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Asymmetric Nash insurance bargaining between risk-averse parties

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Abstract

In this paper, we investigate asymmetric Nash bargaining in the context of proportional insurance contracts between a risk-averse insured and a risk-averse insurer, both seeking to enhance their expected utilities. We obtain a necessary and sufficient condition for the Pareto optimality of the status quo and derive the optimal Nash bargaining solution when the status quo is Pareto dominated. If the insured's and the insurer's risk preference exhibit decreasing absolute risk aversion and the insurer's initial wealth decreases in the insurable risk in the sense of reversed hazard rate order, we show that both the optimal insurance coverage and the optimal insurance premium increase with the insured's degree of risk aversion and the insurer's bargaining power. If the insured's risk preference further follows constant absolute risk aversion, we find that greater insurance coverage is induced as the insurer's constant initial wealth increases.

1. Introduction

Optimal insurance contract theory has been one of the research hotspots in actuarial science. An individual bargains with an insurer to design an insurance contract, which typically consists of an indemnity function (coverage) and an upfront premium. Most papers in actuarial science on optimal insurance focus on stop-loss indemnities, as they are often shown to be optimal; see, for example, Arrow (1963), Van Heerwaarden *et al.* (1989), Gollier and Schlesinger (1996), Gollier (2013), and Chi *et al.* (2024). However, a stop-loss indemnity covers losses above a pre-determined deductible, which can lead to moral hazard. Once the deductible is exceeded, the insured may lack incentives to mitigate further losses (Drèze and Schokkaert 2013). As an alternative, coinsurance is popular, where the insured covers part of the incremental losses. A common form of coinsurance is proportional insurance. Therefore, this paper focuses on the design of proportional insurance contracts.

Many optimal insurance models rely on the insurer's indifference pricing or equilibrium arguments. Boonen and Ghossoub (2023) show that competitive and Bowley equilibria make the insurer or insured indifferent between insuring or not insuring. As a direct alternative, the asymmetric Nash bargaining solution can guarantee that both parties strictly benefit from the insurance transaction (Kalai 1977). The asymmetric Nash bargaining solution, which generalizes the Nash bargaining solution introduced by Nash (1950), is characterized by Kalai (1977). More specifically, the Nash bargaining solution is obtained by maximizing the product of the excess utilities of the two parties over the status quo. The asymmetric Nash bargaining solution removes the symmetry axiom from the axiomatization of Nash (1950). This allows for assigning particular powers to the excess utilities, which can be interpreted as the bargaining power of each agent.

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In this paper, we investigate the asymmetric Nash bargaining on a proportional insurance contract between a risk-averse insured and a risk-averse insurer, where the insurer's initial wealth, subject to fluctuations from existing business operations, is considered random. Following von Neumann and Morgenstern (1953), we assume both agents maximize expected utility. We fully characterize the Pareto optimality of the status quo by introducing a condition. If this condition is not satisfied (i.e., the status quo is Pareto dominated), we derive the optimal asymmetric Nash bargaining solution. If the utility functions of the two agents exhibit decreasing absolute risk aversion (DARA), we show that both the optimal insurance coverage and the optimal insurance premium increase with the insurer's bargaining power and the degree of the insured's risk aversion when the insurer's initial wealth is decreasing with respect to the insurable risk in the sense of reversed hazard rate order. For an insured with an exponential utility function, the optimal insurance coverage increases with the insurer's constant initial wealth.

This paper is related to Chi et al. (2024), who also study Nash insurance bargaining solutions. However, we differ in two fundamental ways. First, we allow for risk aversion of the insurer, while Chi et al. (2024) restrict the insurer to be risk neutral. As a result, we get that once the insurer's initial wealth is negatively dependent with the insurable risk, full insurance cannot be optimal even if the deadweight cost is set to be zero, which contrasts with Chi et al. (2024). Second, we focus on proportional insurance treaties, while Chi et al. (2024) study stop-loss arrangements. Under the assumption of proportional insurance treaties, we can show that the set of feasible Nash bargaining outcomes is convex such that we can use the characterization of asymmetric Nash bargaining solutions in Kalai (1977). To be specific, he shows that the asymmetric Nash bargaining solution is the unique solution that satisfies the following four properties: Feasibility, Invariance under change of scale of utilities, Independence of irrelevant alternatives, and Pareto optimality. Three of these properties are clear and intuitive requirements, but only the *Independence of irrelevant alternatives* can be argued about. This property requires that if some feasible insurance contracts that are not equal to the bargaining solution are removed from our feasible set, then the solution will not change. In other words, removing feasible contracts (other than the bargaining solution) from the feasible set does not alter the solution. Notably, in the bargaining literature, it has been replaced to obtain other bargaining solution concepts, such as the Kalai–Smorodinsky solution (Kalai and Smorodinsky 1975). It should also be emphasized that Nash bargaining under other model settings has been extensively studied. Alternative characterizations of the Nash bargaining solution based on non-cooperative games are provided by van Damme (1986), Britz et al. (2010), and Okada (2010). In a risk-sharing context, Aase (2009) studies the Nash bargaining solution and compares it with the competitive equilibrium. In the context of optimal reinsurance, Boonen et al. (2016) and Anthropelos and Boonen (2020) study asymmetric Nash bargaining solutions with distortion risk measures, and Anthropelos and Boonen (2020) show that it is important for the risk measures to be known by market, as agents have an incentive to misrepresent their risk measures. Moreover, Zhou et al. (2015) and Boonen et al. (2017) study applications to longevity risk transfers with the Nash bargaining solution. Similar to these studies, this paper advances the exploration of Nash bargaining applications in insurance and risk management.

The remainder of this paper is organized as follows. Section 2 defines a proportional insurance contract and provides a brief introduction to the asymmetric Nash bargaining solution. Section 3 provides our main results on the optimal proportional insurance under asymmetric Nash bargaining. Section 4 conducts comparative statics analysis. Section 5 presents detailed examples illustrating our theoretical results. Section 6 concludes, and all the proofs are delegated to the appendix.

¹The risk attitudes of the insurer and the reinsurer may not be too dissimilar, because many major reinsurers do have subsidiaries or affiliated companies that operate as primary insurers, offering first-line insurance directly to customers. Examples of such companies include Munich Re, which owns ERGO Group, and Swiss Re, which has primary insurance operations through its Corporate Solutions unit. This means that a reinsurer can act as an insurer (indirectly) or as a reinsurer in a reinsurance transaction.

2. Model setup

2.1. A proportional insurance contract

An individual endowed with initial wealth w_0 faces an insurable risk X, interpreted as a loss. The risk X, defined on a probability space $(\Omega, \mathcal{G}, \mathbb{P})$, is a non-negative, bounded random variable with an essential infimum of 0 and an essential supremum of M > 0. To reduce her risk exposure, she purchases a proportional insurance contract with indemnity $I_{\theta}(x) := \theta x$ for some proportion $\theta \in [0, 1]$ and insurance premium P, ceding partial risk to an insurer. Note that the admissible insurance premium must be non-negative and cannot exceed M. Thus, the proportional insurance contract is completely determined by the pair $(\theta, P) \in [0, 1] \times [0, M]$. The individual is risk averse and is endowed with a utility function u satisfying $u'(\cdot) > 0$ and $u''(\cdot) < 0$ on the domain $[w_0 - 2M, w_0]$. A rational condition for purchasing this contract is that the individual's expected utility is enhanced, that is,

$$\mathbb{E}[u(w_0 - X + I_{\theta}(X) - P)] \geqslant \mathbb{E}[u(w_0 - X)].$$

Following Raviv (1979), we assume that the insurer is risk averse and endowed with bounded initial wealth W_1 and a utility function v satisfying $v'(\cdot) > 0$ and $v''(\cdot) < 0$ on the relevant domain, where W_1 is affected by the fluctuation of his existing business and may be random. The insurer will not offer this contract unless his welfare is weakly improved, that is,

$$\mathbb{E}[v(W_1 + P - (1+\tau)I_{\theta}(X))] \geqslant \mathbb{E}[v(W_1)],$$

where $\tau \geqslant 0$ is the deadweight cost rate. The factor τ is used to include the overhead costs of providing insurance by the insurer, such as marketing and administration costs.

