

RESEARCH ARTICLE

On approximation of the analytic fixed finite time large t probability distributions in an extreme renewal process with no-mean inter-renewals

Percy H. Brill¹  and Mei Ling Huang² 

¹Departments of Management Science (and) Mathematics and Statistics, University of Windsor, Windsor, ON, Canada.

E-mail: brill@uwindsor.ca

²Department of Mathematics and Statistics, Brock University, St. Catharines, ON, Canada. E-mail: mhaung@brocku.ca.

Keywords: Analytic fixed finite time t pdfs of excess, age, and total life, Finite-mean inter-renewals, Integral equations, Level crossing method, Limiting pdfs of excess, age, and total life, L_1 metric for distance, No-mean inter-renewals, Regenerative process, Renewal process

Abstract

We consider an extreme renewal process with no-mean heavy-tailed Pareto(II) inter-renewals and shape parameter α where $0 < \alpha \leq 1$. Two steps are required to derive integral expressions for the analytic probability density functions (pdfs) of the fixed finite time t excess, age, and total life, and require extensive computations. Step 1 creates and solves a Volterra integral equation of the second kind for the limiting pdf of a basic underlying regenerative process defined in the text, which is used for all three fixed finite time t pdfs. Step 2 builds the aforementioned integral expressions based on the limiting pdf in the basic underlying regenerative process. The limiting pdfs of the fixed finite time t pdfs as $t \rightarrow \infty$ do not exist. To reasonably observe the large t pdfs in the extreme renewal process, we approximate them using the limiting pdfs having simple well-known formulas, in a companion renewal process where inter-renewals are right-truncated Pareto(II) variates with finite mean; this does not involve any computations. The distance between the approximating limiting pdfs and the analytic fixed finite time large t pdfs is given by an L_1 metric taking values in $(0, 1)$, where “near 0” means “close” and “near 1” means “far”.

1. Introduction

The analytic fixed finite time $t > 0$ probability density functions (pdfs) of excess, age, and total life in a renewal process, are of interest theoretically, and to applied probabilists, engineers, and scientists (see [Figure 1](#) in [Section 2.1](#)). Feller [8], Smith [20], Cox [7], Karlin and Taylor [11], Ross [15] pp. 44–45, and others have usually considered these pdfs when the inter-renewals are continuous variates with a finite mean. Recent work in applied probability and stochastic modeling has generated interest in the analytic fixed finite time t pdfs of excess, age, and total life in a renewal processes where the inter-renewals are no-mean, heavy-tailed Pareto(II) variates [2] p. 49 and [12]. Real-world applications of the Pareto(II) distribution are given in, for example, Huang *et al.* [10] and Harris *et al.* [9]. (Note: The Pareto(II) distribution is also called the Lomax distribution [13].)

Here, the extreme renewal process of interest has no-mean Pareto(II) inter-renewals with a shape parameter $\alpha \in (0, 1]$, (formula (3.5), p. 60 in [12]; formula (1.3.5), p. 11 with $\mu = 0$ and $\sigma = 1$ in [2]). We propose a method to approximate the analytic fixed finite time t pdfs of excess, age, and total life for large t in this extreme renewal process. In that process, the analytic fixed finite time t pdfs exist ($0 < t < \infty$), but the corresponding limiting pdfs as $t \rightarrow \infty$ do not exist because the inter-renewals have no mean ([Section 2.2](#)).

Generally, two steps involving extensive computations are required to derive integral expressions for the analytic pdfs of the fixed finite time t excess, age, and total life. Step 1 creates and solves a Volterra integral equation of the second kind for the limiting pdf of a basic underlying regenerative process, denoted by $\{X_{\text{RG}}(s)\}_{s \geq 0}$ (subscript RG indicates “regenerative process”). Process $\{X_{\text{RG}}(s)\}_{s \geq 0}$ is used to derive integral expressions for all three fixed finite time t pdfs (see Figure 3 in Section 3). It has a threshold at the state-space fixed finite level t of interest. Process $\{X_{\text{RG}}(s)\}_{s \geq 0}$ is built of an infinite sequence of i.i.d. (independent and identically distributed) realizations of the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$. In $\{X_{\text{RG}}(s)\}_{s \geq 0}$, the stationary mixed pdf as $s \rightarrow \infty$ exists (e.g., [18,20]). We denote this mixed pdf by $\{\pi_{\text{RG},0}^{(t)}, f_{\text{RG}}^{(t)}(x)\}_{x \in (0,t)}$, where $\pi_{\text{RG},0}^{(t)} = \lim_{s \rightarrow \infty} P(X_{\text{RG}}(s) = 0)$; it is derived using a stochastic LC method (Section 3). Step 2 derives integral expressions for all three analytic fixed finite time t pdfs in $\{Z_n\}_{n=1,2,\dots}$, in terms of $\{\pi_{\text{RG},0}^{(t)}, f_{\text{RG}}^{(t)}(x)\}_{0 < x < t}$ (Section 3.2).

Our goal is to approximate the analytic fixed finite time large t pdfs by applying the corresponding limiting pdfs of a companion renewal process where the inter-renewals are right-truncated Pareto(II) variates. These limiting pdfs have well-known formulas, so their derivations require no computations (see Section 2.5).

We approximate the analytic fixed finite time t pdfs of excess, age, and total life in $\{Z_n\}_{n=1,2,\dots}$ for large t using the corresponding limiting pdfs in the companion renewal process $\{Z_{K,n}^{\text{TR}}\}_{n=1,2,\dots}$ having right-truncated inter-renewals. The right truncation point K of Z_K^{TR} is selected to satisfy a “plausible criterion” for the approximation to be “close” over a nontrivial subset of $(0, K)$ (Section 4).

Section 5.1 gives a measure of distance between the analytic fixed finite time t pdfs in $\{Z_n\}_{n=1,2,\dots}$ and the approximating limiting pdfs in $\{Z_{K,n}^{\text{TR}}\}_{n=1,2,\dots}$, using an L_1 metric taking values in $(0, 1)$. We choose a measure equal to one-half of the area between the two pdfs, because it conveniently takes values in $(0, 1)$. A distance measure “near 0” indicates that the approximation is “close”. A distance measure “near 1” indicates that the approximation is “far”.

Section 2 on preliminaries gives notation and properties about the standard Pareto(II) variable, which are used later in the paper; and for the renewal process whose inter-renewals are Pareto(II) variates. It also gives similar details about the right-truncated Pareto(II) distribution. Section 2.1 defines the fixed finite time- t random variables. Section 2.2 further details the extreme renewal process. Section 2.3 describes the right-truncated Pareto(II) random variables. Section 2.4 gives the expected value of the truncated Pareto(II) random variable, required for the limiting pdfs of the excess, age, and total life variables. Section 2.5 further details the renewal process with right-truncated Pareto(II) inter-renewals, giving the very simple formulas for the limiting pdfs of excess, age, and total life. Section 3 details the underlying regenerative process $\{X_{\text{RG}}(s)\}_{s \geq 0}$. Section 4 approximates the analytic fixed finite time t pdfs (large t) of the extreme renewal process. Section 5 gives an L_1 measure of distance between the analytic fixed finite time t pdfs for large t in the extreme renewal process, and in the corresponding approximating limiting pdfs in the renewal process with right-truncated Pareto(II) inter-renewals.

2. Preliminaries

This section presents preliminary results which are useful in the sequel. It describes the positions of the three fixed finite time t random variables of interest in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$ (Figure 1); specifies the probability distributions of Pareto(II)(x, α), and right-truncated Pareto(II)(x, α) when the truncation point is $K > 0$; details the limiting pdfs of excess, age, and total life in the renewal process with right-truncated inter-renewals (denoted by $\{Z_{K,n}^{\text{TR}}\}_{n=1,2,\dots}$).

2.1. The three fixed finite time t pdfs in the extreme renewal process

Figure 1 shows a sketch of an extreme renewal process where the inter-renewal times are i.i.d. Pareto(II) r.v.s (random variables) $\{Z_n\}_{n=1,2,\dots}$, showing the variates of interest: excess, age, and total life with respect to fixed finite time t .

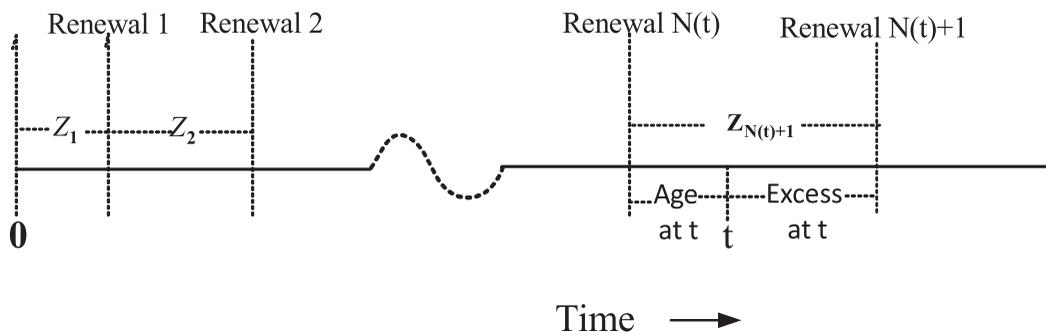


Figure 1. One-dimensional illustration showing the fixed finite time t excess, age, and total life— $Z_{N(t)+1}$ (= age + excess). Figure shows $Z_{N(t)+1}$ relative to the fixed finite time t in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$.

