

AN ASYMPTOTIC EXPANSION FOR A CLASS OF MULTIVARIATE NORMAL INTEGRALS*

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1. Introductory Discussion and Summary.

Let $\mathbf{x} = (x_1, x_2, \dots, x_n)$ be a normal random vector with zero expectation vector and with a variance-covariance matrix which has 1 for its diagonal elements and ρ for its off-diagonal elements. Consider the quantity

$$(1.1) \quad I_n(h; \rho) = (2\pi)^{-n/2} \{1 + (n-1)\rho\}^{-1/2} (1-\rho)^{-(n-1)/2} \int_h^\infty \dots \int_h^\infty e^{-Q(\mathbf{x})/2} dx_1 \dots dx_n,$$

where

$$(1.2) \quad Q(\mathbf{x}) = [1 + (n-1)\rho](1-\rho)^{-1} [1 + (n-2)\rho \sum_{i>j} x_i^2 - 2\rho \sum_{i>j} x_i x_j] \\ = (1-\rho)^{-1} [\sum x_i^2 - \rho \{1 + (n-1)\rho\}^{-1} (\sum x_i)^2].$$

Thus $I_n(h; \rho)$ is the probability that each of n normally distributed, equally correlated and standardized random variables with common correlation ρ shall not fall short of h . Clearly $1 - I_n(h; \rho)$ is also the distribution function of the random variable $\max_i x_i$, and this supplies one application (cf. [3]) of $I_n(h; \rho)$. A second application relates to the familiar one-factor model in factor analysis for the special case of equal weights [8]. Another situation in which knowledge of $I_n(h; \rho)$ is important is in some models of test design in psychology. Other applications will arise or probably exist at present.

In a previous paper [8] (see also [8] for further references), $I_n(h; \rho)$ was expressed as the product of the density function of \mathbf{x} at the cut-off point $\mathbf{h} = (h, h, \dots, h)$ and an infinite power series in h . In this paper it will be shown for $h > 0$ that $I_n(h; \rho)$ can be expressed asymptotically as the product of the density function at \mathbf{h} and an infinite series in negative powers of h . This result can be regarded as the generalization for $n > 1$ of the well-known asymptotic expansion of Mill's ratio

$$(1.3) \quad \int_x^\infty e^{-t^2/2} dt / e^{-x^2/2} \sim x^{-1} (1 - x^{-2} + 1.3x^{-4} - 1.35x^{-6} + \dots) \quad (x > 0).$$

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2. The Asymptotic Development of $I_n(h; \rho)$

Under the transformation

$$(2.1) \quad \begin{aligned} y_1 &= [1 + (n - 1)\rho]^{-1/2} \sum_{j=1}^n b_{1j}x_j, \\ y_i &= (1 - \rho)^{-1/2} \sum_{j=1}^n b_{ij}x_j, \quad (i = 2, 3, \dots, n), \end{aligned}$$

where $((b_{ij}))$, $i, j = 1, 2, \dots, n$, is orthogonal with $b_{1j} = n^{-1/2}$, (1.1) reduces to

$$(2.2) \quad I_n(h; \rho) = (2\pi)^{-n/2} \int_R \dots \int e^{-\sum y_i^2/2} dy_1 \dots dy_n$$

with R defined by

$$(2.3) \quad R: [1 + (n - 1)\rho]^{1/2} [n(1 - \rho)]^{-1/2} y_1 + \sum_{i=2}^n b_{i1}y_i \geq (1 - \rho)^{-1/2} h \quad (i = 1, 2, \dots, n)$$