For any proportion $\theta \in [0, 1]$, these two rationality conditions are equivalent to

$$P_{-}(\theta) \leqslant P \leqslant P_{+}(\theta), \tag{2.1}$$

where $P_{-}(\theta)$ and $P_{+}(\theta)$ are the solutions to the following equations

$$\mathbb{E}[\nu(W_1 + P - (1+\tau)I_{\theta}(X))] = \mathbb{E}[\nu(W_1)], \ P \geqslant 0, \tag{2.2}$$

and

$$\mathbb{E}[u(w_0 - X + I_\theta(X) - P)] = \mathbb{E}[u(w_0 - X)], \ P \geqslant 0, \tag{2.3}$$

respectively. In other words, the proportional insurance contract is acceptable only if the insurance premium is less than the maximum amount the insured is willing to pay and exceeds the minimum level required by the insurer. Thus, compared with the status quo (i.e., $(\theta, P) = (0, 0)$), Equation (2.1) guarantees that positive insurance is demanded only if both insured and insurer will (weakly) benefit from the insurance transaction. Note that $P_{-}(\theta)$ and $P_{+}(\theta)$ are unique due to the strict increasingness of u and v.

Clearly, the final insurance contract depends heavily on the negotiation between the insured and the insurer. In the literature, Nash bargaining is widely used to model the negotiation between two parties. Thus, we will give a brief introduction of Nash bargaining in the next subsection.

2.2. Asymmetric Nash bargaining

A two-person Nash bargaining solution was first introduced by Nash (1950) and then extended by Kalai (1977) to the asymmetric case. Let **S** be the set of all feasible bargaining outcome vectors for the two parties and is a compact convex subset of \mathbb{R}^2 . A bargaining problem is composed of a pair ($\mathbf{a_0}$, \mathbf{S}), where the 2-dimensional vector $\mathbf{a_0} \in \mathbf{S}$ represents the status quo before bargaining ("disagreement point"), and there exists at least one point $\mathbf{x} = (x_1, x_2)^T \in \mathbf{S}$ such that $x_1 > a_{01}$ and $x_2 > a_{02}$. Let \mathcal{B} denote the collection of all 2-person bargaining problems. We use a map $\mu: \mathcal{B} \mapsto \mathbb{R}^2$ to characterize the bargaining process.

²It is a reasonable assumption because the practical insurance loss is usually bounded, and this greatly simplifies the following analysis, especially in the proof of the convex set of the Nash bargaining outcomes. To relax this assumption, more technical discussions may be needed.

According to Kalai (1977), the negotiation is referred to as *asymmetric Nash bargaining* if it satisfies the following four axioms:

- Feasibility: $\mu(a_0, S) \in S$ and $\mu(a_0, S) > a_0$, where $\mu(a_0, S) > a_0$ means a strict inequality in both components.
- Invariance under change of scale of utilities: If $G: \mathbb{R}^2 \mapsto \mathbb{R}^2$ is such that $G(\mathbf{a}) = (c_1 a_1 + b_1, c_2 a_2 + b_2)^T$ where $c_i > 0$, then $G(\mu(\mathbf{a_0}, \mathbf{S})) = \mu(G(\mathbf{a_0}), G(\mathbf{S}))$.
- Independence of irrelevant alternatives: For all two bargaining problems (a_0, S) and (a_0, U) such that $S \subset U$ and $\mu(a_0, U) \in S$, it holds that $\mu(a_0, S) = \mu(a_0, U)$.
- Pareto optimality: If $\mu(\mathbf{a_0}, \mathbf{S}) = (z_1, z_2)^T$ and $y_i \ge z_i$ for i = 1, 2, then either $\mathbf{y} \notin \mathbf{S}$ or $y_i = z_i$ for i = 1, 2.

Kalai (1977) shows that μ satisfies the above four axioms if and only if there exists a $\delta \in (0, 1)$ such that $\mu(\mathbf{a_0}, \mathbf{S})$ is the unique point in \mathbf{S} satisfying

$$\mu(\mathbf{a_0}, \mathbf{S}) = \arg \max_{\mathbf{x} \in \mathbf{S}} (x_1 - a_{01})^{1-\delta} (x_2 - a_{02})^{\delta}.$$

This solution is referred to as the asymmetric Nash bargaining solution. Here, $\delta \in (0, 1)$ represents the bargaining power of the second person relative to the first person. This means that the second person has more power in the negotiation if the value of δ becomes larger (see, e.g., Kalai 1977). Feasibility is an attractive property, as it implies that the asymmetric Nash bargaining solution strictly exceeds \mathbf{a}_0 in every component. This strict inequality does not hold true for Bowley or competitive equilibria in the context of distortion risk measures, as shown by Boonen and Ghossoub (2023).

3. Nash insurance bargaining

In this paper, we analyze the insurance negotiation between a risk-averse insured and a risk-averse insurer using the asymmetric Nash bargaining framework. More specifically, the status quo is characterized by

$$\mathbf{a_0} = \begin{pmatrix} \mathbb{E}[u(w_0 - X)] \\ \mathbb{E}[v(W_1)] \end{pmatrix}.$$

Let the feasible set of insurance contracts be given by

$$\mathcal{F} = \{ (\theta, P) : \theta \in [0, 1], P_{-}(\theta) \leqslant P \leqslant P_{+}(\theta) \}.$$

Then, the set of feasible bargaining outcome vectors under proportional insurance can be given by

$$\mathbf{S} = \left\{ \begin{pmatrix} \mathbb{E}[u(w_0 - X + I_{\theta}(X) - P)] \\ \mathbb{E}[v(W_1 + P - (1 + \tau)I_{\theta}(X))] \end{pmatrix} : (\theta, P) \in \mathcal{F} \right\}.$$

Since $(0,0) \in \mathcal{F}$, it follows that $\mathbf{a}_0 \in \mathbf{S}$.

Let $\mathcal{PO} \subset \mathcal{F}$ be the class of all (θ, P) such that there is no $(\theta', P') \in \mathcal{F}$ satisfying

$$\mathbb{E}[u(w_0 - X + I_{\theta'}(X) - P')] \geqslant \mathbb{E}[u(w_0 - X + I_{\theta}(X) - P)], \tag{3.1}$$

$$\mathbb{E}[\nu(W_1 + P' - (1+\tau)I_{\theta'}(X))] \geqslant \mathbb{E}[\nu(W_1 + P - (1+\tau)I_{\theta}(X))], \tag{3.2}$$

with at least one inequality being strict. Denoting the frontier of S by ∂ S, we have

$$\partial \mathbf{S} = \left\{ \begin{pmatrix} \mathbb{E}[u(w_0 - X + I_{\theta}(X) - P)] \\ \mathbb{E}[v(W_1 + P - (1 + \tau)I_{\theta}(X))] \end{pmatrix} : (\theta, P) \in \mathcal{PO} \right\}.$$

Next, we analyze whether the status quo \mathbf{a}_0 belongs to the frontier of \mathbf{S} .

Proposition 1. Under the assumption of proportional insurance, the status quo is Pareto optimal if and only if

$$\frac{\mathbb{E}[u'(w_0 - X)X]}{\mathbb{E}[u'(w_0 - X)]} \leqslant (1 + \tau) \frac{\mathbb{E}[v'(W_1)X]}{\mathbb{E}[v'(W_1)]}.$$
(3.3)

From the above proposition, we know that no insurance will be purchased if the insured's marginal welfare improvement $\frac{\mathbb{E}[u'(w_0-X)X]}{\mathbb{E}[u'(w_0-X)]}$ is less than the marginal cost from the insurer $(1+\tau)\frac{\mathbb{E}[v'(W_1)X]}{\mathbb{E}[v'(W_1)]}$. In particular, when τ is sufficiently high, insurance becomes unattractive. Furthermore, when W_1 is independent of X, some interesting phenomena can be observed:

- The right-hand side of Equation (3.3) degenerates to $(1 + \tau)\mathbb{E}[X]$ such that this necessary and sufficient condition is no longer affected by the degree of the risk aversion of the insurer. Notably, Proposition 3 in Braun and Muermann (2004) also states that when the insurer is risk neutral, no insurance will be purchased if this condition is satisfied.
- Note that $u'(w_0 X)$ and X are comonotonic, and thus the left-hand side of Equation (3.3) is larger than $\mathbb{E}[X]$ by the Hardy-Littlewood inequality (Hardy *et al.*, 1952). This means that if $\tau = 0$, the status quo is Pareto optimal only if X is almost surely deterministic.