2.2. The Pareto(II) probability distribution for inter-renewals of the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$

Consider the Pareto(II) inter-renewals of the extreme renewal process. Denote the cdf, ccdf (complementary cdf), and pdf of each inter-renewal in $\{Z_n\}_{n=1,2,\dots}$ by: $B(x)$; $\bar{B}(x) := 1 - B(x)$, $0 \leq x < \infty$; and $b(x)$, $0 < x < \infty$, respectively. The inter-renewals $Z_n \stackrel{d}{=} Z \stackrel{d}{=} \text{Pareto(II)}(x, \alpha)$, $0 \leq x < \infty$, $\alpha \in (0, 1]$, where

$$\left. \begin{aligned} B(x) &= 1 - (1 + x)^{-\alpha}, & 0 \leq x < \infty, \\ \bar{B}(x) &= 1 - B(x) = (1 + x)^{-\alpha}, & 0 \leq x < \infty, \\ b(x) &= \frac{d}{dx} B(x) = \alpha(1 + x)^{-\alpha-1}, & 0 < x < \infty, \end{aligned} \right\} \tag{1}$$

where $B(x) \equiv P(Z \leq x)$. Then, Z is heavy-tailed [17]. Since $\alpha \in (0, 1]$, Z has no mean, implying that the limiting pdfs of excess, age, and total life do not exist.

We will use the shape parameter $\alpha = 0.5$ throughout for illustrative purposes. For any value of $\alpha \in (0, 1]$, the analysis would be similar.

Consider the 1-dimensional illustration of the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$ in Figure 1. The fixed finite time point of interest is $t > 0$. Also, $N(t) :=$ number of renewals in the interval $(0, t)$. Let $N_t :=$ number of renewals required to first exceed t , that is, $N_t := \min\{n | \sum_{i=1}^n Z_i > t\}$. Then, N_t is a stopping time for $\{Z_n\}_{n=1,2,\dots}$. Also, $N_t = N(t) + 1$, and $E[N_t] = E[N(t)] + 1 \equiv M(t) + 1$, where $M(t) \equiv E[N(t)]$ is the “renewal function” (see, e.g., [11] pp. 167–169).

In $\{Z_n\}_{n=1,2,\dots}$, the fixed finite time t pdfs of excess, age, and total life exist [4]. But the corresponding limiting pdfs of excess, age, and total life as $t \rightarrow \infty$, do not exist because Z_n has no mean.

2.3. The truncated Pareto(II) probability distribution in the companion renewal process

Define a companion renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ related to, but distinct from $\{Z_n\}_{n=1,2,\dots}$, where the inter-renewals are: $Z_{K,n}^{TR} \stackrel{d}{=} Z_K^{TR} \stackrel{d}{=} \text{right-truncated Pareto(II)}(x, \alpha)$, $0 < x < K$. The superscript “TR” indicates that $K :=$ right truncation point of Z . In $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$, the limiting pdfs of excess, age, and total life, as $t \rightarrow \infty$, do exist because $E[Z_K^{TR}] < \infty$ (finite). Importantly, the limiting pdfs of excess, age, and total life in $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ have relatively simple formulas (Section 2.5).

We use the limiting pdfs of excess, age, and total life in $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ to approximate the analytic fixed finite time t pdfs of interest (large t) in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$. No computations are required since the theoretical formulas of the limiting pdfs are well known. This approximation method saves a great deal of computer time required for the computational derivation of the analytic fixed finite time t pdfs, for even a single large t .

Substituting from formula (1) in Section 2.2, we obtain the truncated cdf $B_K^{TR}(x)$, ccdf $\overline{B}_K^{TR}(x)$, and pdf $b_K^{TR}(x)$, of Z_K^{TR} to be, respectively:

$$\left. \begin{aligned} B_K^{TR}(x) &= \frac{B(x)}{B(K)} = \frac{1 - (1+x)^{-\alpha}}{1 - (1+K)^{-\alpha}}, \quad 0 < x < K, \\ \overline{B}_K^{TR}(x) &= 1 - B_K^{TR}(x) = 1 - \frac{1 - (1+x)^{-\alpha}}{1 - (1+K)^{-\alpha}}, \quad 0 < x < K, \\ b_K^{TR}(x) &= \frac{d}{dx} B_K^{TR}(x) = \frac{\alpha(1+x)^{-\alpha-1}}{1 - (1+K)^{-\alpha}}, \quad 0 < x < K. \end{aligned} \right\} \tag{2}$$

2.4. The expected value of the right-truncated Pareto(II) at level K

For all $\alpha \in (0, 1]$, Z_K^{TR} has a finite mean which is given in formulas (3), when using $\overline{B}_K^{TR}(x)$ in (2).

$$\left. \begin{aligned} E(Z_K^{TR}) &= \int_{x=0}^K \overline{B}_K^{TR}(x) dx = \int_{x=0}^K \left(1 - \frac{1 - (1+x)^{-\alpha}}{1 - (1+K)^{-\alpha}} \right) dx \\ &= \int_{x=0}^K \left(\frac{(1+x)^{-\alpha} - (1+K)^{-\alpha}}{1 - (1+K)^{-\alpha}} \right) dx \\ &= K - \frac{(-\alpha + 1)K - (1+K)^{-\alpha+1} + 1}{(-\alpha + 1)(1 - (1+K)^{-\alpha})}, \quad \text{if } 0 < \alpha < 1; \\ E(Z_K^{TR}) &= \left(1 + \frac{1}{K} \right) \ln(1+K) - 1, \quad \text{if } \alpha = 1. \end{aligned} \right\} \tag{3}$$

In (3), $0 < E(Z_K^{TR}) < K$ (finite) when $K > 0$. (Note: When $\alpha = 1$, $\inf E(Z_K^{TR}) = \inf_{K \downarrow 0} E(Z_K^{TR}) = 0$). Thus, the limiting pdfs of excess, age, and total life in $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ exist as time $\rightarrow \infty$ (Section 2.5).

We explore using the limiting pdfs of $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ to approximate the analytic fixed finite time t pdfs of excess, age, and total life for large t in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$. When constructing the renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ we will need to determine the value of a right truncation point K such that the limiting pdfs of excess, age, and total life approximate the corresponding analytic fixed finite time t pdfs in $\{Z_n\}_{n=1,2,\dots}$ on a nontrivial subset of the interval $(0, K)$ (Section 4). The quality of the approximation will be measured by an L_1 distance measure between the analytic fixed finite time t pdfs in $\{Z_n\}_{n=1,2,\dots}$ and the corresponding (approximating) limiting pdfs in $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ (Section 5).

2.5. The limiting pdfs of excess, age, and total life in the renewal process with right-truncated inter-renewals $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$

Denote the limiting excess, age, and total life by γ_K^{TR} , δ_K^{TR} , and β_K^{TR} , respectively; with corresponding pdfs $f_{\gamma_K}^{TR}(x)$, $0 < x < K$; $f_{\delta_K}^{TR}(x)$, $0 < x < t$; $f_{\beta_K}^{TR}(x)$, $0 < x < K$, where $K > t$. We use the well-known formulas for the limiting pdfs of a standard renewal process where the inter-renewals have a finite mean (e.g., formulas (6.2), (6.5) and (6.6), pp. 193–194, in [11]) and substitute from the formulas in (2). This gives for the renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$:

$$f_{\gamma_K}^{TR}(x) = \frac{1}{E[Z_{K\gamma}^{TR}]} \overline{B}_{K\gamma}^{TR}(x) = \frac{1}{E[Z_{K\gamma}^{TR}]} \left(1 - \frac{1 - (1+x)^{-\alpha}}{1 - (1+K_\gamma)^{-\alpha}} \right), \quad x \in (0, K_\gamma), \quad K_\gamma > t, \tag{4}$$

$$f_{\delta_K}^{TR}(x) = \frac{1}{E[Z_{K\delta}^{TR}]} \overline{B}_{K\delta}^{TR}(x) = \frac{1}{E[Z_{K\delta}^{TR}]} \left(1 - \frac{1 - (1+x)^{-\alpha}}{1 - (1+K_\delta)^{-\alpha}} \right), \quad x \in (0, K_\delta), \quad K_\delta = t, \tag{5}$$