[8]. R is a polyhedral half-cone in y -space with vertex at the point $(r_0, 0, 0, \dots, 0)$, where

$$(2.4) \quad r_0 = [n\{1 + (n - 1)\rho\}]^{1/2} h,$$

such that the angle between any two faces of the cone is $\arccos(-\rho)$; further, the axis of the cone passes through the origin in y -space. $I_n(h; \rho)$ is, then, the probability measure, under an n -dimensional spherical normal distribution with unit standard deviation in any direction, of a regular, symmetrically oriented polyhedral half-cone with common dihedral angle $\arccos(-\rho)$, and with vertex at a distance r_0 from the centre of the distribution. Let P be any point within the cone distant r from the centre of the distribution, η from the axis of the cone and x from the vertex of the cone in a direction parallel to the axis. The probability-mass of an infinitesimal element of volume $d\tau$ at P is

$$(2.5) \quad (2\pi)^{-n/2} e^{-r^2/2} d\tau = (2\pi)^{-1/2} e^{-(r_0+x)^2/2} dx. (2\pi)^{-(n-1)/2} e^{-\eta^2/2} dS,$$

where dS is the measure of an infinitesimal element in the $(n - 1)$ -flat orthogonal to the axis of the cone and distant x from the vertex (cf. [5]). Consider the probability-mass in that portion of the cone (an infinitesimal "slab") demarcated by two adjoining $(n - 1)$ -flats orthogonal to the axis of the cone and distant x and $x + dx$ from the vertex of the cone. It is easily shown that the intersection of the first of these two flats with the cone is a regular $(n - 1)$ -dimensional simplex with centroid at the foot of the perpendicular from P to the axis of the cone and with edges of length

$$[2n\{1 + (n - 1)\rho\}/(1 - \rho)]^{1/2} x.$$

Let $K_N(l)$ denote the probability measure, under an N -dimensional spherical normal distribution with unit standard deviation in any direction, of a regular N -dimensional simplex with centroid at the centre of the distribution and with edges of length l . Then according to (2.5) the probability measure of the infinitesimal slab is

$$(2.6) \quad (2\pi)^{-1/2} e^{-(r_0+x)^2/2} dx \cdot K_{n-1} \left[\left(\frac{2n\{1 + (n-1)\rho\}}{1-\rho} \right)^{1/2} x \right].$$

Consequently, the probability measure of the cone is

$$(2.7) \quad \begin{aligned} I_n(h; \rho) &= \int_0^\infty (2\pi)^{-1/2} e^{-(r_0+x)^2/2} K_{n-1} \left[\left(\frac{2n\{1 + (n-1)\rho\}}{1-\rho} \right)^{1/2} x \right] dx \\ &= (2\pi)^{-1/2} e^{-r_0^2/2} \int_0^\infty e^{-r_0x} \cdot e^{-x^2/2} K_{n-1}(\lambda x) dx, \end{aligned}$$

where

$$(2.8) \quad \begin{aligned} \lambda &\equiv \lambda_n(\rho) \\ &= [2n\{1 + (n-1)\rho\}/(1-\rho)]^{1/2} \end{aligned}$$

and r_0 is given by (2.4). Formula (2.7) which is of considerable intrinsic interest may be used also to develop the required asymptotic expansion of $I_n(h; \rho)$ for $h > 0$. From here on we shall then assume that $h > 0$ ¹.

The K -functions are closely related to Godwin's G -function [1], [2] introduced in connection with the distribution of the absolute mean deviation in normal samples, and some further statistical applications of the functions have been discussed in [4] and [5]. Clearly, $K_N(x)$ is bounded by 1. Again, it has been shown elsewhere [7] that $K_N(x)$ has a power series expansion with infinite radius of convergence. Consequently, Watson's lemma [10] (p. 236) may be used to obtain a valid asymptotic expansion for the integral in (2.7) by expanding $\exp(-x^2/2)K_{n-1}(\lambda x)$ in its Taylor series at $x = 0$ and integrating term by term. In fact, let

$$(2.9) \quad \psi_{n-1}(x) \equiv \psi_{n-1}(x; \lambda) \equiv e^{-x^2/2} K_{n-1}(\lambda x) = \sum_{i=0}^\infty c_{n-1,i} x^i / i!,$$

where the $c_{n-1,i}$ are functions of λ (and therefore of ρ). Then (2.7) gives with the aid of Watson's lemma,

