u_0, \dots, u_d \in S^{d-1}$ in general position. We then have $J(u_0, \dots, u_d) = 1$ if and only if $z(u_0, \dots, u_d)$ is pointing along the edge generated by u_1, \dots, u_d and starting at the point 0. We define a probability measure on the interval $[-1, 1]$ by

$$\mathbb{W} := c_0^{-1} \int \cdots \int \mathbb{1}_{\{J(u_0, \dots, u_d) | z(u_1, \dots, u_d) \in \cdot\}} \Delta_d(u_0, \dots, u_d) \mathbb{S}(du_0) \cdots \mathbb{S}(du_d), \tag{5.2}$$

where c_0 is as defined in Theorem 1.1(v) and \mathbb{S} is the uniform distribution on the unit sphere S^{d-1} . Note that $\mathbb{W}(\{-1\}) = \mathbb{W}(\{1\}) = 0$. In view of Theorem 1.1(v) and Corollary 2.1, we can interpret \mathbb{W} as the conditional distribution of the signed distance of the typical vertex

from the centre of a randomly chosen edge emanating from the vertex, given that the distance from the typical vertex to its neighbours is 1. Another interpretation of \mathbb{W} and a more explicit formula will be given in Proposition 5.2.

The following proposition generalizes a result of [8] and provides the joint distribution of the typical edge and its neighbours. The result can be considered a preliminary one: a more explicit version will be given in Theorem 5.1. As above, it is convenient to introduce the shorthand notation $(L', L'', \rho_1, \Phi_1) \equiv (L'(0), L''(0), \rho_1(0), \Phi_1(0))$, where we recall from Section 4 that $\Phi_1(x)$ is the set of the normalized directions to the d neighbours of the edge with centre $x \in N_1$.

Proposition 5.1. *The following assertions hold under $P_{N_1}^0$.*

- (i) *The random variables (L', L'', ρ_1) and Φ_1 are independent.*
- (ii) *The distribution of Φ_1 is $\mathbb{Q}_1(\cdot \times S^{d-1})$, where \mathbb{Q}_1 is as given in (1.3) for $k = 1$.*
- (iii) *For any measurable mappings $h_1: \mathbb{R}^2 \rightarrow [0, \infty)$ and $h_2: [0, \infty) \rightarrow [0, \infty)$,*

$$\begin{aligned}
 & E_{N_1}^0[h_1(L', L'')h_2(\rho_1)] \\
 &= \frac{d\lambda(\lambda\kappa_d)^d}{2\Gamma(d)} \int_{-1}^1 \int_0^\infty \int_0^\infty (h_1(rw - t, rw) + h_1(-rw, -rw + t)) \\
 &\quad \times h_2(r\sqrt{1 - w^2})V^*(t, r, w) \\
 &\quad \times \exp[-\lambda V(t, r, \sqrt{t^2 + r^2 - 2trw})] \\
 &\quad \times r^{d^2-1} dt dr \mathbb{W}(dw), \tag{5.3}
 \end{aligned}$$

where $V^*(t, r, w)$ is as defined in (5.1) and \mathbb{W} is as given in (5.2).

Proof. Using Proposition 4.2 instead of Proposition 4.1, the first two assertions can be proved as were the equivalent assertions of Theorem 4.1. To determine the distribution of (L', L'', ρ_1) under $P_{N_1}^0$, we will now use Corollary 2.1 and an idea that goes back to [4] (see also [8]). Assume that $x_0, \dots, x_d \in \mathbb{R}^d$ are in general position. Then, for any $i \in \{0, \dots, d\}$, define

$$z_i(x_0, \dots, x_d) := z(x_0, \dots, x_{i-1}, x_{i+1}, \dots, x_d)$$

and note that

$$z_i(ax_0, \dots, ax_d) = az_i(x_0, \dots, x_d), \quad a \neq 0.$$

For points $x_0, \dots, x_d \in \mathbb{R}^d$ that are not in general position, we define $z_i(x_0, \dots, x_d) := 0$. For a measurable function $h: N \rightarrow [0, \infty)$, using Corollary 2.1 we obtain

$$2\lambda_1 E_{N_1}^0[h] = \lambda_0 \sum_{i=0}^d E_{N_0}^0[h(N - Rz_i(U_0, \dots, U_d))], \tag{5.4}$$

where $R \equiv R_0$ and $(U_0, \dots, U_d) \equiv (U_{0,0}, \dots, U_{0,d})$.