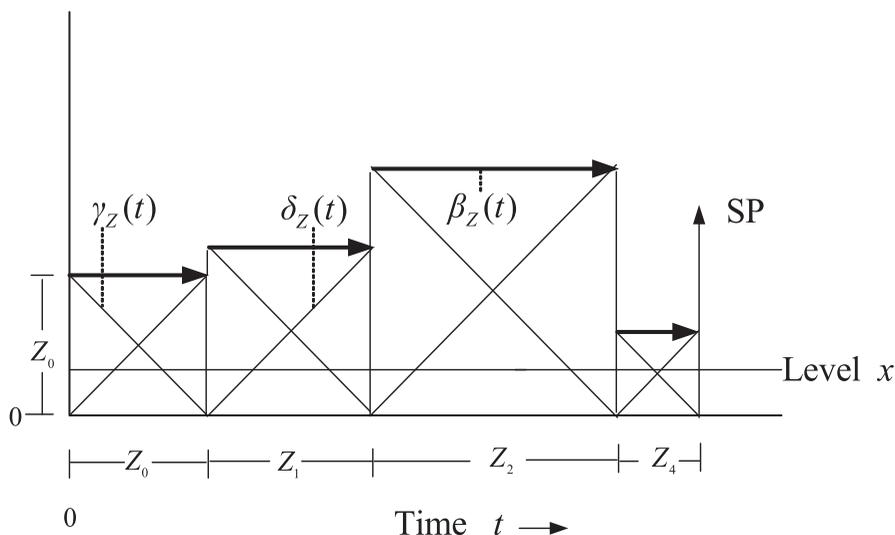


Figure 2. In figure $Z_n \stackrel{d}{=} Z_{K_\zeta}^{TR}$, $\zeta = \gamma, \delta, \beta$. $\gamma_Z(t)$ (slope -1) := limiting excess. $\delta_Z(t)$ (slope $+1$) := limiting age. $\beta_Z(t)$ (slope 0) := limiting total life. Expected regenerative cycles $E[Z_{K_\zeta}^{TR}]$ are all equal.

$$f_{\beta_K}^{TR}(x) = \frac{1}{E[Z_{K_\beta}^{TR}]} x b_{K_\beta}^{TR}(x) = \frac{1}{E[Z_{K_\beta}^{TR}]} x \left(\frac{\alpha(1+x)^{-\alpha-1}}{1 - (1+K_\beta)^{-\alpha}} \right), \quad x \in (0, K_\beta), \quad K_\beta > t, \quad (6)$$

where $E[Z_{K_\zeta}^{TR}]$ is given in (3) on replacing K by K_ζ , where $\zeta = \gamma, \delta$, or β .

We can quickly check formulas (4)–(6) using the elementary renewal theorem (e.g., [16] pp. 432–433) and the renewal reward theorem (e.g., [21] p. 33ff). Note that the sample paths of all three regenerative processes have the same rate out of level 0 (see Figure 2), and have the same expected cycle, for example, $E[Z_{K_\zeta}^{TR}]$. That is, by the elementary renewal theorem the rate out of level 0 is $1/E[Z_{K_\zeta}^{TR}]$, $\zeta = \gamma, \delta, \beta$ and all three regenerative cycles are equal.

Clarifying formula (4), the limiting pdf of excess at time t (same as the limiting pdf of remaining service time at instant t (see Figure 2)). The expected number of upcrossings of arbitrary state-space level x , ($0 < x < K$) in a regenerative cycle is equal to 1 with probability $\overline{B_{K_\gamma}^{TR}}(x)$ (probability that an upward jump representing a service time overshoots level x). By the renewal reward theorem (first two equalities immediately below) and the basic level crossing theorem (third equality immediately below [3]):

$$\frac{E[\text{No. of } x\text{-upcrossings in cycle}]}{E[\text{cycle}]} = \frac{\overline{B_{K_\gamma}^{TR}}(x)}{E[Z_{K_\gamma}^{TR}]} = \lim_{s \rightarrow \infty} \frac{\mathcal{D}_s(x)}{s} = f_{\gamma_K}^{TR}(x), \quad 0 < x < K_\gamma,$$

where $\mathcal{D}_s(x)$ is the number of sample-path downcrossings of level x during a time interval $(0, s)$.

The age at time t is the time that a customer has been in service at instant t (Figure 2). The sample path of the regenerative process for limiting age at t overlays the sample path of the limiting excess at time t . Using the renewal reward theorem and basic LC theorem similarly as for the limiting excess at time t yields the formula

$$\frac{\overline{B_{K_\delta}^{TR}}(x)}{E[Z_{K_\delta}^{TR}]} = \lim_{s \rightarrow \infty} \frac{\mathcal{U}_s(x)}{s} = f_{\delta_K}^{TR}(x), \quad 0 < x < K_\delta,$$

where $\mathcal{U}_s(x)$ is the number of sample path upcrossings of level x during time interval $(0, s)$.

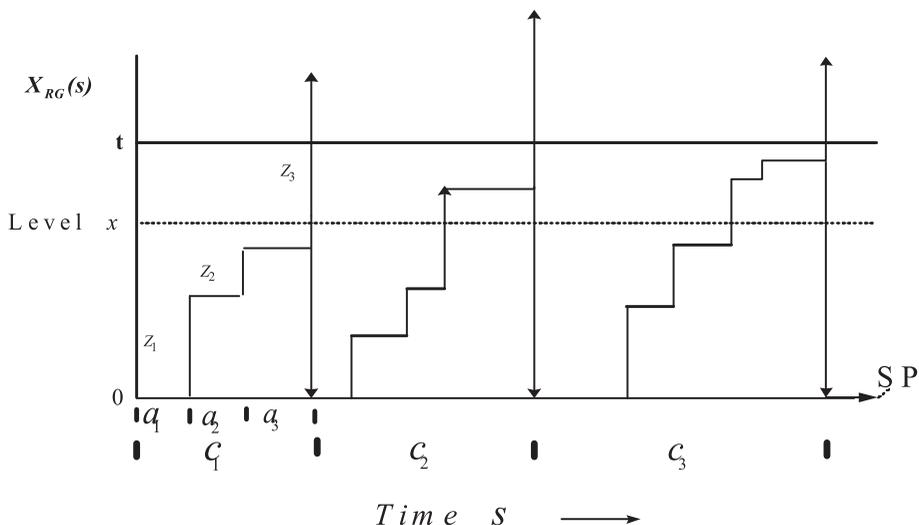


Figure 3. Sample path of underlying regenerative process $\{X_{RG}(s)\}_{s \geq 0}$ with i.i.d. renewal processes $\{Z_n\}_{n=1,2,\dots}$ in the vertical direction within the horizontal regenerative cycles. When an inter-renewal jumps over level t it “double jumps” immediately down to level 0, beginning a new regenerative cycle. The a_i s are the times between vertical jumps. The vertical-jump sizes are inter-renewal times. The a_i s are $\stackrel{d}{=} \text{Exp}_\mu$ random variables, where Exp_μ is an exponential random variable with rate μ . Also shows SP (system point—leading point of the sample path).

The limiting total life at time t is equal to limiting excess + limiting age. It is constant during the service time of the customer in service at time t . It is also a regenerative process (Figure 2). The sample path of limiting total life overlays the sample paths of the limiting excess and limiting age. The renewal reward theorem gives:

$$\frac{\overline{B_{K_\beta}^{\text{TR}}}(x)}{E[Z_{K_\beta}^{\text{TR}}]} = \lim_{t \rightarrow \infty} \frac{\mathcal{U}_t(x)}{t} = \lim_{t \rightarrow \infty} \frac{\mathcal{D}_t(x)}{t} = \int_{y=x}^K \frac{1}{y} f_{\beta_K}^{\text{TR}}(y) dy, \quad 0 < x < K_\beta.$$

Taking d/dx in the left-most and right-most terms in the immediately-above formula gives

$$-\frac{b_{K_\beta}^{\text{TR}}(x)}{E[Z_{K_\beta}^{\text{TR}}]} = -\frac{1}{x} f_{\beta_K}^{\text{TR}}(x), \quad 0 < x < K_\beta,$$

or

$$f_{\beta_K}^{\text{TR}}(x) = \frac{x b_{K_\beta}^{\text{TR}}(x)}{E[Z_{K_\beta}^{\text{TR}}]}, \quad 0 < x < K_\beta.$$

3. Underlying regenerative process and its limiting mixed pdf $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{0 < x < t}$

Construct a sample path of the regenerative process $\{X_{RG}(s)\}_{s \geq 0}$ with state space $[0, \infty)$ having a threshold at the fixed finite level of interest t . The sample path is built of i.i.d. probabilistic replicas of the extreme renewal process $\{Z_n\}_{n=1,\dots,N_t}$ (Figure 3).

The properties of the sample path of $\{X_{RG}(s)\}_{s \geq 0}$ are: (1) $X(0) = 0$; (2) SP (system point—leading point of the sample path which can jump vertically in the state space, for example, “not in Time”; see Chapter 2 in [5]) makes upward jumps $\stackrel{d}{=} Z$ (Pareto(II) variate with shape parameter $\alpha \in (0, 1]$) at an

arbitrary Poisson rate $1/a$ (we select $a := 1$ for simplicity). (3) Whenever a sample-path upward jump upcrosses level t , the sample path of $\{X_{RG}(s)\}_{s \geq 0}$ is prescribed to jump downward instantaneously into level 0, completing a regenerative cycle, and immediately beginning a new independent regenerative cycle. (4) The sample path on each regenerative cycle is a nondecreasing random step function starting at level 0 during every cycle. (5) The sample path's structure rotates the time axis of the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$ in Figure 3 by 90° counterclockwise. This construction results in a sequence of adjoining independent realizations of $\{Z_n\}_{n=1,2,\dots}$ comprising the regenerative cycles (see [4]; also discussed in Section 10.3.1 in [5]).