¹ The centre of the distribution is interior or exterior to the halfcone according as to whether $h < 0$ or $h > 0$. The integral formula for $I_n(h; \rho)$ in (2.7) is valid for all h , but for the asymptotic expansion developed subsequently (equ. (2.22)) $h > 0$. ($I_n(0; \rho)$ is known to be equal to the normed measure of a regular $(n-1)$ -dimensional spherical simplex with common dihedral angle $\arccos(-\rho)$. The reader is referred to [9] where tables of such normed measures are provided for $n = 1(1) 51 - i$ and $\rho = 1/i, i = 1(1) 12$.)

$$(2.10) \quad I_n(h; \rho) \sim (2\pi)^{-1/2} e^{-r_0^2/2} \sum_{i=0}^{\infty} c_{n-1,i} r_0^{-(i+1)}.$$

This is the required formula. It should be noted that the probability density in the original distribution at the point (h, h, \dots, h) is

$$(2.11) \quad (2\pi)^{-n/2} \{1 + (n - 1)\rho\}^{-1/2} (1 - \rho)^{-(n-1)/2} e^{-r_0^2/2},$$

thereby justifying the assertion at the end of the introductory Section.

It now remains to determine the coefficients $c_{n-1,i}$ in (2.10) ($c_{n-1,i} = \psi_{n-1}^{(i)}(0)$). On differentiating (2.9) j times at $x = 0$ we obtain after some simplification

$$(2.12) \quad \begin{aligned} c_{n-1,n-1+2k} &= \sum_{s=0}^k \left(-\frac{1}{2}\right)^{k-s} \frac{(n-1+2k)!}{(k-s)!} \lambda^{n-1+2s} a_{n-1,n-1+2s} \quad (k = 0, 1, \dots), \\ c_{n-1,m} &= 0 \quad (m = 0, 1, 2, \dots, n-2), \\ c_{n-1,n+2r} &= 0 \quad (r = 0, 1, \dots), \end{aligned}$$

where the a 's are defined by

$$K_N(x) = \sum_{j=0}^{\infty} a_{N,j} x^j \quad (N = 0, 1, 2, \dots)$$

($a_{N,j} = K_N^{(j)}(0)/j!$). In the derivation of (2.12) use has been made of the fact that

$$(2.13) \quad \begin{aligned} a_{N,j} &= 0 \quad (j = 0, 1, 2, \dots, N-1), \\ a_{N,N+2r+1} &= 0 \quad (r = 0, 1, 2, \dots). \end{aligned}$$

Formula (2.13) in its turn derives by induction from the following recursion relationship between the a 's proved elsewhere [7]:

$$(2.14) \quad a_{N,s} = (2s)^{-1} \{(N+1)/(N\pi)\}^{1/2} \sum_{q=0}^{\lfloor (s-1)/2 \rfloor} \{-4N(N+1)\}^{-q} a_{N-1,s-1-2q} / q! \quad (s = 1, 2, \dots),$$

$\lfloor (s-1)/2 \rfloor$ denoting, as usual, the integral part of $(s-1)/2$. Though (2.14) may be exploited to derive explicit expressions for the non-zero a 's these are more easily obtained recursively by repeated application of (2.14) on noting that, trivially,

$$(2.15) \quad \begin{aligned} a_{0,j} &= 0 \quad (j = 1, 2, \dots), \\ &= 1 \quad (j = 0). \end{aligned}$$

This yields for the first three non-zero $a_{n-1,j}$,

$$(2.16) \quad a_{n-1,n-1} = \frac{n^{1/2}}{2^{n-1}\pi^{(n-1)/2}} \frac{1}{(n-1)!},$$

$$(2.17) \quad a_{n-1,n+1} = -\frac{n^{1/2}}{2^{n-1}\pi^{(n-1)/2}} \frac{n-1}{4} \frac{1}{(n+1)!},$$

$$(2.18) \quad a_{n-1,n+3} = \frac{n^{1/2}}{2^{n-1}\pi^{(n-1)/2}} \frac{(n-1)(n^2+7n-6)}{32n} \frac{1}{(n+3)!}$$

((2.14) shows that the non-zero a 's oscillate in sign).