Further, if W_1 is stochastically increasing in X (denoted by $W_1 \uparrow_{st} X$), we have

$$\mathbb{E}[\nu'(W_1)X]/\mathbb{E}[\nu'(W_1)] \leqslant \mathbb{E}[X]$$

such that Equation (3.3) fails to be satisfied for a relatively small τ . This is intuitive: if the insurer can hedge existing business by underwriting new risks, he will provide insurance when the deadweight cost is low. For the opposite case of $W_1 \downarrow_{st} X$, it is harder to evaluate because the right-hand side of Equation (3.3) also exceeds $\mathbb{E}[X]$.

To ensure the feasibility of Nash bargaining, we impose the following assumption, which requires the deadweight cost rate τ to be sufficiently small.

Assumption 1.
$$\frac{\mathbb{E}[u'(w_0-X)X]}{\mathbb{E}[u'(w_0-X)]} > (1+\tau) \frac{\mathbb{E}[v'(W_1)X]}{\mathbb{E}[v'(W_1)]}$$

In addition to the Pareto inefficiency of the status quo, Nash bargaining problems also require the set **S** to be compact and convex. This is shown in the following proposition.

Proposition 2. The set **S** is convex and compact.

From the above proposition, Nash bargaining is feasible if Assumption 1 is satisfied. The asymmetric Nash bargaining solution solves the following optimization problem:

$$\max_{\substack{\theta \in [0,1]\\P_{-}(\theta) \leqslant P \leqslant P_{+}(\theta)}} \left\{ \mathbb{E}[\nu(W_{1} + P - (1+\tau)I_{\theta}(X))] - \mathbb{E}[\nu(W_{1})] \right\}^{\delta} \\
\times \left\{ \mathbb{E}[u(w_{0} - X + I_{\theta}(X) - P)] - \mathbb{E}[u(w_{0} - X)] \right\}^{1-\delta}$$
(3.4)

for some $\delta \in (0, 1)$, where δ represents the bargaining power of the insurer. Obviously, the optimization objective function is continuous in θ and P and equals to zero if P is equal to either $P_{-}(\theta)$ or $P_{+}(\theta)$. Thus, the Nash bargaining solutions (θ^*, P^*) must exist and satisfy

$$\theta^* \in \Theta := \{ \theta \in (0, 1] : P_+(\theta) > P_-(\theta) \} \quad \text{and} \quad P_-(\theta^*) < P^* < P_+(\theta^*)$$
 (3.5)

due to Assumption 1. The set Θ plays an important role in deriving optimal solutions, and we provide an alternative characterization in the following proposition.

Proposition 3. Under Assumption 1,

$$\Theta = \begin{cases} (0, 1], & if P_{+}(1) > P_{-}(1), \\ (0, \theta_{0}), & otherwise \end{cases}$$

for some $\theta_0 \in (0, 1]$.

³That is, the distribution of $[W_1|X=x]$ is increasing in x in the sense of usual stochastic order, where random variable Z_1 is said to be smaller than Z_2 in the usual stochastic order if $\mathbb{P}(Z_1 \leq z) \geqslant \mathbb{P}(Z_2 \leq z)$ for all $z \in \mathbb{R}$.

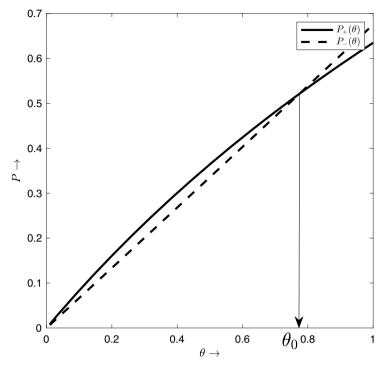


Figure 1. The functions $P_{+}(\theta)$ and $P_{-}(\theta)$ corresponding to Example 1.

Example 1. Let

$$w_0 = 20, u(w) = -w^{-1}, v(w) = w^{0.6},$$

and $W_1 = 100$ almost surely. Moreover, assume that the loss X has the cumulative distribution function

$$F_X(x) = \frac{5}{6} + \frac{4}{7} \int_0^x \frac{10^3}{(y+10)^4} dy, \ x \in [0, 10].$$

Clearly, $\mathbb{P}(X=0) = \frac{5}{6}$ and M=10. Assumption 1 is satisfied whenever $\tau < 0.817$, since $\frac{\mathbb{E}[u'(w_0-X)X]}{\mathbb{E}[u'(w_0-X)]\mathbb{E}[X]} = 1.817$.

Based on the above assumptions, it is easy to calculate $P_+(\theta)$ and $P_-(\theta)$ numerically. For $\tau=0.4$, we display these two functions in Figure 1. The set Θ can be determined for different values of the cost rate τ . More specifically, when $\tau=0.3$, we have $\theta_0=1$, and thus $\Theta=(0,1]$. The supremum of Θ changes to approximately 0.773 if the value of τ is set to be 0.4 (see Figure 1). If τ further increases to 0.5, then θ_0 decreases to approximately 0.540. A similar set Θ is obtained when the insurer's risk attitude is changed from power utility function to

$$v(w) = 1 - e^{-aw}$$

for some positive a. The outcomes of θ_0 in those experiments are presented in Table 1.

Now, we can solve the optimization problem (3.4) and obtain the Nash bargaining solutions in the following proposition.

Table 1. The supremum of Θ , given by θ_0 , when the insurer's risk preference follows an exponential utility function. Here, the value $-\infty$ corresponds to the cases that Assumption 1 is violated.

		a		
τ	0.001	0.01	0.1	1
0.3	1	0.984	0.513	$-\infty$
0.4	0.801	0.726	0.381	$-\infty$
0.5	0.559	0.507	0.268	$-\infty$

Proposition 4. Under Assumption 1, the optimal proportional insurance contract (θ^*, P^*) that solves Problem (3.4) is unique. It satisfies the first-order condition

$$\begin{cases} (1-\delta) \times \frac{\mathbb{E}[u'(w_0 - X + I_{\theta^*}(X) - P^*)]}{\mathbb{E}[u(w_0 - X + I_{\theta^*}(X) - P^*)] - \mathbb{E}[u(w_0 - X)]} = \delta \times \frac{\mathbb{E}[v'(W_1 + P^* - (1+\tau)I_{\theta^*}(X))]}{\mathbb{E}[v(W_1 + P^* - (1+\tau)I_{\theta^*}(X))] - \mathbb{E}[v(W_1)]}, \\ \frac{\mathbb{E}[u'(w_0 - X + I_{\theta^*}(X) - P^*)X]}{\mathbb{E}[u'(w_0 - X + I_{\theta^*}(X) - P^*)]} = (1+\tau) \frac{\mathbb{E}[v'(W_1 + P^* - (1+\tau)I_{\theta^*}(X))X]}{\mathbb{E}[v'(W_1 + P^* - (1+\tau)I_{\theta^*}(X))]}, \end{cases}$$
(3.6)

if

$$P_{+}(1) \leq P_{-}(1) \quad or \quad \mathbb{E}[X] < (1+\tau) \frac{\mathbb{E}\left[v'(W_{1} + \hat{P} - (1+\tau)X)X\right]}{\mathbb{E}\left[v'(W_{1} + \hat{P} - (1+\tau)X)\right]}, \tag{3.7}$$

and equals to $(1, \hat{P})$ otherwise, where \hat{P} is a solution to the following equation

$$\frac{(1-\delta)u'(w_0-P)}{u(w_0-P) - \mathbb{E}[u(w_0-X)]} = \frac{\delta \mathbb{E}[v'(W_1+P-(1+\tau)X)]}{\mathbb{E}[v(W_1+P-(1+\tau)X)] - \mathbb{E}[v(W_1)]}, P \geqslant 0.$$
(3.8)

From the above proposition, we can see that full insurance is unlikely to be optimal, except for very extreme cases such as that of a quite small τ and the insurer's risk exposure $W_1 - (1 + \tau)X$ being stochastically increasing in X. Especially when $W_1 \downarrow_{st} X$, which includes the independent case between W_1 and X, we have

$$\frac{\mathbb{E}\left[v'(W_1+\hat{P}-(1+\tau)X)X\right]}{\mathbb{E}\left[v'(W_1+\hat{P}-(1+\tau)X)\right]} > \mathbb{E}[X]$$

such that condition (3.7) is satisfied even if the deadweight cost rate τ is zero, then partial insurance is optimal. Moreover, the optimal insurance solution depends not only on the degree of risk aversion of both parties but also on the bargaining power δ .