Next we introduce some random variables as seen from the typical vertex. From $i \in \{0, \dots, d\}$, we define

$$\begin{aligned} J_i &:= J(U_i, U_0, \dots, U_{i-1}, U_{i+1}, \dots, U_d), \\ Z_i &:= z_i(U_0, \dots, U_d), \quad \eta_i := |Z_i|^{-1} Z_i, \\ L_i &:= L''(N - RZ_i, 0) - L'(N - RZ_i, 0) = L''(N, RZ_i) - L'(N, RZ_i). \end{aligned}$$

We also define $U_{(i)} := U_{i+1}$ for $i \leq d - 1$ and $U_{(d)} := U_0$. Let $t > 0$. A crucial fact is the equivalence

$$L_i > t \iff N \cap [B(tJ_i\eta_i, |RU_{(i)} - tJ_i\eta_i|) \setminus B(0, R)] = \emptyset.$$

Assertions (i) and (iii) of Theorem 1.1 imply that

$$P_{N_0}^0(L_i > t \mid \Psi_0, R) = \exp\left[-\lambda V\left(t, R, \sqrt{R^2 + t^2 - 2tRJ_i\langle U_{(i)}, \eta_i \rangle}\right)\right] e^{\lambda\kappa_d R^d},$$

i.e.

$$P_{N_0}^0(L_i \in dt \mid \Psi_0, R) = F(J_i\langle U_{(i)}, \eta_i \rangle, t, R) dt, \tag{5.5}$$

where

$$\begin{aligned} F(w, t, r) &:= -e^{\lambda\kappa_d r^d} \frac{d}{dt} \exp[-\lambda V(t, r, \sqrt{t^2 + r^2 - 2trw})] \\ &= \lambda e^{\lambda\kappa_d r^d} V^*(t, r, w) \exp[-\lambda V(t, r, \sqrt{t^2 + r^2 - 2trw})]. \end{aligned}$$

Under $P_{N_1}^0$ we consider the set, $E := \{L'U(0), L''U(0)\}$, of vertices of the typical edge. From (5.4) and (5.5), for any nonnegative, measurable function g with suitable domain we then obtain

$$\begin{aligned} &2\lambda_1 E_{N_1}^0[g(E, \rho_1)] \\ &= \lambda_0 \sum_{i=0}^d E_{N_0}^0[g(\{-RZ_i, -RZ_i + L_i J_i \eta_i\}, R\sqrt{1 - |Z_i|^2})] \\ &= \lambda_0 \sum_{i=0}^d E_{N_0}^0\left[\int g(\{-RZ_i, -RZ_i + tJ_i \eta_i\}, R\sqrt{1 - |Z_i|^2}) F(J_i\langle U_{(i)}, \eta_i \rangle, t, R) dt\right]. \end{aligned}$$

It is now easy to check that

$$\langle U_{(i)}, z_i(U_0, \dots, U_d) \rangle = |z_i(U_0, \dots, U_d)|^2 = |Z_i|^2, \quad i = 0, \dots, d.$$

Hence, we have $\langle U_{(i)}, \eta_i \rangle = |Z_i|$. Letting $\xi_i := J_i|Z_i|$, and noting that $Z_i = |Z_i|\eta_i = J_i\xi_i\eta_i$, we can rewrite our previous result as

$$\begin{aligned} &2\lambda_1 E_{N_1}^0[g(E, \rho_1)] \\ &= \lambda_0 \sum_{i=0}^d E_{N_0}^0\left[\int g(\{-R\xi_i J_i \eta_i, -R\xi_i J_i \eta_i + tJ_i \eta_i\}, R\sqrt{1 - \xi_i^2}) F(\xi_i, t, R) dt\right]. \end{aligned}$$