The process $\{X_{RG}(s)\}_{s \geq 0}$ has a key limiting mixed pdf (e.g., [18,20]). The limiting mixed pdf is denoted by $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{0 < x < t}$, where $\pi_{RG,0}^{(t)} := \lim_{s \rightarrow \infty} P(X_{RG}(s) = 0)$ and $f_{RG}^{(t)}(x)$, $0 < x < t$, is the absolutely continuous part.

The limiting mixed pdf $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{0 < x < t}$ is the time-average limiting pdf. It is the same as the limiting pdf at the sample-path upward jump instants due to the prescribed Poisson arrivals of renewal upward jumps, by the PASTA principle (Poisson Arrivals See Time Averages, Wolff [22]). The limiting cdf of $\{X_{RG}(s)\}_{s \geq 0}$ as $s \rightarrow \infty$, is

$$F_{RG}^{(t)}(x) = \pi_{RG,0}^{(t)} + \int_{y=0}^x f_{RG}^{(t)}(y) dy, \quad 0 \leq x < t, \quad 0 < t < \infty, \tag{7}$$

with normalizing condition

$$F_{RG}^{(t)}(t) = 1. \tag{8}$$

The pdf $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{0 < x < t}$ is the basis for deriving integral expressions for the analytic fixed finite time t pdfs of the excess, age, and total life in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$. (An advantage of the method used here to derive the integral expressions is: the connection with the sample path of the underlying regenerative process gives concrete meaning to the derivations of the integral expressions.)

3.1. Obtaining an integral equation for the key limiting mixed pdf $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{0 < x < t}$ in the underlying regenerative process $\{X_{RG}(s)\}_{s \geq 0}$

Applying the stochastic level crossing method in Sections 3.2.1–3.2.2, p. 198, in Brill [4] (or Section 10.3 in [5]), leads to the Volterra integral equation of the second kind for $f_{RG}^{(t)}(x)$:

$$f_{RG}^{(t)}(x) = \pi_{RG,0}^{(t)} \alpha (1+x)^{-\alpha-1} + \alpha \int_{y=0}^x (1+x-y)^{-\alpha-1} f_{RG}^{(t)}(y) dy, \quad 0 < x < t, \quad \alpha > 0, \tag{9}$$

and the normalizing condition

$$\pi_{RG,0}^{(t)} + \int_0^t f_{RG}^{(t)}(x) dx = 1. \tag{10}$$

Formulas (3.13) and (3.14) in Brill [4] connect the solution of Eqs. (9) and (10) to the classical renewal function $M(x)$ (e.g., [11] p. 169):

$$\pi_{RG}^{(t)} = \frac{1}{1 + M(t)}, \quad f_{RG}^{(t)}(x) = \frac{M'(x)}{1 + M(t)}, \quad 0 < x < t, \tag{11}$$

where $M(t)$ is equal to a series of self-convolutions of the inter-renewal cdf $B(x)$. This series may be time-consuming computationally, and tedious to derive, especially for large t .

Therefore, we use a relatively straightforward numerical procedure (Section 3.2) to obtain a computational solution for the pdf $f_{RG}^{(t)}(x)$, $0 < x < t$ by solving integral Eq. (9). This is followed by the application of Eq. (10) to obtain a computational solution for $\pi_{RG,0}^{(t)}$. (As a by-product, this also gets a computational solution for $M(t)$ by virtue of formula (11).)

This proposed numerical procedure used here is sufficient to gain insight into the approximation of the analytic fixed finite time t pdfs of excess, age, and total life in the renewal process $\{Z_n\}_{n=1,2,\dots}$ for large t by applying the corresponding computational limiting pdfs in the companion renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$.

Section 5 considers a plausible L_1 measure of the distance (discrepancy) between these two pdfs for each time- t random variable, and gives a computed numerical value of the corresponding measure of distance.

(The literature contains various numerical solution techniques to solve Volterra integral equations of the second kind, e.g., [6,14]—and references therein.)

3.2. Details about solving for the limiting mixed pdf $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{x \in (0,t)}$ of the regenerative process $\{X_{RG}(s)\}_{s \geq 0}$

We solve Eq. (9) with the normalizing condition (10) to derive $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{x \in (0,t)}$ using directly the Riemann–Stieltjes definition of an integral on a finite interval (e.g., [1] p. 141). The resulting numerical solution is a step function with respect to a preassigned partition of state-space interval $(0, t)$ with a norm $h > 0$. To get a good computational solution for $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{x \in (0,t)}$, we choose small h such that $t = Nh$ where N is a positive integer.

3.2.1. Use of Pareto(II) shape parameter $\alpha = 0.5$ and large fixed finite time $t = 400$

In the remainder of this paper, we consider a generic case where the shape parameter α of the Pareto(II)-distributed inter-renewals of the process $\{Z_n\}_{n=1,2,\dots}$ is $\alpha = 0.5$ and the “large fixed finite time of interest” is $t := 400$, unless otherwise stated. These values help to increase insight in approximating the analytic fixed finite time t pdfs of excess, age, and total life for large t in the renewal process $\{Z_n\}_{n=1,2,\dots}$ by using the limiting pdfs in the companion renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$, given in formulas (4)–(6) in Section 2.5.

3.2.2. Details about computer program to compute limiting mixed pdf $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{x \in (0,t)}$ of the regenerative process $\{X_{RG}(s)\}_{s \geq 0}$

Let $f_{RG*}^{(t)}(x) = f_{RG}^{(t)}(x)/\pi_{RG,0}^{(t)}$, $x \in (0, t)$. This transforms Eq. (9) into the following integral equation for $f_{RG*}^{(t)}(x)$:

$$f_{RG*}^{(t)}(x) = \alpha(1+x)^{-\alpha-1} + \alpha \int_{y=0}^x (1+x-y)^{-\alpha-1} f_{RG*}^{(t)}(y) dy, \quad 0 < x < t. \tag{12}$$

The computation consists of the following five steps.

Step 1. Compute and store in computer memory:

$$b(ih) := \alpha(1+ih)^{-\alpha-1}, \quad i = 0, \dots, N, \quad h > 0. \tag{13}$$

Step 2. Using Eq. (12), start the computation with $i = 0$, that is,

$$f_{RG*}^{(t)}(x) = b(0) + 0 = \alpha, \quad 0 < x < h. \tag{14}$$

Compute for $i = 1, \dots, N$,

$$\begin{aligned} f_{RG*}^{(t)}(ih) &= \alpha(1+ih)^{-\alpha-1} + \int_{y=0}^{ih} \alpha(1+ih-y)^{-\alpha-1} f_{RG*}^{(t)}(y) dy \\ &= \alpha(1+ih)^{-\alpha-1} + \sum_{j=1}^i \int_{y=(j-1)h}^{jh} \alpha(1+ih-y)^{-\alpha-1} f_{RG*}^{(t)}(y) dy, \quad i = 1, \dots, N. \end{aligned} \tag{15}$$

where $f_{RG^*}^{(t)}(y) = f_{RG^*}^{(t)}((i-1)h)$, $(i-1)h < y < ih$, $i = 1, \dots, N$.

This gives the values of $f_{RG^*}^{(t)}(ih)$, $i = 1, \dots, N$.

Step 3. From Step 2, construct the step function

$$f_{RG^*}^{(t)}(x) = f_{RG^*}^{(t)}((i-1)h), \quad x \in ((i-1)h, ih), \quad i = 1, \dots, N. \tag{16}$$

Step 4. Compute the estimation for

$$\int_{x=0}^t f_{RG^*}^{(t)}(x) dx \approx h \sum_{i=1}^N f_{RG^*}^{(t)}(ih). \tag{17}$$

Step 5. Denote the computational solution for $\{\pi_{RG,0}^{(t)}, f_{RG}^{(t)}(x)\}_{x \in (0,t)}$ by $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{x \in (0,t)}$. Then,

$$\widehat{f}_{RG}^{(t)}(x) = \widehat{\pi}_{RG,0}^{(t)} \widehat{f}_{RG^*}^{(t)}(x), \quad 0 < x < t, \tag{18}$$

$$\begin{aligned} \widehat{\pi}_{RG,0}^{(t)} + \int_0^t \widehat{f}_{RG}^{(t)}(x) dx &= \widehat{\pi}_{RG,0}^{(t)} + \widehat{\pi}_{RG,0}^{(t)} \int_0^t \widehat{f}_{RG^*}^{(t)}(x) dx = 1 \\ &\approx \widehat{\pi}_{RG,0}^{(t)} + \widehat{\pi}_{RG,0}^{(t)} h \sum_{i=1}^N \widehat{f}_{RG^*}^{(t)}(ih) dx = 1, \end{aligned} \tag{19}$$

giving

$$\widehat{\pi}_{RG,0}^{(t)} \approx \frac{1}{1 + h \sum_{i=1}^N \widehat{f}_{RG^*}^{(t)}(ih)}. \tag{20}$$

To make the computational solutions concrete, we assume, in addition to $\alpha = 0.5$, $t = 400$, that $h = 0.1$, $N = 4000$.