On applying (2.16), (2.17) and (2.18) in (2.12), the first three non-zero c 's are obtained:

$$(2.19) \quad \begin{aligned} c_{n-1,n-1} &= (n-1)! \lambda^{n-1} a_{n-1,n-1} \\ &= n^{1/2} 2^{-(n-1)} \pi^{-(n-1)/2} \lambda^{n-1}, \end{aligned}$$

$$(2.20) \quad \begin{aligned} c_{n-1,n+1} &= (n+1)! \left\{ -\frac{1}{2} \lambda^{n-1} a_{n-1,n-1} + \lambda^{n+1} a_{n-1,n+1} \right\} \\ &= -n^{1/2} 2^{-(n-1)} \pi^{-(n-1)/2} \left\{ \frac{1}{2} n(n+1) \lambda^{n-1} + \frac{1}{4} (n-1) \lambda^{n+1} \right\}, \end{aligned}$$

$$(2.21) \quad \begin{aligned} c_{n-1,n+3} &= (n+3)! \left\{ \frac{1}{8} \lambda^{n-1} a_{n-1,n-1} - \frac{1}{2} \lambda^{n+1} a_{n-1,n+1} + \lambda^{n+3} a_{n-1,n+3} \right\} \\ &= n^{1/2} 2^{-(n-1)} \pi^{-(n-1)/2} \left\{ \frac{1}{8} n(n+1)(n+2)(n+3) \lambda^{n-1} \right. \\ &\quad \left. + \frac{1}{8} (n-1)(n+2)(n+3) \lambda^{n+1} \right. \\ &\quad \left. + \frac{1}{32} \frac{(n-1)(n^2+7n-6)}{n} \lambda^{n+3} \right\}. \end{aligned}$$

Thus from (2.10),

$$(2.22) \quad I_n(h; \rho) \sim (2\pi)^{-1/2} e^{-r_0^2/2} \left\{ c_{n-1,n-1} r_0^{-n} + c_{n-1,n+1} r_0^{-(n+2)} + c_{n-1,n+3} r_0^{-(n+4)} + \dots \right\},$$

where the first three coefficients in the asymptotic expansion are given by (2.19), (2.20) and (2.21) (further coefficients may be obtained in the manner shown). A slightly more convenient form of (2.22) is

$$(2.23) \quad \begin{aligned} I_n(h; \rho) &\sim \left(\frac{1}{2}n\right)^{1/2} \pi^{-n/2} e^{-r_0^2/2} (t/r_0)^{n-1} r_0^{-1} \\ &\times \left[1 - \left\{\frac{1}{2}(n)\right\}_2 + (n-1)t^2 r_0^{-2} \right. \\ &\left. + \left\{\frac{1}{8}(n)\right\}_4 + \frac{1}{2}(n+2)_2(n-1)t^2 + \frac{1}{2}(n-1)(n^2+7n-6)n^{-1}t^4 r_0^{-4} - \dots \right], \end{aligned}$$

where

$$(2.24) \quad \begin{aligned} t &\equiv t_n(\rho) = \lambda/2 \\ &= [n\{1 + (n-1)\rho\}/2(1-\rho)]^{1/2} \end{aligned}$$

and $(n)_m$ denotes $n(n+1)\dots(n+m-1)$. It will be noted that the present asymptotic expansion is particularly suitable for large h (i.e., the cut-off point is not near the centre of the distribution) and algebraically small ρ .