Let A_i denote the event that $-J_i Z_i < J_i Z_i$, and let A_i^c be its complement. By definition of (L', L'') , for any measurable function $h: \mathbb{R}^2 \times [0, \infty) \rightarrow [0, \infty)$ we obtain

$$\begin{aligned}
 & 2\lambda_1 E_{N_1}^0 [h(L', L'', \rho_1)] \\
 &= \lambda_0 \sum_{i=0}^d E_{N_0}^0 \left[1_{A_i} \int h\left(-R\xi_i, -R\xi_i + t, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right] \\
 &+ \lambda_0 \sum_{i=0}^d E_{N_0}^0 \left[1_{A_i^c} \int h\left(R\xi_i - t, R\xi_i, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right]. \tag{5.6}
 \end{aligned}$$

To conclude the proof of assertion (iii), we apply a symmetry argument based on \mathbb{Q}_0 , the distribution of $\{U_0, \dots, U_d\}$ (see (1.3)) and the independence of (U_0, \dots, U_d) and R . For all $u_0, \dots, u_d \in S^{d-1}$, we have

$$z_i(-u_0, \dots, -u_d) = -z_i(u_0, \dots, u_d)$$

and, hence,

$$J_i(-u_0, \dots, -u_d) = J_i(u_0, \dots, u_d).$$

Furthermore, we have $\Delta_d(-u_0, \dots, -u_d) = \Delta_d(u_0, \dots, u_d)$. It follows that

$$\begin{aligned}
 & \sum_{i=0}^d E_{N_0}^0 \left[1_{A_i} \int h\left(-R\xi_i, -R\xi_i + t, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right] \\
 &= \sum_{i=0}^d E_{N_0}^0 \left[1_{A_i^c} \int h\left(-R\xi_i, -R\xi_i + t, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right] \\
 &= \frac{1}{2} \sum_{i=0}^d E_{N_0}^0 \left[\int h\left(-R\xi_i, -R\xi_i + t, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right].
 \end{aligned}$$

This and a similar formula for the second summand on the right-hand side of (5.6) yields

$$\begin{aligned}
 & 2\lambda_1 E_{N_1}^0 [h(L', L'', \rho_1)] \\
 &= \frac{\lambda_0}{2} \sum_{i=0}^d E_{N_0}^0 \left[\int h\left(-R\xi_i, -R\xi_i + t, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right] \\
 &+ \frac{\lambda_0}{2} \sum_{i=0}^d E_{N_0}^0 \left[\int h\left(R\xi_i - t, R\xi_i, R\sqrt{1 - \xi_i^2}\right) F(\xi_i, t, R) dt \right].
 \end{aligned}$$

By definition of \mathbb{W} (see (5.2)) and Theorem 1.1(v), we have

$$\mathbb{W} = \frac{1}{d + 1} \sum_{i=0}^d P_{N_0}^0 (\xi_i \in \cdot). \tag{5.7}$$

Again, using the independence of R and (ξ_0, \dots, ξ_d) , this implies that

$$\begin{aligned} &2\lambda_1 E_{N_1}^0[h(L', L'', \rho_1)] \\ &= \frac{(d+1)\lambda_0}{2} E_{N_0}^0 \left[\iint h(-Rw, -Rw+t, R\sqrt{1-w^2})F(w, t, R) dt \mathbb{W}(dw) \right] \\ &+ \frac{(d+1)\lambda_0}{2} E_{N_0}^0 \left[\iint h(Rw-t, Rw, R\sqrt{1-w^2})F(w, t, R) dt \mathbb{W}(dw) \right]. \end{aligned}$$

Recalling that $2\lambda_1 = (d+1)\lambda_0$ and that R^d has a gamma distribution with shape parameter d and scale parameter $\lambda\kappa_d$, we obtain assertion (iii).