In the sequel, we denote pdfs and related quantities which we derive computationally, by using a “hat”, for example, “pdf”. Quantities which are known or derived from theory will not have a “hat”.

The computational mixed pdf $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{0 < x < t}$ is plotted in Figure 4.

The computational solutions can be improved in various ways, for example, making the norm size h smaller and the corresponding N larger; using trapezoidal areas instead of rectangular areas in intervals $((i-1)h, ih)$, $i = 1, \dots, N$; etc. (see [6]). Section 4 displays the analytic fixed finite time t pdfs of excess, age, and total life in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$, with the corresponding limiting pdfs of the companion renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$, given in Section 2.5. (Note: The limiting pdfs are known from theory and do not require any extra computation here.)

3.3. Obtaining integral expressions for the analytic fixed finite time t pdfs in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$, $0 < t < \infty$

Using the computational solutions for the mixed pdf $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{0 < x < t}$ of the regenerative process $\{X_{RG}(s)\}_{s \geq 0}$, we create integral expressions of the analytic fixed finite time t pdfs in $\{Z_n\}_{n=1,2,\dots}$.

Denote the analytic fixed finite time t excess, age, and total life by γ_t , δ_t , and β_t , respectively, with corresponding pdfs: $\widehat{f}_{\gamma_t}(x)$, $0 < x < \infty$; $\{\widehat{\pi}_{\delta_t}, \widehat{f}_{\delta_t}(x)\}_{0 < x < t}$ (Note: the atom $\widehat{\pi}_{\delta_t} \approx P(\delta_t = t)$; the computational pdf of total life is denoted by $\widehat{f}_{\beta_t}(x)$, $0 < x < \infty$.)

For each large $t > 0$, evaluating the pdf $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{0 < x < t}$ may require a time-consuming computational solution of a Volterra integral equation of the second kind, due to the nature of the Pareto(II) distribution. Increasing values of t require corresponding increasing computation times.)

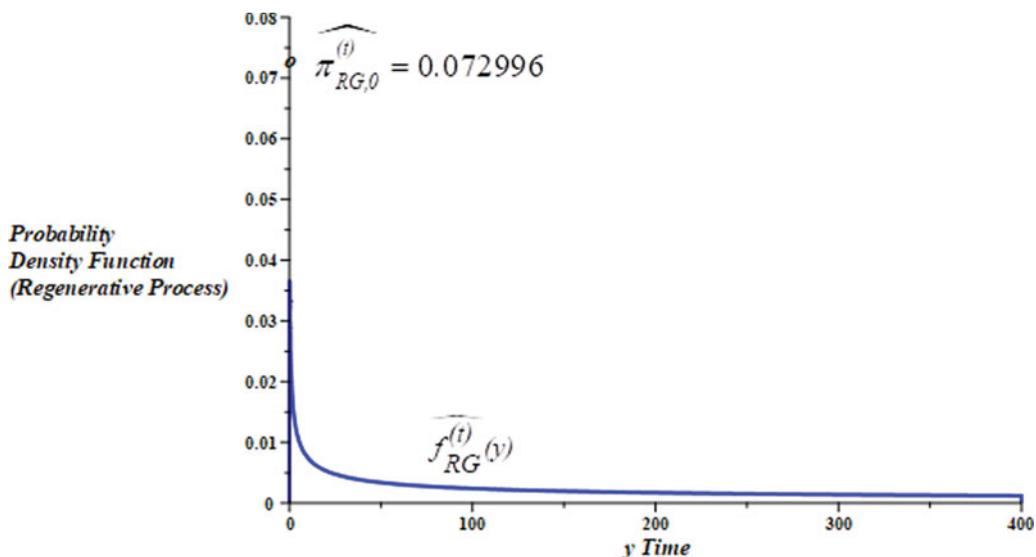


Figure 4. Computational mixed limiting pdf of the regenerative process $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{x \in (0,t)}$ (see Eq. (9)). $\alpha = 0.5, t = 400; \widehat{\pi}_{RG,0}^{(t)} = 0.072996$.

Formulas (21)–(23) are obtained from formulas derived in Brill [4]. For example, formula (4.4a) in [4] gives

$$f_{\gamma_t}(x) = b(t+x) + \int_{y=0}^t b(t+x-y) \frac{f_{RG}^{(t)}(y)}{\pi_{RG,0}^{(t)}} dy, \quad 0 < x < \infty. \tag{21}$$

Formula (4.11) in [4] gives

$$f_{\delta_t}(x) = \bar{B}(x) \frac{f_{RG}^{(t)}(t-x)}{\pi_{RG,0}^{(t)}}, \quad 0 < x < t; \quad \pi_{\delta_t} = \bar{B}(t). \tag{22}$$

The formula for $f_{\beta_t}(x), 0 < x < \infty$, is given in formula (4.13) in [4], and also in formula (23) for convenience.

$$f_{\beta_t}(x) = \begin{cases} b(x) \int_{y=t-x}^t \frac{f_{RG}^{(t)}(y)}{\pi_{RG,0}^{(t)}} dy, & 0 < x < t, \\ b(x) \left(1 + \int_{y=t-x}^t \frac{f_{RG}^{(t)}(y)}{\pi_{RG,0}^{(t)}} dy \right), & t \leq x < \infty. \end{cases} \tag{23}$$

The formulas (21)–(23), and those for $\bar{B}(x)$ and $b(x)$ in formula (1), are valid when the inter-renewals are $\overset{d}{=}$ Pareto(II) variate with shape parameter $\alpha \in (0, 1]$.

Remark 1. Formula (22) shows that δ_t has an atom at level t with probability $\pi_{\delta_t} = (1+t)^{-\alpha}$. Formula (23) shows that $f_{\beta_t}(x)$ has a discontinuity at level $x = t$ of size

$$f_{\beta_t}(t^+) - f_{\beta_t}(t^-) = b(t) = \alpha(1+t)^{-\alpha-1}. \tag{24}$$

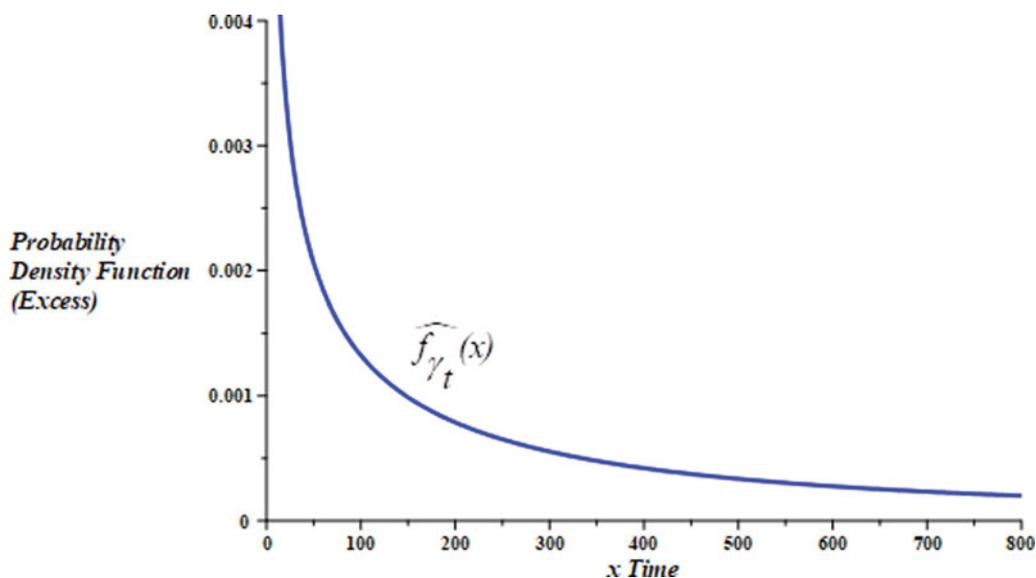


Figure 5. $\widehat{f}_{\gamma_t}(x)$, $0 < x < 800$, $\alpha = 0.5$, $t = 400$, $h = 0.1$, $N = 4000$.