4. Comparative statics analysis

In this section, we will carry out comparative statics analysis to investigate the effect of some interesting factors on the Nash bargaining solution.

First, we investigate the effect of the degree of the insured's risk aversion on the optimal rate and insurance premium of proportional insurance. In the literature, the insured's degree of risk aversion is often measured by Arrow-Pratt coefficient of absolute risk aversion

$$A_u(w) = -\frac{u''(w)}{u'(w)}.$$

The insured's risk preference is called to exhibit constant absolute risk aversion (CARA) if $A_u(w)$ is a constant function. It is equivalent to that the insured has an exponential utility function with $A_u(w) = a$ for some positive a.

To proceed, we consider a very special case. If there are no deadweight costs and risk preferences of the two agents in the contract exhibit CARA, then the optimal proportion has a closed-form expression and is independent of the bargaining power.

Proposition 5. Assume that the insurer's initial wealth W_1 is independent of X. For $\tau=0$ and CARA utility functions u and v with Arrow-Pratt coefficients of λ_u and λ_v , the asymmetric Nash bargaining solution is given by (θ^*, P^*) , where $\theta^* = \frac{1/\lambda_v}{1/\lambda_u + 1/\lambda_v}$ and $P^* \in (P_-(\theta^*), P_+(\theta^*))$.

The result in the above proposition can also be verified through Equation (3.6). Under these strict assumptions, we can see that the optimal proportion θ^* does not depend on the bargaining power δ and increases in λ_u . In other words, as the insured becomes more risk averse, more insurance coverage is demanded. We claim that this effect can also be held for general cases, while the optimal parameters θ^* and P^* may not have closed-form expressions. Another insured's risk attitude with increasing and concave utility function \tilde{u} is called to be more risk averse than the insured's with utility function u if $A_{\tilde{u}}(w) \geqslant A_u(w)$ for any w. Equivalently, there exists a twice differentiable increasing concave function κ such that $\tilde{u}(w) = \kappa(u(w))$; see Proposition 2 in Gollier (2001). Furthermore, the utility function u is said to exhibit a DARA risk preference if $A_u(w)$ is decreasing.

Proposition 6. Let $-W_1$ be increasing in X in the sense of hazard rate order (i.e., $-W_1 \uparrow_{hr} X$). Under the DARA assumptions on utility functions u and v, both the optimal proportion and the optimal insurance premium increase as the insured becomes more risk averse in the Arrow-Pratt sense.

The above proposition can be explained as follows: As the insured becomes more risk averse, she is willing to pay more insurance premium to mitigate the risk and transfer more risk to the insurer. Notably, a similar finding is obtained by Proposition 2 in Chi *et al.* (2024), which assumes a risk-neutral insurer. The above proposition extends the result to the case of a risk-averse insurer.

Next, we analyze the effects on the welfare improvement and the Nash bargaining solution by the change of the insurance market structure.

Proposition 7. Under Assumption 1, as the insurer becomes more powerful in the bargaining (i.e., δ increases), the increment of the insurer's expected utility increases while the welfare improvement of the insured decreases. Further, if $-W_1 \uparrow_{hr} X$ and the risk preferences of the insured and the insurer exhibit DARA, both the optimal insurance premium P^* and the optimal proportion θ^* are increasing in δ .

The above proposition indicates that the insurer will receive a larger reward from the contract negotiation as he becomes more powerful. However, this comes at the cost of a reduction in the insured's welfare improvement. As more reward is asked by the insurer, the insured has to pay more extra cost to cede the risk. Equivalently, the insured's initial wealth is relatively reduced. Under the DARA assumption of the insured's risk preference, the insured becomes more risk averse and would like to cede more risk. In other words, the increase in the insurer's bargaining power leads to more insurance demand. Thus, this proposition can be used to explain the phenomenon of the overinsurance preference observed in practice without incorporating behavioral elements (e.g., Braun and Muermann 2004). It is necessary to point out that a similar finding is also obtained by Proposition 5 in Chi *et al.* (2024) under the assumption of a risk-neutral insurer. However, we consider a risk-averse insurer with random initial wealth satisfying $-W_1 \uparrow_{hr} X$ instead.

⁴Random variable Z_1 is said to be smaller than Z_2 in the sense of hazard rate order if $\frac{\mathbb{P}(Z_1>z)}{\mathbb{P}(Z_2>z)}$ is decreasing in z. It is well-known that the hazard rate order is more strict than the usual stochastic order. Thus, $-W_1 \uparrow_{hr} X$ implies that W_1 is stochastically decreasing in X. In addition, $-W_1 \uparrow_{hr} X$ is equivalent to that W_1 is decreasing in X in the reversed hazard order. We refer to Shaked and Shanthikumar (2007) for more details on stochastic orders.

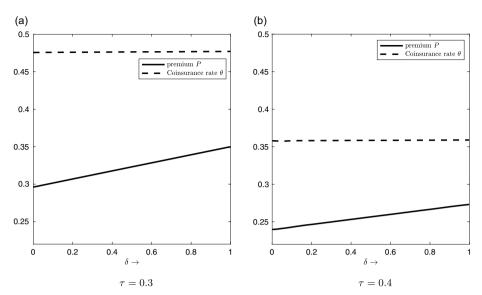


Figure 2. The optimal parameter pair (θ, P) in the asymmetric Nash bargaining solution as a function of the insurer's bargaining power $\delta \in (0, 1)$, with $\tau = 0.3$ (left) and $\tau = 0.4$ (right).

Finally, we assume that $W_1 = w_1$ almost surely and attempt to analyze the effect on the optimal insurance coverage by the change of the insurer's initial wealth w_1 .

Proposition 8. Set the insurer's initial wealth W_1 to be a constant w_1 almost surely, and assume that the insurer's risk preference exhibits DARA and that the insured has a CARA utility function u. Under Assumption 1, the optimal rate of proportional insurance θ^* increases in the insurer's initial wealth w_1 .

When the insurer's initial wealth increases, he will become less risk averse under the DARA assumption of his risk preference and charge less insurance premium. As the insurance becomes less costly, the insured would cede more risk, as stated in the above proposition.

5. Numerical examples

In this section, we provide examples to illustrate the theoretical results from the previous two sections. Specifically, we adopt the setting of Example 1 as our benchmark assumptions. That is,

$$w_0 = 20$$
, $u(w) = -w^{-1}$, $v(w) = w^{0.6}$ and $\tau = 0.4$,

and $W_1 = 100$ almost surely.