Finally, observe that for $n = 1$ (2.22) reduces to (1.3), since $\psi_0(x) = \exp(-x^2/2)$ and

$$(2.25) \quad c_{0,2j} = (-\frac{1}{2})^j (2j)! / j!$$

(The polyhedral half-cone is here the interval $[h, \infty)$.) For $n = 2$, (2.22) reduces to

$$(2.26) \quad I_2(h; \rho) \sim \pi^{-1} e^{-h^2/2} (t/r_0^2) [1 - (3+t^2)r_0^{-2} + (15 + 10t^2 + 3t^4)r_0^{-4} - \dots].$$

This agrees with a formula obtained previously [6] for the probability measure, $W(r_0; \alpha)$, $r_0 > 0$, under a standardized circular normal distribution, of a sector of angle α , vertex at a distance r_0 from the centre of the distribution and with one arm of the sector passing through the latter point. The relationship between I_2 and W is

$$(2.27) \quad I_2(h; \rho) = 2W(r_0; \theta/2)$$

where $\theta = 2 \arctan t_2(\rho) = 2 \arctan\{(1 + \rho)/(1 - \rho)\}^{1/2}$. It has been shown in [6] that the bivariate normal integral for arbitrary cut-off point may be expressed in terms of the difference of two W -functions (and therefore of two I_2 -functions).

3. The Accuracy of the Asymptotic Expansion

In this section we obtain upper bounds to the error induced by taking the first m terms of the asymptotic expansion as an approximation to $I_n(h; \rho)$.

Let ϕ be the angle between the axis of the half-cone and the line joining P and the vertex of the cone, and let ξ be the distance of P from this vertex. Then (using the notation of Section 2)

$$r^2 = r_0^2 + \xi^2 + 2r_0\xi \cos \phi,$$

and the probability-mass of an infinitesimal volume-element of content $d\tau$ at P is

$$(3.1) \quad \frac{(2\pi)^{-n/2} \exp[-\frac{1}{2}r^2] d\tau}{(2\pi)^{-n/2} \exp[-\frac{1}{2}(r_0^2 + \xi^2 + 2r_0\xi \cos \phi)] \xi^{n-1} d\xi d\omega},$$

where $d\omega$ is the solid angle subtended at the vertex of the cone by the volume-element (or, equivalently, the surface-content of an infinitesimal element on the surface of a unit sphere whose centre coincides with the vertex of the cone). Thus the probability-mass of the half-cone is

$$(3.2) \quad I_n(h; \rho) = (2\pi)^{-n/2} e^{-r_0^2/2} \int_0^\infty \int_\Omega e^{-(r_0 \cos \phi)\xi} \xi^{n-1} e^{-\xi^2/2} d\xi d\omega,$$

where Ω is the $(n - 1)$ -dimensional regular spherical simplex (with common dihedral angle arc $\cos(-\rho)$) formed by the intersection of the half-cone and the surface of the unit sphere. Again, if

$$G_{n-1}(\xi) = \xi^{n-1} e^{-\xi^2/2},$$

then the derivatives of $G_{n-1}(\xi)$ at the origin, $G_{n-1}^{(i)}(0)$, are given by

$$G_{n-1}^{(n-1+2i)}(0) = (-1)^i \frac{(n-1+2i)!}{2^i i!} \quad (i = 0, 1, 2, \dots)$$

with all other derivatives vanishing. Therefore, repeated integration by parts yields

$$(3.3) \quad \int_0^\infty e^{-(r_0 \cos \phi)\xi} G_{n-1}(\xi) d\xi = \sum_{i=0}^{m-1} (-1)^i \frac{(n-1+2i)!}{2^i i!} \frac{1}{(r_0 \cos \phi)^{n+2i}} + R_m(r_0 \cos \phi),$$

where

$$(3.4) \quad \begin{aligned} R_m(r_0 \cos \phi) &= (r_0 \cos \phi)^{-(n+2m-2)} \int_0^\infty e^{-(r_0 \cos \phi)\xi} G_{n-1}^{(n+2m-2)}(\xi) d\xi \\ &= (r_0 \cos \phi)^{-(n+2m-1)} \int_0^\infty e^{-(r_0 \cos \phi)\xi} G_{n-1}^{(n+2m-1)}(\xi) d\xi, \end{aligned}$$