Assertion (iii) of the previous theorem implies the following natural symmetry property:

$$P_{N_1}^0((L', L'', \rho_1) \in \cdot) = P_{N_1}^0((-L'', -L', \rho_1) \in \cdot). \tag{5.8}$$

To derive an alternative to (5.3), we introduce the $[-1, 1]$ -valued random variables

$$\eta' := -\frac{L'}{\sqrt{L'^2 + \rho_1^2}}, \quad \eta'' := \frac{L''}{\sqrt{L''^2 + \rho_1^2}}.$$

We have $\eta' = \cos \alpha'$ and $\eta'' = \cos \alpha''$ for some random angles $\alpha', \alpha'' \in [0, \pi]$. Here α' is the angle spanned by $U(0)$ and the vector between the edge $L'U(0)$ and one of the neighbours of $C_1(0)$, while α'' is the angle spanned by $-U(0)$ and the vector between the edge $L''U(0)$ and one of the neighbours of $C_1(0)$ (see Figure 2). Since $\sin(\alpha' + \alpha'') \geq 0$, it follows that $q(\eta', \eta'') \geq 0$, where

$$q(w, v) := v\sqrt{1-w^2} + w\sqrt{1-v^2}. \tag{5.9}$$

This inequality also follows more directly from $L'' - L' \geq 0$. Under $P_{N_1}^0$ there is (almost surely) a one-to-one correspondence between (L', L'', ρ_1) and (η', η'', L) .

The next two propositions can be found in [8]. For any $s \in \mathbb{R}$, we denote by $[s]$ the largest integer k such that $k \leq s$.

Proposition 5.2. *We have*

$$P_{N_1}^0(\eta' \in \cdot) = P_{N_1}^0(\eta'' \in \cdot) = \mathbb{W}. \tag{5.10}$$

Moreover, the measure \mathbb{W} is given by

$$\mathbb{W}(dw) = c^*(1-w^2)^{(d^2-d-2)/2}b(w) dw, \tag{5.11}$$

where

$$c^* = \frac{(d^2-d)\Gamma((d+1)/2)\Gamma(d^2/2)}{\Gamma(d/2)\Gamma((d^2+1)/2)}, \quad b(w) := \sum_{i=0}^{[(d-1)/2]} b_i(w). \tag{5.12}$$

The $b_i(w)$ are given by

$$b_0(w) := \begin{cases} \pi^{-1}((\pi - \arccos(w))w + \sqrt{1-w^2}) & \text{if } d \text{ is even,} \\ (1+w)/2 & \text{if } d \text{ is odd,} \end{cases}$$

and, for $i \in \{1, \dots, [(d - 1)/2]\}$, by

$$b_i(w) := \begin{cases} -\frac{1}{4\sqrt{\pi}} \frac{\Gamma(i)}{\Gamma(i + \frac{3}{2})} (1 - w^2)^{i+1/2} & \text{if } d \text{ is even,} \\ -\frac{1}{4\sqrt{\pi}} \frac{\Gamma(i - \frac{1}{2})}{\Gamma(i + 1)} (1 - w^2)^i & \text{if } d \text{ is odd.} \end{cases}$$

Proof. Here we prove only the first assertion. The second will be derived in the proof of Proposition 5.3. We choose $\omega, \omega' \in N$ such that $0 \in N_0(\omega)$ and $\omega' = \omega - x$ for some (unique) $x \in N_1(\omega)$ satisfying $0 \in C_1(\omega, x)$. We then define $\alpha(\omega, \omega') \in [0, \pi]$ as the angle spanned by the vector between 0 and one of the neighbours of $C_1(\omega, x)$ and the (directed) edge $C_1(\omega, x)$ starting at 0. For all other $\omega, \omega' \in N$, we give $\alpha(\omega, \omega')$ some fixed value in $[0, \pi]$. For $j = 0$ and $k = 1$, Proposition 2.2 yields, for any measurable function $f: [-1, 1] \rightarrow [0, \infty)$,

$$\lambda_0 E_{N_0}^0 \left[\sum_{x \in N_1, 0 \in C_1(x)} f(\cos \alpha(N, N - x)) \right] = \lambda_1 E_{N_1}^0 \left[\sum_{x \in N_0, x \in C_1(0)} f(\cos \alpha(N - x, N)) \right]. \tag{5.13}$$