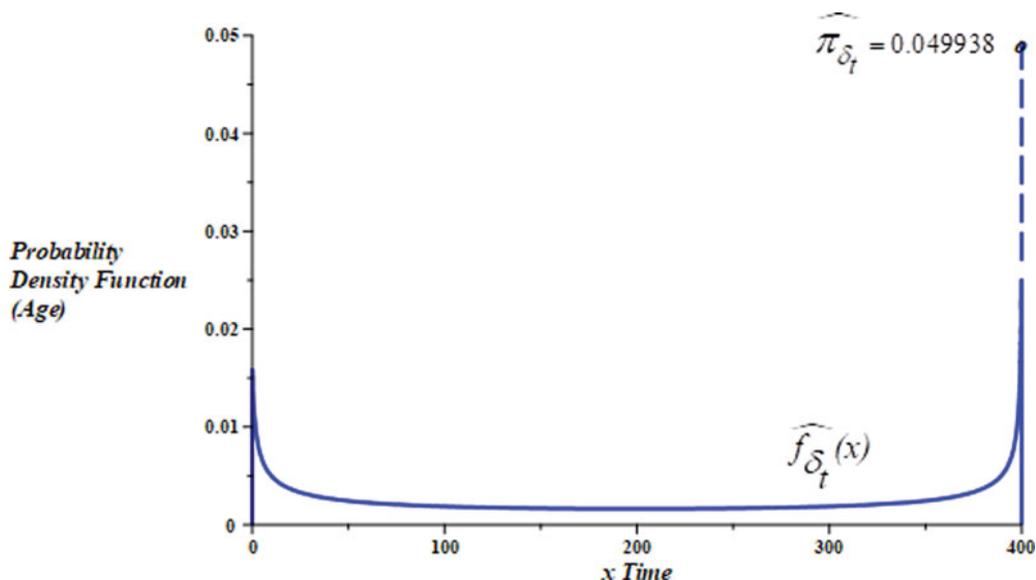


Figure 6. $\{\widehat{\pi}_{\delta_t}, \widehat{f}_{\delta_t}(x)\}_{0 < x < t}$, $\alpha = 0.5$, $t = 400$, $h = 0.1$, $N = 4000$. $\widehat{\pi}_{\delta_t} = 0.049938$.

We obtain computational results for the analytic fixed finite time t pdfs in formulas (21)–(23) by substituting in them the computational step function $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{0 < x < t}$ given in formulas (18) and (20) for the pdf of the regenerative process $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{0 < x < t}$, where $t \in (0, \infty)$.

The computational analytic fixed finite time t pdfs in $\{Z_n\}_{n=1,2,\dots} : \widehat{f}_{\gamma_t}(x), x > 0$, $\{\widehat{\pi}_{\delta_t}, \widehat{f}_{\delta_t}(x)\}_{0 < x < t}$, and $\widehat{f}_{\beta_t}(x), x > 0$, are specified directly in terms of $\{\widehat{\pi}_{RG,0}^{(t)}, \widehat{f}_{RG}^{(t)}(x)\}_{0 < x < t}$ in $\{X_{RG}(s)\}_{s \geq 0}$ (Section 3.2.2). Examples of these estimated analytic fixed finite time t pdfs are plotted in Figures 5, 6, and 7, respectively.

We give further discussion on Figures 6 and 7 in Section 5.

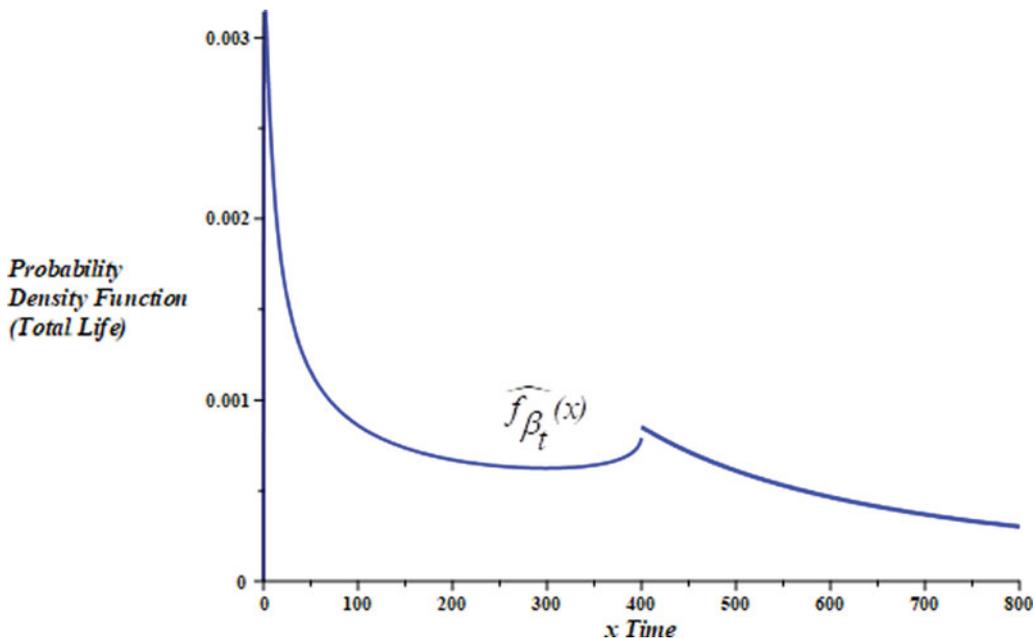


Figure 7. $\widehat{f}_{\beta_t}(x)$, $0 < x < 800$, $\alpha = 0.5$, $t = 400$, $h = 0.1$, $N = 4000$. Discontinuity at time point t is equal to $b(t) = \alpha(1 + t)^{-\alpha-1}$.

4. Using limiting pdfs of $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ to approximate analytic fixed finite time t pdfs of $\{Z_n\}_{n=1,2,\dots}$ (fixed finite large t)

We use the limiting pdfs $f_{\zeta_K}^{TR}(x)$, $0 < x < \zeta_K$, ($\zeta = \gamma, \delta, \beta$) given by formulas (4)–(6) in Section 2.5, to approximate the corresponding analytic fixed finite time t pdfs $\widehat{f}_{\zeta_t}(x)$, $0 < x < K_\zeta$ ($\zeta = \gamma, \delta, \beta$) in the extreme renewal process $\{Z_n\}_{n=1,2,\dots}$ (t large). Computing analytic pdfs $\widehat{f}_{\zeta_t}(x)$, $\zeta = \gamma, \delta, \beta$, for $\{Z_n\}_{n=1,2,\dots}$ is time-consuming, requiring numerical solutions of Volterra integral equations and related quantities (Section 3.3). This tediousness increases motivation to apply the no-computation well-known limiting pdfs $f_{\zeta_K}^{TR}(x)$ s to approximate $\widehat{f}_{\zeta_t}(x)$, $\zeta = \gamma, \delta, \beta$, (t large).

4.1. Selecting right truncation points of the inter-renewals in the renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$

We derive the right truncation points of the Pareto(II) variate, K_γ , K_δ , and K_β , which are useful for approximating the fixed finite t pdfs of excess, age, and total life, respectively. The resulting right-truncated inter-renewals, denoted by $Z_{K_\gamma}^{TR}$, $Z_{K_\delta}^{TR}$, and $Z_{K_\beta}^{TR}$, respectively, have pdfs given by formulas (4)–(6) on interval $(0, K_\zeta)$, $\zeta = \gamma, \delta, \beta$. The $Z_{K_\zeta}^{TR}$ s in the renewal process $\{Z_{K_\zeta,n}^{TR}\}_{n=1,2,\dots}$ have the same shape parameter α as the Pareto(II) variate in the original extreme renewal process of interest $\{Z_n\}_{n=1,2,\dots}$.

4.2. Details for approximating the analytic fixed finite time pdfs of γ_t , δ_t , and β_t

4.2.1. Details for approximating the pdf of γ_t

Letting $x \downarrow 0$ in formula (21) gives

$$f_{\gamma_t}(0^+) = \alpha(1 + t)^{-\alpha-1} + \alpha \int_{y=0}^t (1 + t - y)^{-\alpha-1} \frac{f_{RG}^{(t)}(y)}{\pi_{RG,0}^{(t)}} dy, \quad 0 < t < \infty. \tag{25}$$

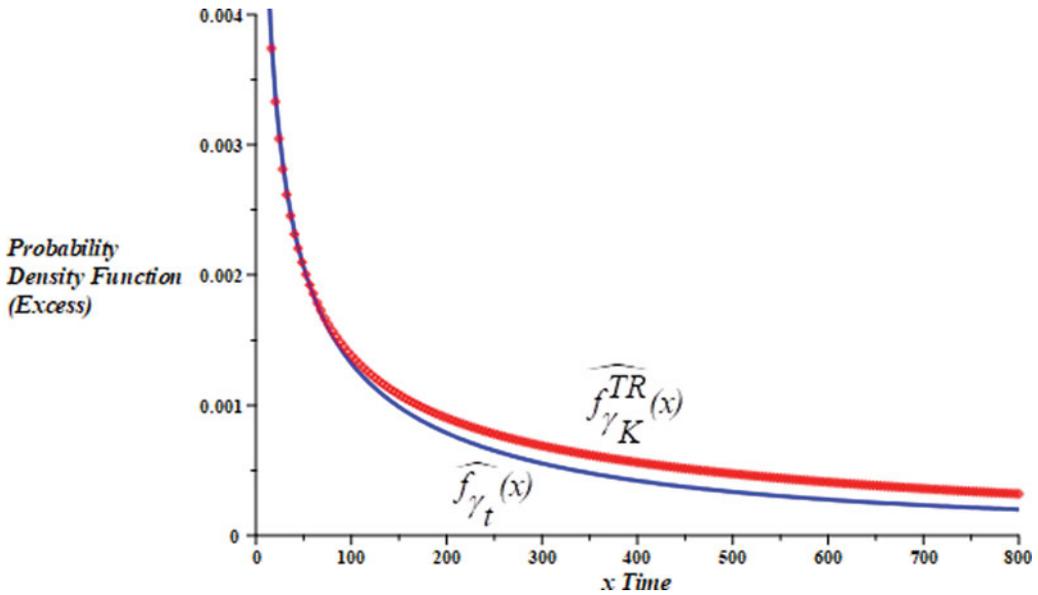


Figure 8. Analytical time t pdf of excess $\widehat{f}_{\gamma_t}(x)$ (solid blue line) versus approximating pdf $f_{\gamma_K}^{TR}(x)$ (red line), $0 < x < 800$. $\alpha = 0.5$, fixed finite $t = 400$.