We show the insurance contract corresponding to the asymmetric Nash bargaining solution in Figure 2. In this figure, we display the optimal insurance contract (θ,P) under the asymmetric Nash bargaining for $\tau=0.3$ and $\tau=0.4$. The contracts are shown as functions of the bargaining power of the insurer δ . Consistent with Proposition 7, the numerical results show that both the coinsurance rate θ and the premium P strictly increase with δ . Thus, greater insurance coverage is induced when the insurer has more bargaining power. In addition, we note that the result in Proposition 5 may not hold if the assumptions on $\tau=0$ and the CARA risk preferences of the insured and the insurer are relaxed. The difference between the cases $\tau=0.3$ and $\tau=0.4$ can be described as follows. When the deadweight cost rate τ is lower, insurance becomes more attractive, as reflected by a larger coinsurance rate of approximately 0.48 for $\tau=0.3$, compared to approximately 0.36 when $\tau=0.4$. As the insurance coverage gets larger, the corresponding insurance premium also becomes higher.

Next, we present a sensitivity analysis to illustrate the comparative statics. For related problems involving different parameter choices for δ or minor adjustments to the relative risk-aversion parameters in u or v, we find that the pattern of the optimal insurance contract (θ, P) is roughly consistent with

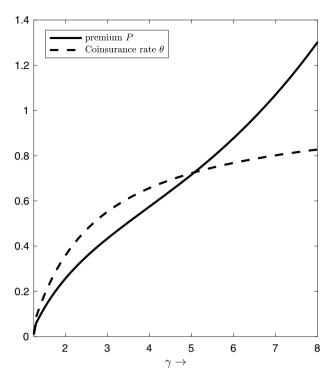


Figure 3. The optimal parameter pair (θ, P) in the asymmetric Nash bargaining solution as a function of the risk-aversion parameter γ of the insured, with $\delta = 0.5$ and $\tau = 0.4$.

Figure 2. That is, both θ and P increase with the bargaining power, albeit only slightly. We now present four additional examples, focusing on the following sensitivities: (1) the impact of the insured's risk aversion, (2) the use of a different class of utility functions, (3) the effect of the insurer's initial wealth, and (4) the effect of background risk. Throughout the following examples, we fix $\tau = 0.4$ and $w_0 = 20$.

First, we examine the effect of the insured's risk-aversion parameter. Let $W_1 = 100$ a.s., $v(w) = w^{0.6}$, and let the bargaining power be $\delta = 0.5$. The utility function u of the insured is given by

$$u(w) = \begin{cases} \frac{w^{1-\gamma}}{1-\gamma} & \text{if } \gamma \neq 1, \\ \ln(w), & \text{if } \gamma = 1, \end{cases}$$
 (5.1)

which is the constant relative risk aversion (CRRA) utility function with coefficient $\gamma > 0$. Thus, the utility function of the insurer $v(w) = w^{0.6}$ belongs to the same class, with parameter $\gamma = 0.4$. In Figure 3, we show the optimal insurance contract as a function of the risk-aversion parameter γ of the insured. We observe that as the insured becomes more risk averse, the insurance coverage increases, and the insurance premium rises disproportionately.

Second, we present a case in which the insurance coverage no longer increases in the bargaining power. We assume that the utility function of the insured is given as

$$u(w) = w - \frac{\beta}{2}w^2. {(5.2)}$$

Here, $\beta = 0.05$ controls the degree of risk aversion. We can easily verify that this utility function is increasing for wealth levels within the relevant range. However, the Arrow-Pratt coefficient of absolute risk aversion A_u is increasing in w and the DARA assumption is not met. The optimal insurance contract is displayed in Figure 4. We find that the optimal insurance coverage θ is indeed decreasing in the bargaining power, while the insurance premium is increasing. Thus, if the insurer gets more bargaining

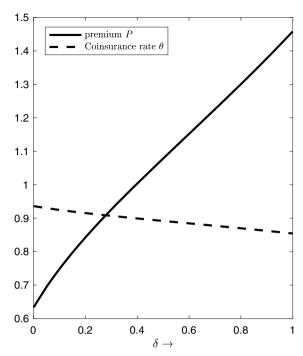


Figure 4. The optimal parameter pair (θ, P) in the asymmetric Nash bargaining solution as a function of the bargaining power $\delta \in (0, 1)$, where $\tau = 0.4$, $v(w) = w^{0.6}$, and the utility function u is as in Equation (5.2) with $\beta = 0.05$.

power, he will charge a higher premium for less coverage. This represents a clear advantage for the insurer

Third, we study the effect of the initial wealth of the insurer. Let the utility function of the insured u be given as in Equation (5.1) with coefficient $\gamma=2$. For two different utility functions v of the insurer, we show the optimal insurance contracts as a function of the initial wealth of the insurer in Figure 5. Proposition 8 shows that if v is DARA and u CARA, then under Assumption 1, the optimal rate of proportional insurance θ^* increases in the insurer's initial wealth w_1 . This pattern is also evident in Figure 5(a), where both the insured and the insurer have DARA utility functions. However, Figure 5(b) shows that this pattern does not hold when v is not a DARA utility function but instead takes the functional form in Equation (5.2) with $\beta=0.01$, which results in an increasing Arrow-Pratt coefficient of absolute risk aversion \mathcal{A}_v for relevant wealth levels.

Fourth, and finally, we study the effect of background risk of the insurer. Recall that in our baseline example, we let the utility functions of the insured u and the insurer v be given as in Equation (5.1) with $\gamma=2$ for the insured and $\gamma=0.4$ for the insurer. We model background risk via the following assumption:

$$W_1 = 100 - kX, (5.3)$$

Here, $k \in \mathbb{R}$ measures the extent of background risk. If $k \geqslant 0$, then $-W_1 \uparrow_{hr} X$. We show the optimal insurance contract as a function of $k \in [-2, 2]$ in Figure 6. We interpret a larger value of k as indicating greater background risk. Additionally, $k \geqslant 0$ implies that new insurance risk is not attractive for the insurer due to a lack of diversification opportunities. From Figure 6, we observe that larger background risk leads to less coverage and a lower corresponding insurance premium. Additionally, we do not observe any discontinuity at k = 0. Thus, a small change from positive to negative dependence, or vice versa, has little impact on the optimal insurance contract

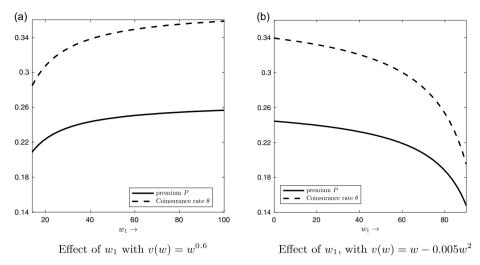


Figure 5. The optimal parameter pair (θ, P) in the asymmetric Nash bargaining solution as a function of the initial wealth of the insurer w_1 , with $\tau = 0.4$ and $\delta = 0.5$. Here, we use u as in Equation (5.1) with $\gamma = 2$.

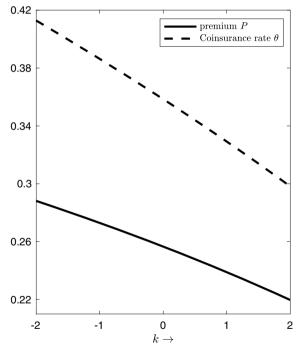


Figure 6. The optimal parameter pair (θ, P) in the asymmetric Nash bargaining solution as a function of the weight k that measures background risk (see Equation (5.3)), where $\tau = 0.4$, $\delta = 0.5$, $u(w) = -w^{-1}$, and $v(w) = w^{0.6}$.

6. Concluding remarks

In this paper, we study optimal insurance design under asymmetric Nash bargaining, assuming that both the insured and the insurer are risk averse and allowing the insurer's initial wealth to be random. To simplify the analysis, ensure the feasibility of asymmetric Nash bargaining, and facilitate comparative statics, we focus on proportional insurance contracts. After obtaining a necessary and sufficient condition for Pareto optimality of the status quo, we derive the optimal Nash bargaining solutions when this condition is not satisfied. We show that when the insurer's initial wealth is negatively dependent with the insurable risk, full insurance cannot be optimal, even if the deadweight cost is set to zero. We further find that insurance coverage increases with the insured's risk aversion or the insurer's bargaining power if the insurer's initial wealth decreases in the insurable risk in the sense of the reversed hazard rate order and both parties exhibit DARA risk preferences. In particular, when the insured has a CARA utility function, greater insurance coverage is induced by higher initial wealth of the insurer.