after a further single integration by parts. On using (3.3) and (3.4) in (3.2),

$$(3.5) \quad \begin{aligned} I_n(h; \rho) &= (2\pi)^{-n/2} e^{-r_0^2/2} \left\{ \sum_{i=0}^{m-1} (-1)^i \frac{(n-1+2i)!}{2^i i!} \alpha_{n,i} r_0^{-(n+2i)} \right. \\ &\quad \left. + \int_\Omega R_m(r_0 \cos \phi) d\omega, \right. \end{aligned}$$

where

$$(3.6) \quad \alpha_{n,i} = \int_\Omega \sec^{n+2i} \phi d\omega.$$

In (3.5), the remainder after m terms is

$$(3.7) \quad E_m = (2\pi)^{-n/2} e^{-r_0^2/2} \int_\Omega R_m(r_0 \cos \phi) d\omega.$$

An upper bound to $|E_m|$ can be obtained from an upper bound to $R_m(r_0 \cos \phi)$ in (3.4). The latter upper bound is itself obtained by deriving first an upper bound to $G_{n-1}^{(n+2m-1)}(\xi)$ for $\xi \geq 0$. If, then,

$$(3.8) \quad |G_{n-1}^{(n+2m-1)}(\xi)| \leq A_{n-1, 2m},$$

(3.4) gives

$$(3.9) \quad |R_m(r_0 \cos \phi)| < A_{n-1, 2m}(r_0 \cos \phi)^{-(n+2m)},$$

whence by (3.7)

$$(3.10) \quad |E_m| < (2\pi)^{-n/2} e^{-r_0^2/2} A_{n-1, 2m} \int_{\Omega} (r_0 \cos \phi)^{-(n+2m)} d\omega \\ = A_{n-1, 2m} (2\pi)^{-n/2} e^{-r_0^2/2} \alpha_{n, m} r_0^{-(n+2m)},$$

which is proportional to the $(m + 1)$ th term of the series

$$(3.11) \quad (2\pi)^{-n/2} e^{-r_0^2/2} \cdot \sum_{i=0}^{\infty} (-1)^i \frac{(n - 1 + 2i)!}{2^i i!} \alpha_{n, i} r_0^{-(n+2i)}.$$

Consequently, (3.11) is a valid asymptotic expansion when $r_0 > 0$ of $I_n(h; \rho)$. Moreover, the series (3.11) must be identical with the series (2.22), since a given function determines uniquely (if at all) a series of the form $\sum c_p/r_0^p$, so that (3.10) provides an upper bound to the error in using (2.22).

We now proceed to determine a value² for $A_{n-1, 2m}$. Let

$$(3.12) \quad \xi^{n-1} = \beta_{n-1, 0} H_0(\xi) + \beta_{n-1, 1} H_1(\xi) + \dots + \beta_{n-1, n-1} H_{n-1}(\xi),$$

where $H_j(\xi)$ are the Tchebycheff-Hemite polynomials orthogonal to the weight function $\exp(-\xi^2/2)$ and normalized so that the coefficient of ξ^j in $H_j(\xi)$ is 1. On multiplying (3.12) by $H_j(\xi) \exp(-\xi^2/2)$, and integrating over the real line, we find

$$(3.13) \quad \beta_{n-1, j} = \int_{-\infty}^{\infty} \xi^{n-1} H_j(\xi) e^{-\xi^2/2} d\xi / \int_{-\infty}^{\infty} H_j^2(\xi) e^{-\xi^2/2} d\xi.$$

The value of the denominator in (3.13) is well-known to be $\sqrt{(2\pi)j!}$. In order to evaluate the numerator, define

$$\gamma_{n-1, j} = \int_{-\infty}^{\infty} \xi^{n-1} H_j(\xi) e^{-\xi^2/2} d\xi \quad (j = 0, 1, \dots, n-1).$$

Integration by parts gives the recursion relationship

$$(3.14) \quad \gamma_{n-1, j} = (n - 1)\gamma_{n-2, j-1},$$

and on successive application of (3.14)

$$\gamma_{n-1, j} = (n - 1)(n - 2) \dots (n - j)\gamma_{n-1-j, 0} \\ = (n - 1)(n - 2) \dots (n - j) \int_{-\infty}^{\infty} \xi^{n-1-j} e^{-\xi^2/2} d\xi,$$

whence

² That $A_{n-1, 2m} < \infty$ is evident from the fact that all derivatives of $G_{n-1}(\xi)$ are products of polynomials in ξ and $\exp(-\xi^2/2)$.