Since $\overline{B_{K_\gamma}^{TR}}(0) = 1$ in (4), we equate the values of formula (25) and (4) at $x = 0^+$ and solve for K_γ in the equation

$$f_{\gamma_t}(0^+) = \frac{1}{E[Z_{K_\gamma}^{TR}]}, \tag{26}$$

where $E[Z_{K_\gamma}^{TR}]$ is given in formula (3). The solution of (26) is K_γ for the inter-renewals in $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$. We shall use the Policy expressed by Eq. (26) for the approximation procedure because it is plausible and of theoretical interest.

When $t = 400$ (considered here as a large t) and $\alpha = 0.5$, Eq. (26) gives $K_\gamma = 3877.5672$ and $E(Z_{K_\gamma}^{TR}) = 61.2781$. The right truncation point K_γ is significantly greater than $t = 400$. Pdf $\widehat{f}_{\gamma_t}(x)$ versus pdf $f_{\gamma_K}^{TR}(x)$ is plotted in Figure 8. The closeness of the two pdfs appears to be good.

4.2.2. Details for approximating the pdf of δ_t

From the discussion in Section 2.5 and Figure 2, we have

$$E[Z_{K_\delta}^{TR}] = E[Z_{K_\beta}^{TR}] = E[Z_{K_\gamma}^{TR}]. \tag{27}$$

Note that $K_\delta = K_\gamma = 3877.5672$. In pdf, $\{\widehat{\pi}_{\delta_t}, \widehat{f}_{\delta_t}(x)\}_{0 < x < t, \widehat{\pi}_{\delta_t} > 0}$ is a small probability when t is large. The absolutely continuous part of the pdf $\widehat{f}_{\delta_t}(x)$ versus $f_{\delta_K}^{TR}(x)$, $0 < x < t$ ($t = 400$), is plotted in Figure 9. The closeness between the two pdfs appears to be good on the nontrivial interval $(0, 200)$, the discrepancy increases rapidly on $(200, 400)$.

4.2.3. Details for approximating the pdf of β_t

Using an equation similar to Eq. (26) and letting $x = 0$ gives “ $0 = 0$ ”, which does not provide an equation directly for K_β . Hence, we will use a plausible $K_\beta > 0$. We now compare analytic pdf $\widehat{f}_{\beta_t}(x)$ with the corresponding approximating pdf $f_{\beta_K}^{TR}(x)$. When $t = 400$ and $\alpha = 0.5$, we use $K_\beta = 3877.5672$; giving $E[Z_{K_\beta}] = 61.2781$, the same as for K_γ and for K_δ . Pdf $\widehat{f}_{\beta_t}(x)$ versus limiting pdf $f_{K_\beta}(x)$ is plotted in Figure 10.

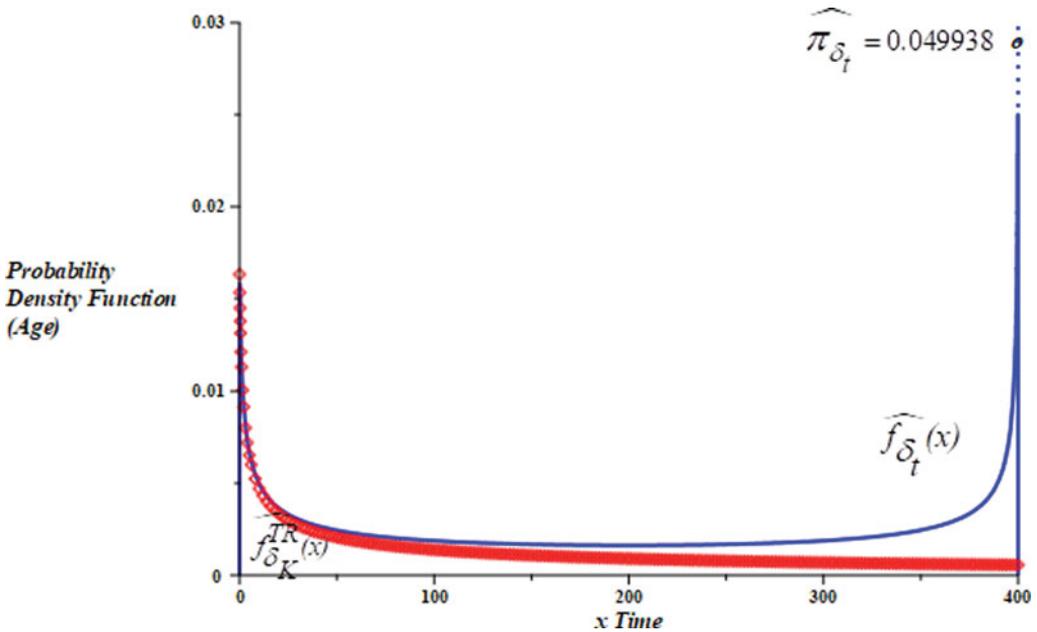


Figure 9. Analytic time t pdf of $\widehat{f}_{\delta_t}(x)$ (solid blue line) versus $f_{\delta_K}^{TR}(x)$ (red line), $0 < x < 400$. $\alpha = 0.5$, fixed finite $t = 400$.

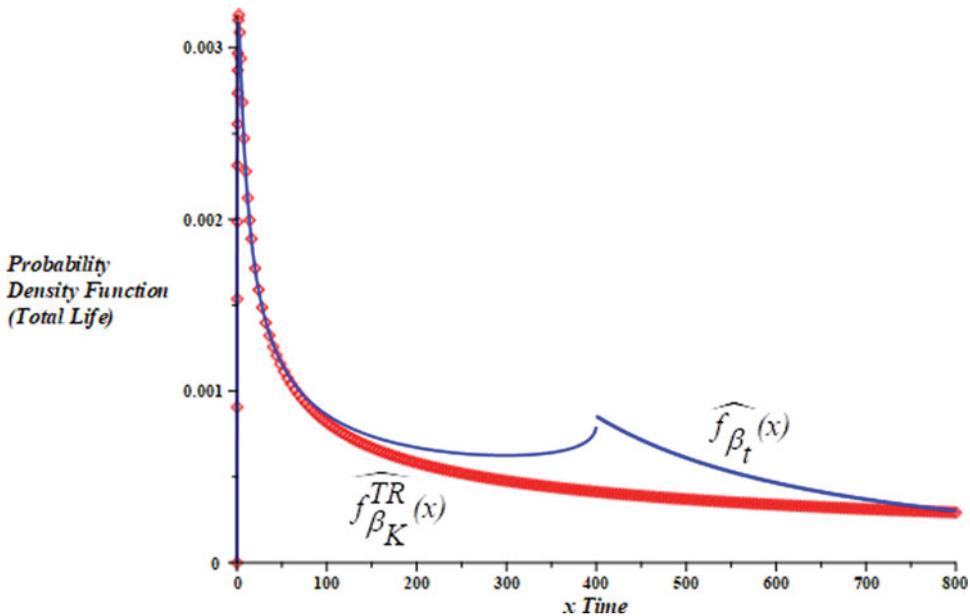


Figure 10. $\widehat{f}_{\beta_t}(x)$ (blue line) versus $f_{\beta_K}^{TR}(x)$ (red line), $0 < x < 800$. $\alpha = 0.5$, fixed finite $t = 400$.

5. Measure of distance between the analytic fixed finite time t pdfs and the limiting pdfs of the renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$

We quantify the notion: “The limiting pdf $f_{\zeta_K}(\cdot)$ of the renewal process $\{Z_{K,n}^{TR}\}_{n=1,2,\dots}$ approximates the analytic fixed finite time t pdf $\widehat{f}_{\zeta_t}(\cdot)$ for $\zeta = \gamma, \delta, \beta$ ”, by using an L_1 measure based on the metric $|\widehat{f}_{\zeta_t}(x) - f_{\zeta_K}^{TR}(x)|$, $0 < x < K_\zeta$. This metric leads to an integral measure for the distance between the

analytic pdfs $\widehat{f}_{\zeta}(x)$ and the corresponding approximating limiting pdfs $f_{\zeta_K}^{TR}(x)$ for $x \in (0, K_{\zeta})$, $\zeta = \gamma, \delta, \beta$.