We acknowledge that our analysis relies heavily on the assumption of proportional insurance. It would be interesting to revisit this problem using other types of insurance contracts. However, stop-loss insurance, which is widely used in the actuarial science literature, may not be a suitable choice. This is because the corresponding set of bargaining outcomes may not be convex, which prevents the direct application of the representation theorem for asymmetric Nash bargaining in Kalai (1977). On the other hand, this paper analyzes the effect of changes in the insurer's constant initial wealth on the optimal insurance coverage under the strict assumption that the insured has a CARA utility function. It would be valuable to investigate whether this comparative statics result remains valid when the CARA assumption on the insured's risk preference is relaxed. We leave these questions to future research.

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A Proofs

A.1 Proof of Proposition 1

First, we show that the Pareto optimality of the status quo is equivalent to

$$0 \in \arg\max_{\theta \in [0,1]} \psi(\theta) := \mathbb{E}[u(w_0 - X + I_{\theta}(X) - P_{-}(\theta))], \tag{A1}$$

where $P_{-}(\theta)$ is the solution to Equation (2.2). More specifically, according to the definition of $P_{+}(\theta)$ in Equation (2.3), the Pareto optimality of the status quo is equivalent to $P_{-}(\theta) \geqslant P_{+}(\theta)$ for any $\theta \in [0, 1]$. It is further equivalent to

$$\mathbb{E}[u(w_0 - X + \theta X - P_{-}(\theta))] \leq \mathbb{E}[u(w_0 - X + \theta X - P_{+}(\theta))]$$

= $\mathbb{E}[u(w_0 - X)] = \mathbb{E}[u(w_0 - X - P_{-}(0))],$

where the last equality is derived by the fact $P_{-}(0) = 0$. In other words, 0 is a solution to the optimization problem $\max_{\theta \in [0,1]} \psi(\theta)$.

Next, we can show that the function $\psi(\theta)$ is concave. More specifically, for any $\theta_1, \theta_2 \in [0, 1]$ and any $\lambda \in [0, 1]$, the concavity of ν implies

$$\begin{split} \mathbb{E}[\nu(W_1)] &= \lambda \mathbb{E}[\nu(W_1 + P_-(\theta_1) - (1 + \tau)\theta_1 X)] + (1 - \lambda) \mathbb{E}[\nu(W_1 + P_-(\theta_2) - (1 + \tau)\theta_2 X)] \\ &\leq \mathbb{E}\left[\nu(W_1 + \lambda P_-(\theta_1) + (1 - \lambda)P_-(\theta_2) - (1 + \tau)(\lambda\theta_1 + (1 - \lambda)\theta_2)X)\right], \end{split}$$

which together with Equation (2.2) implies

$$\lambda P_{-}(\theta_1) + (1-\lambda)P_{-}(\theta_2) \geqslant P_{-}(\lambda\theta_1 + (1-\lambda)\theta_2).$$

In other words, $P_{-}(\theta)$ is convex. As a result, the concavity of u implies

$$\psi(\lambda\theta_{1} + (1-\lambda)\theta_{2}) = \mathbb{E}\left[u(w_{0} - X + (\lambda\theta_{1} + (1-\lambda)\theta_{2})X - P_{-}(\lambda\theta_{1} + (1-\lambda)\theta_{2}))\right]
\geqslant \mathbb{E}\left[u(w_{0} - X + \lambda(\theta_{1}X - P_{-}(\theta_{1})) + (1-\lambda)(\theta_{2}X - P_{-}(\theta_{2})))\right]
\geqslant \lambda\mathbb{E}\left[u(w_{0} - X + \theta_{1}X - P_{-}(\theta_{1}))\right] + (1-\lambda)\mathbb{E}\left[u(w_{0} - X + \theta_{2}X - P_{-}(\theta_{2}))\right]
= \lambda\psi(\theta_{1}) + (1-\lambda)\psi(\theta_{2}).$$

Finally, it is easy to get from Equation (2.2) and the Implicit Function Theorem that $P_{-}(\theta)$ is differentiable with

$$P'_{-}(\theta) = \frac{(1+\tau)\mathbb{E}[v'(W_1 + P_{-}(\theta) - (1+\tau)\theta X)X]}{\mathbb{E}[v'(W_1 + P_{-}(\theta) - (1+\tau)\theta X)]}$$

almost everywhere, which in turn implies

$$\begin{split} \psi'(\theta) &= \mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P_{-}(\theta))] \\ &\times \left\{ \frac{\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P_{-}(\theta))X]}{\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P_{-}(\theta))]} - \frac{(1 + \tau)\mathbb{E}[v'(W_1 + P_{-}(\theta) - (1 + \tau)\theta X)X]}{\mathbb{E}[v'(W_1 + P_{-}(\theta) - (1 + \tau)\theta X)]} \right\}. \end{split}$$

Recall that $\psi(\theta)$ is a concave function. A necessary and sufficient condition for Equation (A1) is $\lim_{\theta\downarrow 0} \psi'(\theta) \leq 0$. Equivalently,

$$\begin{split} 0 &\geqslant \lim_{\theta \downarrow 0} \left\{ \frac{\mathbb{E}[u'(w_0 - X + I_\theta(X) - P_-(\theta))X]}{\mathbb{E}[u'(w_0 - X + I_\theta(X) - P_-(\theta))]} - \frac{(1 + \tau)\mathbb{E}[v'(W_1 + P_-(\theta) - (1 + \tau)\theta X)X]}{\mathbb{E}[v'(W_1 + P_-(\theta) - (1 + \tau)\theta X)]} \right\} \\ &= \frac{\mathbb{E}[u'(w_0 - X)X]}{\mathbb{E}[u'(w_0 - X)]} - (1 + \tau) \frac{\mathbb{E}[v'(W_1)X]}{\mathbb{E}[v'(W_1)]}. \end{split}$$

It is indeed Equation (3.3). The proof is thus completed.

A.2 Proof of Proposition 2

We first show that the set **S** is compact. For any convergence sequence $\{\mathbf{a}_n \in \mathbf{S}: n = 1, 2, ...\}$, there exist $\theta_n \in [0, 1]$ and $P_n \in [P_-(\theta_n), P_+(\theta_n)]$ such that

$$\mathbf{a}_n = \begin{pmatrix} \mathbb{E}[u(w_0 - X + I_{\theta_n}(X) - P_n)] \\ \mathbb{E}[v(W_1 + P_n - (1 + \tau)I_{\theta_n}(X))] \end{pmatrix}.$$

Due to $\theta_n \in [0, 1]$ and $P_n \in [0, M]$, we can get a convergence subsequence $\{(\theta_{n_k}, P_{n_k})\}_{k=1}^{\infty}$. Without loss of generality, we denote

$$\theta = \lim_{k \to \infty} \theta_{n_k}$$
 and $P = \lim_{k \to \infty} P_{n_k}$,

then $\theta \in [0, 1]$ and $P_{-}(\theta) \leq P \leq P_{+}(\theta)$. Due to the continuity of the expected utilities with respect to the proportion and the insurance premium, we can get

$$\lim_{n\to\infty} \mathbf{a}_n = \lim_{k\to\infty} \mathbf{a}_{n_k} = \begin{pmatrix} \mathbb{E}[u(w_0 - X + I_\theta(X) - P)] \\ \mathbb{E}[v(W_1 + P - (1 + \tau)I_\theta(X))] \end{pmatrix} \in \mathbf{S}.$$

Thus, we can claim that S is compact.