Using the notation established in the proof of Proposition 5.1, the left-hand side of this equation can be written as

$$\lambda_0 E_{N_0}^0 \left[\sum_{i=0}^d f(\xi_i) \right] = (d + 1)\lambda_0 \int f(w) \mathbb{W}(dw);$$

see (5.7). However, the right-hand side of (5.13) equals

$$\lambda_1 E_{N_1}^0 [f(\eta') + f(\eta'')].$$

Now, (5.8) implies that η' and η'' have the same distribution under $P_{N_1}^0$. Therefore,

$$(d + 1)\lambda_0 \int f(w) \mathbb{W}(dw) = 2\lambda_1 E_{N_1}^0 [f(\eta')].$$

As $(d + 1)\lambda_0 = 2\lambda_1$, the assertion in (5.10) follows.

We define $L := L'' - L'$ as the length of the typical edge.

Proposition 5.3. *The joint distribution of η', η'' , and L under $P_{N_1}^0$ is given by*

$$\begin{aligned} &P_{N_1}^0((\eta', \eta'', L) \in \cdot) \\ &= \frac{d^2(\lambda\kappa_d)^{d+1}}{\Gamma(d)} c^* \int_0^\infty \int_{-1}^1 \int_{-1}^1 \mathbf{1}_{\{(w,v,t) \in \cdot\}} \mathbf{1}_{\{q(w,v) > 0\}} b(w)b(v) \\ &\quad \times \frac{(1 - w^2)^{(d^2-2)/2} (1 - v^2)^{(d^2-2)/2}}{q(w, v)^{d^2+d}} t^{d^2+d-1} \\ &\quad \times \exp \left[-\lambda t^d V \left(1, \frac{\sqrt{1 - v^2}}{q(w, v)}, \frac{\sqrt{1 - w^2}}{q(w, v)} \right) \right] dw dv dt, \end{aligned} \tag{5.14}$$

where $c^*, b(w)$, and $q(w, v)$ are as given in (5.12) and (5.9), respectively.

Proof. As

$$\exp[-\lambda V(t, r, \sqrt{t^2 + r^2 - 2trw})] \rightarrow 0 \quad \text{as } t \rightarrow \infty$$

and $[d(\lambda\kappa_d)^d / \Gamma(d)]r^{d^2-1}e^{-\lambda\kappa_d r^d}$ is a density, for all $w \in [-1, 1]$ we obtain

$$\frac{d\lambda(\lambda\kappa_d)^d}{\Gamma(d)} \int_0^\infty \int_0^\infty V^*(t, r, w) \exp[-\lambda V(t, r, \sqrt{t^2 + r^2 - 2trw})]r^{d^2-1} dt dr = 1. \quad (5.15)$$

For measurable mappings $h_1 : [-1, 1]^2 \rightarrow [0, \infty)$ and $h_2 : [0, \infty) \rightarrow [0, \infty)$, from (5.3) we obtain

$$\begin{aligned} & E_{N_1}^0[h_1(\eta', \eta'')h_2(L)] \\ &= \frac{d\lambda(\lambda\kappa_d)^d}{2\Gamma(d)} \iiint \left(h_1\left(\frac{t - rw}{\sqrt{r^2 + t^2 - 2trw}}, w\right) + h_1\left(w, \frac{-rw + t}{\sqrt{r^2 + t^2 - 2trw}}\right) \right) \\ &\quad \times h_2(t)V^*(t, r, w) \exp[-\lambda V(t, r, \sqrt{t^2 + r^2 - 2trw})] \\ &\quad \times r^{d^2-1} dr \mathbb{W}(dw) dt. \end{aligned} \quad (5.16)$$