5.1. The metric $|\widehat{f}_{\zeta}(x) - f_{\zeta_K}^{TR}(x)|$, $\zeta = \gamma, \delta, \beta$

We use a measure of distance (discrepancy) motivated by the areas between the fixed time t pdfs $\widehat{f}_{\zeta}(x)$ and the approximating pdfs $f_{\zeta_K}^{TR}(x)$. The areas between the pdfs must be in the interval $(0, 2)$ since the areas below each pdf is equal to 1. The L_1 measure between the pdfs accounts for all $x \in (0, K_{\zeta})$, $\zeta = \gamma, \delta, \beta$. Denote this measure by

$$\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR}) = \frac{1}{2} \int_{x=0}^{K_{\zeta}} |\widehat{f}_{\zeta}(x) - f_{\zeta_K}^{TR}(x)| dx, \quad \zeta = \gamma, \delta, \beta, \tag{28}$$

where

$$0 \leq \rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR}) < 1. \tag{29}$$

The measure $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR}) = 0$ iff $\widehat{f}_{\zeta}(x) = f_{\zeta_K}^{TR}(x)$ for all $x \in (0, K_{\zeta})$. Moreover, $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR}) < 1$. The less the distance $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR})$, the better is the approximation (i.e., $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR})$ is nearer to 0). The greater the distance $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR})$, the worse is the approximation (i.e., $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR})$ is nearer to 1).

(Other measures of discrepancy between pdfs are given on p. 35 and in Section 3.7 in [19].)

5.2. Example values of the measure $\rho(\widehat{f}_{\zeta}, f_{\zeta_K}^{TR})$

In the three examples here, the shape parameter of the Pareto(II) distribution is $\alpha = 0.5$.

5.2.1. Excess when fixed time $t = 400$

For the example in Figure 8, we obtain $\rho(\widehat{f}_{\gamma_t}, f_{\gamma_K}^{TR}) = \mathbf{0.198468}$. The distance is uniformly quite close on the entire time interval $(0, 800)$. The approximation of pdf \widehat{f}_{γ_t} by using pdf $\rho(\widehat{f}_{\gamma_t}, f_{\gamma_K}^{TR})$ is good.

5.2.2. Age when fixed time $t = 400$

For the example in Figure 9, we obtain $\rho(\widehat{f}_{\delta_t}, f_{\delta_K}^{TR}) = \mathbf{0.475558}$. The distance between the two pdfs is quite close on $(0, 200)$. It is not uniformly close on $(0, 400)$; starting at approximately time 200, the discrepancy increases rapidly on the interval $(200, 400)$. In addition, there is an atom with probability 0.049938 at time t . The approximation of pdf \widehat{f}_{δ_t} using $\rho(\widehat{f}_{\delta_t}, f_{\delta_K}^{TR})$ is medium. Generally, $\rho(\widehat{f}_{\delta_t}, f_{\delta_K}^{TR})$ gives the analyst an idea of where the pdf $\widehat{f}_{\delta_t}(x)$ is located relative to $f_{\delta_K}^{TR}(x)$.

5.2.3. Total life when fixed time $t = 400$

For the example in Figure 10, we obtain $\rho(\widehat{f}_{\beta_t}, f_{\beta_K}^{TR}) = \mathbf{0.281637}$. The distance between the two pdfs is close on intervals $(0, 150)$ and on $(650, 800)$. The overall distance is not uniformly close. The approximation of pdf \widehat{f}_{β_t} using $\rho(\widehat{f}_{\beta_t}, f_{\beta_K}^{TR})$ is between good and medium. There is a rise in discrepancy starting at approximately time 150 and increasing rapidly on $(150, 400)$. This rise appears similar to the discrepancy pattern in Figure 9 on interval $(200, 400)$, for approximating $\widehat{f}_{\delta_t}(x)$. The variate of “total life” is equal to “age + excess” (see Figure 1). In Figure 10, on interval $(400, 800)$, the distance decreases steadily, resembling Figure 8 for $\rho(\widehat{f}_{\gamma_t}, f_{\gamma_K}^{TR})$. The discrepancy patterns for all three fixed finite time t pdfs appear to be related.

6. Conclusion

- (1) This article proposes a technique for approximating the analytic fixed finite time t pdfs for large t , of excess, age, and total life, in a renewal process where the inter-renewals have a no-mean heavy-tailed distribution (extreme renewal process).

- (2) For the inter-renewals, we use the Pareto(II) variate (e.g., [2]) with shape parameter $\alpha \in (0, 1]$, as a generic example throughout the article, because of its common occurrence in recent research and applications.
- (3) Obtaining the analytic fixed finite time t pdfs (t large) of excess age, and total life requires time-consuming computations. The greater t is, the more time-consuming is the computation. The approximating pdfs are the limiting pdfs of the fixed finite time t variates as $t \rightarrow \infty$ in a companion renewal process, denoted by $\{Z_{K,n}^{\text{TR}}\}_{n=1,2,\dots}$, where the inter-renewals are right-truncated Pareto(II) variates with finite mean, finite right truncation point $K > 0$ and the same shape parameter $\alpha \in (0, 1]$. The limiting pdfs in $\{Z_{K,n}^{\text{TR}}\}_{n=1,2,\dots}$ exist because the inter-renewals have a finite mean, and have well-known formulas available from theory, not requiring any computation in the presented technique.
- (4) An L_1 distance measure between the analytic fixed finite time t pdfs of the excess, age, and total life (large t) in the extreme renewal process of interest $\{Z_n\}_{n=1,2,\dots}$, and the approximating limiting pdfs of excess, age, and total life in the companion renewal process $\{Z_{K,n}^{\text{TR}}\}_{n=1,2,\dots}$ is considered in Section 5.

Acknowledgments. The authors thank the reviewers for insightful comments which helped to improve the paper.

Funding statement. This research is supported by the Natural Sciences and Engineering Research Council of Canada [PHB-RGPIN-2014-05697; MLH-DDG-2019-04206].

References

- [1] Apostol, T.M. (1974). *Mathematical analysis*, 2nd ed. Reading, MA: Addison Wesley Publishing Company.
- [2] Arnold, B.C. (2015). *Pareto distributions*, 2nd ed. Boca Raton: CRC Press, Taylor & Francis Group.
- [3] Brill, P.H. (1975). A new methodology for analyzing a broad class of exponential queues. In *5th Conference on Stochastic Processes & their Applications*. University of Maryland, in *Advances in Applied Probability* 8(2), June 1976, p. 242 (invited by J. Keilson and R. Syski).
- [4] Brill, P.H. (2014). Alternative analysis of finite-time probability distributions of renewal theory. *Probability in the Engineering and Informational Sciences* 28(2): 183–201.
- [5] Brill, P.H. (2017). *Level crossing methods in stochastic models*, 2nd ed. New York: Springer.
- [6] Brunner, H. (2010). Seven lectures on theory and numerical solution of Volterra integral equations. <http://www.math.hkbu.edu.hk/hbrunner/harbin10/HL1.pdf>
- [7] Cox, D.R. (1962). *Renewal theory (reprinted 1970)*. London: Methuen.
- [8] Feller, W. (1966). *An introduction to probability theory and its applications*, vol. II. New York: John Wiley.
- [9] Harris, C.M., Brill, P.H., & Fischer, M.J. (2000). Internet-type queues with power-tailed interarrival times and computational methods for their analysis. *INFORMS Journal on Computing* 12(4): 261–271.
- [10] Huang, M.L., Coia, V., & Brill, P.H. (2013). A cluster truncated Pareto distribution and its applications. *ISRN Probability and Statistics* 2013: Article ID 265373, 10 pages. doi:10.1155/2013/265373
- [11] Karlin, S. & Taylor, H.M. (1975). *A first course in stochastic processes*, 2nd ed. New York: Academic Press.
- [12] Kleiber, C. & Kotz, S. (2003). *Statistical size distribution in economics and actuarial sciences*. New York: John Wiley.
- [13] Lomax, K.S. (1954). Business failures: Another example of the analysis of failure data. *Journal of the American Statistical Association* 49: 847–852.
- [14] Press, W.H. & Teukolsky, S.A. (1990). Fredholm and Volterra integral equations of the second kind. *Computers in Physics* 4: 554. doi:10.1063/1.4822946
- [15] Ross, S.M. (1970). *Applied probability with optimization applications*. San Francisco: Holden Day.
- [16] Ross, S.M. (2010). *Introduction to probability models*, 10th ed. Elsevier.
- [17] Sigman, K. (1999). A primer on heavy-tailed distributions. *Queueing Systems* 33: 261–275.
- [18] Sigman, K. & Wolff, R.W. (1993). A review of regenerative processes. *SIAM Review* 35(2): 269–286.
- [19] Silverman, B.W. (1986). *Density estimation for statistics and data analysis*. New York: Chapman & Hall.
- [20] Smith, W.L. (1955). Regenerative stochastic processes. *Proceedings of the Royal Society A* 232: 6–31.
- [21] Tijms, H.C. (2003). *A first course in stochastic models*. Wiley.
- [22] Wolff, R.W. (1982). Poisson arrivals see time averages. *Operations Research* 30(2): 223–231.