Next, we proceed to prove that **S** is convex. Note that for any $\mathbf{a}_i \in \mathbf{S}$, there exist $\theta_i \in [0, 1]$ and $P_i \in [P_-(\theta_i), P_+(\theta_i)]$ such that

$$\mathbf{a}_{i} = \begin{pmatrix} \mathbb{E}[u(w_{0} - X + I_{\theta_{i}}(X) - P_{i})] \\ \mathbb{E}[v(W_{1} + P_{i} - (1 + \tau)I_{\theta_{i}}(X))] \end{pmatrix}, i = 1, 2.$$

Given any $\lambda \in (0, 1)$, we define $\tilde{P}_{-}(\theta)$ and $\tilde{P}_{+}(\theta)$ as the solutions to the following equations

$$\begin{split} & \lambda \mathbb{E}[v(W_1 + P_1 - (1 + \tau)I_{\theta_1}(X))] + (1 - \lambda)\mathbb{E}[v(W_1 + P_2 - (1 + \tau)I_{\theta_2}(X))] \\ &= \mathbb{E}[v(W_1 + P - (1 + \tau)I_{\theta}(X))], \ P \in \mathbb{R} \end{split}$$

and

$$\lambda \mathbb{E}[u(w_0 - X - P_1 + I_{\theta_1}(X))] + (1 - \lambda) \mathbb{E}[u(w_0 - X - P_2 + I_{\theta_2}(X))]$$

= $\mathbb{E}[u(w_0 - X - P + I_{\theta}(X))], P \in \mathbb{R},$

respectively for each $\theta \in [0, 1]$. Using the concavity property of u and v, we can easily find that $\tilde{P}_{-}(\theta)$ is convex while $\tilde{P}_{+}(\theta)$ is concave, and get

$$\tilde{P}_{-}(\lambda\theta_1 + (1-\lambda)\theta_2) \leqslant \lambda P_1 + (1-\lambda)P_2 \leqslant \tilde{P}_{+}(\lambda\theta_1 + (1-\lambda)\theta_2).$$

Furthermore, due to the rationality conditions (2.2) and (2.3), we obtain

$$\tilde{P}_+(0) \leqslant 0 \leqslant \tilde{P}_-(0)$$
.

More precisely, $\tilde{P}_{+}(\theta) - \tilde{P}_{-}(\theta)$ is a concave function, which is non-positive at 0 and non-negative at $\lambda \theta_{1} + (1 - \lambda)\theta_{2}$. Thus, there must exist a $\theta \in [0, \lambda \theta_{1} + (1 - \lambda)\theta_{2}]$ such that $\tilde{P}_{+}(\theta) = \tilde{P}_{-}(\theta) \geqslant \tilde{P}_{-}(0) \geqslant 0$. Setting $P = \tilde{P}_{+}(\theta)$, we can verify that the rationality condition (2.1) is satisfied by the proportional

insurance (I_{θ}, P) , and

$$\begin{pmatrix}
\mathbb{E}[u(w_0 - X + I_\theta(X) - P)] \\
\mathbb{E}[v(W_1 + P - (1 + \tau)I_\theta(X))]
\end{pmatrix} = \lambda \mathbf{a}_1 + (1 - \lambda)\mathbf{a}_2.$$

Thus, the set S is convex.

A.3 Proof of Proposition 3

From the proof of Proposition 1 in Appendix A.1, we can see that

$$\lim_{\theta \downarrow 0} \psi'(\theta) = \mathbb{E}[u'(w_0 - X)] \left(\frac{\mathbb{E}[u'(w_0 - X)X]}{\mathbb{E}[u'(w_0 - X)]} - (1 + \tau) \frac{\mathbb{E}[v'(W_1)X]}{\mathbb{E}[v'(W_1)]} \right) > 0$$

under Assumption 1. In other words, for the θ very close to 0, $\psi'(\theta) > 0$ such that

$$\psi(\theta) > \psi(0) = \mathbb{E}[u(w_0 - X)] = \mathbb{E}[u(w_0 - X + I_{\theta}(X) - P_{+}(\theta))],$$

which in turn implies $P_{-}(\theta) < P_{+}(\theta)$. Therefore, all such θ belong to the set Θ and hence the point 0 is on the boundary of Θ .

Next, we show that $P_+(\theta)$ is concave. More specifically, for any $\theta_1, \theta_2 \in [0, 1]$ and $\lambda \in [0, 1]$, we have

$$\mathbb{E}[u(w_{0} - X)]$$

$$= \lambda \mathbb{E}\left[u(w_{0} - X - P_{+}(\theta_{1}) + I_{\theta_{1}}(X))\right] + (1 - \lambda) \mathbb{E}\left[u\left(w_{0} - X - P_{+}(\theta_{2}) + I_{\theta_{2}}(X)\right)\right]$$

$$\leq \mathbb{E}[u\left(w_{0} - X - \lambda P_{+}(\theta_{1}) - (1 - \lambda)P_{+}(\theta_{2}) + I_{\lambda\theta_{1} + (1 - \lambda)\theta_{2}}(X)\right)],$$

where the inequality is derived by the concavity of u. Thus, we have

$$P_{\perp}(\lambda\theta_1 + (1-\lambda)\theta_2) \geqslant \lambda P_{\perp}(\theta_1) + (1-\lambda)P_{\perp}(\theta_2)$$

Finally, since $P_{-}(\theta)$ has been shown to be convex in Appendix A.1, then $P_{+}(\theta) - P_{-}(\theta)$ is a concave continuous function, equals to zero when $\theta = 0$, and is positive at a neighbourhood of 0. Further, if $P_{+}(1) > P_{-}(1)$, then this function is positive over (0,1] such that $\Theta = (0,1]$. Otherwise, this function has a zero point $\theta_{0} \in (0,1]$ such that it is positive over $(0,\theta_{0})$ and non-positive afterwards. For this case, it is natural to have $\Theta = (0,\theta_{0})$. The proof is thus completed.

A.4 Proof of Proposition 4

Define the set

$$\mathcal{D} := \{ (\theta, P) : \theta \in \Theta, P_{-}(\theta) < P < P_{+}(\theta) \}.$$

For any two points (θ_1, P_1) and (θ_2, P_2) in \mathcal{D} , recalling that $P_+(\theta)$ is concave and $P_-(\theta)$ is convex, we have

$$\begin{split} P_{+}(\lambda\theta_{1}+(1-\lambda)\theta_{2}) &\geqslant \lambda P_{+}(\theta_{1})+(1-\lambda)P_{+}(\theta_{2}) \\ &> \lambda P_{1}+(1-\lambda)P_{2} \\ &> \lambda P_{-}(\theta_{1})+(1-\lambda)P_{-}(\theta_{2}) \\ &\geqslant P_{-}(\lambda\theta_{1}+(1-\lambda)\theta_{2}) \end{split}$$

for any $\lambda \in (0, 1)$. We can get from Proposition 3 and the above equation that $(\lambda \theta_1 + (1 - \lambda)\theta_2, \lambda P_1 + (1 - \lambda)P_2) \in \mathcal{D}$. Thus, \mathcal{D} is a convex set.

Over the set \mathcal{D} , the optimization objective of Problem (3.4) is positive. Taking its log yields

$$\Psi(\theta, P) := \delta \ln\{\mathbb{E}[\nu(W_1 + P - (1 + \tau)I_{\theta}(X))] - \mathbb{E}[\nu(W_1)]\}$$

+ $(1 - \delta) \ln\{\mathbb{E}[\mu(w_0 - X + I_{\theta}(X) - P)] - \mathbb{E}[\mu(w_0 - X)]\}.$

Noting that $\mathbb{E}[v(W_1 + P - (1 + \tau)I_{\theta}(X))]$ and $\mathbb{E}[u(w_0 - X + I_{\theta}(X) - P)]$ are strictly concave in parameter vector (θ, P) and that the log function is an increasing concave function, we can get that $\Psi(\theta, P)$ is also strictly concave in (θ, P) such that the optimal solution to Problem (3.4) must be unique.