\sum_j' denoting summation over all non-negative integral $j \leq n - 1$ such that $n - 1 - j$ is even.

If n is odd, set $j = 2i$ in (3.20). Then

$$|G_{n-1}^{(n-1+2m)}(\xi)| \leq \pi^{-1/2} \sum_{i=0}^{(n-1)/2} \frac{(n-1)!}{2^{(n-1-2i)/2} (\frac{1}{2}(n-1) - i)! (2i)!} \cdot 2^{m+i+(n-1)/2} \Gamma(m+i+\frac{1}{2}n) \quad (n = 1, 3, \dots),$$

and, on using the gamma duplication formula in the form

$$\pi^{-1/2} \Gamma(m+i+\frac{1}{2}n) = (2m+n-1+2i)! / \{ (m+\frac{1}{2}(n-1)+i)! 2^{2m+n-1+2i} \},$$

the latter inequality simplifies to

$$(3.21) \quad |G_{n-1}^{(n-1+2m)}(\xi)| \leq \frac{(n-1)!}{2^{m+n-1}} \sum_{i=0}^{(n-1)/2} \frac{(2m+n-1+2i)!}{(m+\frac{1}{2}(n-1)+i)! (2i)! (\frac{1}{2}(n-1)-i)!} \quad (n = 1, 3, \dots).$$

Similarly if n is even, set $j = 2i + 1$ in (3.20). Then

$$(3.22) \quad |G_{n-1}^{(n-1+2m)}(\xi)| \leq \pi^{-1/2} \sum_{i=0}^{(n-2)/2} \frac{(n-1)!}{2^{(n-2-2i)/2} (\frac{1}{2}(n-2)-2i)! (2i+1)!} \cdot 2^{m+i+n/2} \Gamma(n+i+\frac{1}{2}(n+1)) \quad (n = 2, 4, \dots),$$

and, on using the gamma duplication formula in the form

$$\pi^{-1/2} \Gamma(m+i+\frac{1}{2}(n+1)) = (2m+n+2i)! / \{ (m+\frac{1}{2}n+i)! 2^{2m+n+2i} \},$$

the last inequality reduces to

$$(3.23) \quad |G_{n-1}^{(n-1+2m)}(\xi)| \leq \frac{(n-1)!}{2^{m+n-1}} \sum_{i=0}^{(n-2)/2} \frac{(2m+n+2i)!}{(m+\frac{1}{2}n+i)! (2i+1)! (\frac{1}{2}(n-2)-i)!} \quad (n = 2, 4, \dots).$$

Formula (3.21) and (3.23) provide the required inequalities in the sense that their right-hand members (refer to (3.8)) may be substituted for $A_{n-1,2m}$ in (3.10) to supply the desired upper bound for the error after m terms. A weaker (but at the same time simpler) upper bound may be obtained by noting that in (3.21)

$$(n-1)! (2m+n-1+2i)! / (2i)! = (n-1)! (2i+1)(2i+2) \cdots (2m+n-1+2i) \leq (2n-2+2m)!,$$

whence

$$\begin{aligned}
 |G_{n-1}^{(n-1+2m)}(\xi)| &\leq \frac{(2n-2+2m)!}{2^{m+n-1}} \sum_{i=0}^{(n-1)/2} \frac{1}{(m+\frac{1}{2}(n-1)+i)! (\frac{1}{2}(n-1)-i)!} \\
 &= \frac{(2n-2+2m)!}{2^{m+n-1}} \sum_{s=0}^{(n-1)/2} \frac{1}{(m+n-1-s)! s!} \\
 (3.24) \quad &= \frac{(2n-2+2m)!}{(n-1+m)!} \cdot \left(\frac{1}{2}\right)^{m+n-1} \sum_{s=0}^{(n-1)/2} \binom{n-1+m}{s} \\
 &\hspace{15em} (n=1, 3, \dots).
 \end{aligned}$$