To modify the inner integration we write

$$r(w, t, v) := \frac{t\sqrt{1 - v^2}}{q(w, v)}, \quad r \geq 0, w, v \in [-1, 1],$$

provided that $q(w, v) > 0$. It is easy to check that, for any fixed $t > 0$ and $w \in (-1, 1)$, $v \mapsto r(w, t, v)$ is a differentiable bijection between $(0, \infty)$ and $\{v \in (-1, 1) : q(w, v) > 0\}$. Moreover, we have

$$\begin{aligned} \frac{\partial}{\partial v} r(w, t, v) &= -\frac{t\sqrt{1 - v^2}}{q(w, v)^2\sqrt{1 - v^2}}, \\ t - r(w, t, v)w &= \frac{tv\sqrt{1 - v^2}}{q(w, v)}, \\ t^2 + r(w, t, v)^2 - 2r(w, t, v)wt &= \frac{t^2(1 - w^2)}{q(w, v)^2}. \end{aligned}$$

Substituting these equations into (5.16) yields

$$E_{N_1}^0[h_1(\eta', \eta'')h_2(L)] = \iiint (h_1(v, w) + h_1(w, v))h_2(t)H(w, t, v) dv dt \mathbb{W}(dw), \quad (5.17)$$

where

$$\begin{aligned} H(w, t, v) &:= \frac{d\lambda(\lambda\kappa_d)^d}{2\Gamma(d)} 1_{\{q(w, v) > 0\}} \frac{t\sqrt{1 - w^2}}{q(w, v)^2\sqrt{1 - v^2}} V^*\left(t, \frac{t\sqrt{1 - v^2}}{q(w, v)}, w\right) \\ &\quad \times \left(\frac{t\sqrt{1 - v^2}}{q(w, v)}\right)^{d^2-1} \exp\left[-\lambda V\left(t, \frac{t\sqrt{1 - v^2}}{q(w, v)}, \frac{t\sqrt{1 - w^2}}{q(w, v)}\right)\right]. \end{aligned} \quad (5.18)$$

In particular,

$$E_{N_1}^0 [h_1(\eta', \eta'')] = \iint (h_1(v, w) + h_1(w, v))H^*(w, v) \, dv \mathbb{W}(dw), \tag{5.19}$$

where

$$H^*(w, v) := \int_0^\infty H(w, t, v) \, dt.$$

Equation (5.15) implies that

$$\int_{-1}^1 H^*(w, v) \, dv = \frac{1}{2}, \quad w \in [-1, 1].$$

Hence, it follows from (5.19) that

$$P(\eta' \in \cdot) = \frac{1}{2} \mathbb{W} + \iint 1_{\{v \in \cdot\}} H^*(w, v) \, dv \mathbb{W}(dw).$$

Using (5.10), we obtain the invariance relationship

$$\mathbb{W} = \int_{-1}^1 \int_{-1}^1 1_{\{v \in \cdot\}} 2H^*(w, v) \, dv \mathbb{W}(dw). \tag{5.20}$$

We next compute the kernel function $H^*(w, v)$. From [8, Equation (11)],

$$V^* \left(t, \frac{t\sqrt{1-v^2}}{q(w, v)}, w \right) = d\kappa_d t^{d-1} \left(\frac{\sqrt{1-w^2}}{q(w, v)} \right)^{d-1} b(v),$$

where $b(v)$ is as given in (5.12). Recalling definition (5.18), we hence obtain

$$2H(w, t, v) = S(w, t, v)(1-v^2)^{(d^2-d-2)/2} b(v), \tag{5.21}$$

where

$$\begin{aligned} S(w, t, v) &:= 1_{\{q(w, v) > 0\}} \frac{d^2(\lambda\kappa_d)^{d+1}}{\Gamma(d)} \frac{(1-w^2)^{d/2}(1-v^2)^{d/2}}{q(w, v)^{d^2+d}} t^{d^2+d-1} \\ &\times \exp \left[-\lambda V \left(t, \frac{t\sqrt{1-v^2}}{q(w, v)}, \frac{t\sqrt{1-w^2}}{q(w, v)} \right) \right]. \end{aligned} \tag{5.22}$$

As $V(t, ta, tb) = t^d V(1, a, b)$ for all $a, b, t \geq 0$, we find from an easy calculation that

$$\int_0^\infty S(w, t, v) \, dt = 1_{\{q(w, v) > 0\}} d^2 \kappa_d^{d+1} \frac{(1-w^2)^{d/2}(1-v^2)^{d/2}}{q(w, v)^{d^2+d}} G(w, v)^{-d-1},$$

where

$$G(w, v) := V \left(1, \frac{\sqrt{1-v^2}}{q(w, v)}, \frac{\sqrt{1-w^2}}{q(w, v)} \right).$$

This is a symmetric function of w and v , and it immediately follows that the right-hand side of (5.11) does satisfy (5.20). Moreover, $H^*(w, v)$ is continuous and positive on $\{(w, v) \in (-1, 1)^2 : q(w, v) > 0\}$, from which (5.11) follows since standard results on discrete-time

Markov processes imply that there exists at most one probability measure \mathbb{W} on $[-1, 1]$ satisfying both (5.20) and $\mathbb{W}(\{-1\}) = \mathbb{W}(\{1\}) = 0$. Indeed, we may consider $2H^*(w, v) dv$ a stochastic kernel on $(-1, 1)$. The associated Markov process has the regeneration set $[-1/2, 1/2]$, for instance. By Theorem VII.3.5 of [1], the process can have at most one invariant measure.

The value (see (5.12)) of the constant c^* in (5.11) is a consequence of [8, Equation (13)]. Finally, we may substitute (5.21), (5.22), and (5.11) into (5.17) to obtain (5.14).

We can now summarize the main results of this section, as follows.

Theorem 5.1. Under $P_{N_1}^0$ the following assertions hold.

- (i) The random variables (η', η'') , $G(\eta', \eta'')L^d$, and Φ_1 are independent.
- (ii) The random variable $G(\eta', \eta'')L^d$ has a gamma distribution with shape parameter $d + 1$ and scale parameter λ .
- (iii) The distribution of (η', η'') has the density

$$d^2 \kappa_d^{d+1} c^* 1_{\{q(w,v) > 0\}} b(w)b(v) \frac{(1-w^2)^{(d^2-2)/2} (1-v^2)^{(d^2-2)/2}}{q(w,v)^{d^2+d}} G(w,v)^{-d-1}.$$

- (iv) The distribution of Φ_1 is $\mathbb{Q}_1(\cdot \times S^{d-1})$, where \mathbb{Q}_1 is as given in (1.3) for $k = 1$.

Proof. Assertion (iv) follows directly from Proposition 5.1 while assertions (ii) and (iii) follow from (5.14) and a simple substitution. Combining Proposition 5.1 and (5.14) yields assertion (i).

Remark 5.1. Under $P_{N_1}^0$ the origin is almost surely the centre of the typical edge. The random variable $G(\eta', \eta'')L^d$ is the volume of the union of the two balls, respectively centred at the endpoints of the typical edge, whose radii are given by the respective distances from the endpoints to one of the neighbours of the edge. This volume, the angles $\arccos \eta'$ and $\arccos \eta''$, and the normalized directions in Φ_1 provide a complete geometric description of the typical edge and its neighbours. Theorem 5.1 gives remarkably explicit formulae for the distribution of these random variables. With a little more effort, Theorem 5.1 can be deduced from Propositions 4.2 and 2.1.

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