Similarly, for n even, observe that in (3.23)

$$\begin{aligned}
 (n-1)!(2m+n+2i)!/(2i+1)! &= (n-1)!(2i+2)(2i+3) \cdots (2m+n+2i) \\
 &\leq (2n-2+2m)!,
 \end{aligned}$$

whence

$$\begin{aligned}
 |G_{n-1}^{(n-1+2m)}(\xi)| &\leq \frac{(2n-2+2m)!}{2^{m+n-1}} \sum_{i=0}^{(n-2)/2} \frac{1}{(m+\frac{1}{2}n+i)! (\frac{1}{2}(n-2)-i)!} \\
 (3.25) \quad &= \frac{(2n-2+2m)!}{2^{m+n-1}} \sum_{s=0}^{(n-2)/2} \frac{1}{(m+n-1-s)! s!} \\
 &= \frac{(2n-2+2m)!}{(n-1+m)!} \cdot \left(\frac{1}{2}\right)^{m+n-1} \sum_{s=0}^{(n-2)/2} \binom{n-1+m}{s} \\
 &\hspace{15em} (n=2, 4, \dots).
 \end{aligned}$$

The inequalities (3.24) and (3.25) may be combined in the following single inequality valid for all n (odd or even):

$$\begin{aligned}
 (3.26) \quad |G_{n-1}^{(n-1+2m)}(\xi)| &\leq \frac{(2n-2+2m)!}{(n-1+m)!} \cdot \left(\frac{1}{2}\right)^{m+n-1} \sum_{s=0}^{[(n-1)/2]} \binom{n-1+m}{s} \\
 &\hspace{15em} (n=1, 2, \dots).
 \end{aligned}$$

Thus an upper bound to the $(n-1+2m)$ th derivative of $G_{n-1}(\xi)$ is provided by the product of $(2n-2+2m)!/(n-1+m)!$ and the cumulative sum of the first (or last) $[(n+1)/2]$ probabilities in a binomial distribution with index $n-1+m$ and parameter $1/2$. (The latter cumulative sum is, of course, readily available from various statistical tables.) This upper bound may now be substituted for $A_{n-1,2m}$ in (3.10) to give the desired simplified upper bound to the error after m terms in the asymptotic expansion as a multiple of the $(m+1)$ th term.

References

- [1] H. J. Godwin, "On the distribution of the estimate of the mean deviation obtained from samples from a normal population," *Biometrika*, Vol. 33 (1945), pp. 254—256.
- [2] H. J. Godwin, "A further note on the mean deviation", *Biometrika*, Vol. 35 (1948), pp. 304—309.
- [3] A. Kuddô, "On the distribution of the maximum value of an equally correlated sample from a normal population", *Sankhyâ*, Vol. 20 (1958), pp. 309—316.
- [4] K. R. Nair, "The distribution of the extreme deviate from the sample mean and its studentised form", *Biometrika*, Vol. 35 (1948), pp. 118—144.
- [5] Harold Ruben, "Probability content of regions under spherical normal distributions, I," *Ann. Math. Stat.*, Vol. 31 (1960), pp. 598—618.
- [6] Harold Ruben, "Probability content of regions under spherical normal distributions, III: The bivariate normal integral," *Ann. Math. Stat.*, Vol. 32 (1961), pp. 171—186.
- [7] Harold Ruben, "A power series expansion for a class of Schlâfli functions," *J. London Math. Soc.*, Vol. 36 (1961), pp. 69—77.
- [8] Harold Ruben, "On the numerical evaluation of a class of multivariate normal integrals," *Proc. Roy. soc. Edin.*, Vol. 65, (1961), pp. 272—281.
- [9] Harold Ruben, "On the moments of order statistics in samples from normal populations," *Biometrika*, Vol. 41 (1954), pp. 200—227.
- 10] G. N. Watson, *Bessel Functions*. Cambridge University Press, 1922.

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