For any $\theta \in \Theta$, we denote by $P^*(\theta)$ the solution to

$$\max_{P_{-}(\theta) < P < P_{+}(\theta)} \Psi(\theta, P)$$

Recalling that $\Psi(\theta, P)$ is concave in (θ, P) , we can get

$$0 = \frac{\partial \Psi(\theta, P^*(\theta))}{\partial P} \\ = \frac{\delta \mathbb{E}[\nu'(W_1 + P^*(\theta) - (1 + \tau)I_{\theta}(X))]}{\mathbb{E}[\nu(W_1 + P^*(\theta) - (1 + \tau)I_{\theta}(X))] - \mathbb{E}[\nu(W_1)]} - \frac{(1 - \delta)\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P^*(\theta))]}{\mathbb{E}[u(w_0 - X + I_{\theta}(X) - P^*(\theta))] - \mathbb{E}[u(w_0 - X)]},$$

which in turn implies that $P^*(\theta)$ is an increasing function. Denoting by $\Phi(\theta) = \Psi(\theta, P^*(\theta))$, we have

$$\begin{split} & \Phi'(\theta) \\ & = \frac{(1-\delta)\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P^*(\theta))]}{\mathbb{E}[u(w_0 - X + I_{\theta}(X) - P^*(\theta))] - \mathbb{E}[u(w_0 - X)]} \\ & \times \left\{ -\frac{(1+\tau)\mathbb{E}[v'(W_1 + P^*(\theta) - (1+\tau)I_{\theta}(X))X]}{\mathbb{E}[v'(W_1 + P^*(\theta) - (1+\tau)I_{\theta}(X))]} + \frac{\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P^*(\theta))X]}{\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P^*(\theta))]} \right\} \end{split}$$

and

$$\begin{split} \Phi''(\theta) &= \frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial \theta^2} + \frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial \theta \partial P} P^{*'}(\theta) \\ &= \frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial \theta^2} - \frac{\left(\frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial \theta \partial P}\right)^2}{\frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial P^2}} \\ &= \frac{1}{\frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial P^2}} \left(\frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial \theta^2} \frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial P^2} - \left(\frac{\partial^2 \Psi(\theta, P^*(\theta))}{\partial \theta \partial P}\right)^2\right) \\ &\leq 0. \end{split}$$

where the last inequality follows from Theorem 4.5 in Rockafellar (1970). Thus, $\Phi(\theta)$ is concave. If $P_+(1) > P_-(1)$, then it follows from Proposition 3 that $1 \in \Theta$ and $\hat{P} = P^*(1)$ according to Equation (3.8). Thus, we have

$$\Phi'(1) = \frac{(1 - \delta)u'(w_0 - \hat{P})}{u(w_0 - \hat{P}) - \mathbb{E}[u(w_0 - X)]} \times \left\{ -\frac{(1 + \tau)\mathbb{E}[v'(W_1 + \hat{P} - (1 + \tau)X)X]}{\mathbb{E}[v'(W_1 + \hat{P} - (1 + \tau)X)]} + \mathbb{E}[X] \right\}.$$

If condition (3.7) is not satisfied, then $\Phi'(1) \ge 0$ such that $(\theta^*, P^*) = (1, \hat{P})$. Otherwise, the optimal insurance solution cannot appear on the boundary $\{(1, P) : P_-(1) < P < P_+(1)\}$. In other words, the optimal solution must be an interior point of \mathcal{D} and satisfy the first-order condition

$$\begin{cases} \frac{\partial \Psi(\theta, P)}{\partial P} = 0, \\ \frac{\partial \Psi(\theta, P)}{\partial \theta} = 0, \end{cases}$$
(A2)

where

$$\frac{\partial \Psi(\theta, P)}{\partial P} = \frac{\delta \mathbb{E}[v'(W_1 + P - (1 + \tau)I_{\theta}(X))]}{\mathbb{E}[v(W_1 + P - (1 + \tau)I_{\theta}(X))] - \mathbb{E}[v(W_1)]} - \frac{(1 - \delta)\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P)]}{\mathbb{E}[u(w_0 - X + I_{\theta}(X) - P)] - \mathbb{E}[u(w_0 - X)]},$$

and

$$\frac{\partial \Psi(\theta, P)}{\partial \theta} = -\frac{\delta(1+\tau)\mathbb{E}[v'(W_1+P-(1+\tau)I_{\theta}(X))X]}{\mathbb{E}[v(W_1+P-(1+\tau)I_{\theta}(X))] - \mathbb{E}[v(W_1)]} + \frac{(1-\delta)\mathbb{E}[u'(W_0-X+I_{\theta}(X)-P)X]}{\mathbb{E}[u(W_0-X+I_{\theta}(X)-P)] - \mathbb{E}[u(W_0-X)]}.$$

As a result, Equation (3.6) can be obtained by simplification. The proof is thus completed.

A.5 Proof of Proposition 5

Since it is assumed that W_1 is independent of X, then condition (3.7) is satisfied, and we can introduce an indirect utility function $\tilde{v}(w) = \mathbb{E}[v(W_1 + w)]$ for any $w \in \mathbb{R}$ such that

$$\left\{ \begin{array}{l} \mathbb{E}\left[\tilde{v}(P-(1+\tau)I_{\theta}(X))\right] = \mathbb{E}\left[v(W_1+P-(1+\tau)I_{\theta}(X))\right], \\ \mathbb{E}\left[\tilde{v}'(P-(1+\tau)I_{\theta}(X))X^k\right] = \mathbb{E}\left[v'(W_1+P-(1+\tau)I_{\theta}(X))X^k\right], \end{array} \right.$$

for k = 0, 1. The function $\tilde{v}(w)$ is increasing concave and exhibits CARA because v has the CARA property. Thus, Equation (3.6) implies that the optimal Nash bargaining solution for v(w), and W_1 is same as that for $\tilde{v}(w)$ and a deterministic W_1 .

It follows from Kalai (1977) that any asymmetric Nash bargaining solution is Pareto optimal. For Pareto optimal insurance contracts with exponential utilities and $\tau = 0$, the result is well known (cf. Example 16 in Gerber and Pafumi 1998).

A.6 Proof of Proposition 6

Since it is assumed that \tilde{u} is more risk averse than u, then there exists an increasing concave function κ such that $\tilde{u}(w) = \kappa(u(w))$. Thus, we have

$$\begin{split} \frac{\mathbb{E}[\tilde{u}'(w_0-X)X]}{\mathbb{E}[\tilde{u}'(w_0-X)]} &= \frac{\mathbb{E}[\kappa'(u(w_0-X))u'(w_0-X)X]}{\mathbb{E}[\kappa'(u(w_0-X))u'(w_0-X)]} \\ &\geqslant \frac{\mathbb{E}[u'(w_0-X)X]}{\mathbb{E}[u'(w_0-X)]}, \end{split}$$

where the last inequality is derived by

$$\begin{split} & \mathbb{E}[\kappa'(u(w_0 - X))u'(w_0 - X)X] \times \mathbb{E}[u'(w_0 - X)] \\ & - \mathbb{E}[\kappa'(u(w_0 - X))u'(w_0 - X)] \times \mathbb{E}[u'(w_0 - X)X] \\ & = \mathbb{E}\left[u'(w_0 - X)u'(w_0 - Y)\kappa'(u(w_0 - X))(X - Y)\right] \\ & = \frac{1}{2}\mathbb{E}\left[u'(w_0 - X)u'(w_0 - Y)\left(\kappa'(u(w_0 - X)) - \kappa'(u(w_0 - Y))\right)(X - Y)\right] \geqslant 0. \end{split}$$

Here, Y is an independent copy of random variable X. In other words, Assumption 1 is more likely held for \tilde{u} than for u.

Next, we assume that Assumption 1 holds for utility function u. Since it is assumed that $-W_1 \uparrow_{hr} X$, then we have $W_1 \downarrow_{st} X_1$ such that condition (3.7) is met, then the optimal insurance solution must satisfy the first-order condition. We will prove the result by using a slight modification to the proof of Proposition 2 in Chi *et al.* (2024). The proof is very lengthy and will be divided into several steps.

First, we will show that the upper premium bound $P_+^u(\theta)$ for the utility function u is smaller than that for \tilde{u} . More specifically, we have

$$\mathbb{E}[\tilde{u}(w_{0} - X + I_{\theta}(X) - P_{+}^{u}(\theta))] - \mathbb{E}[\tilde{u}(w_{0} - X)] \\
= \mathbb{E}\left[\kappa(u(w_{0} - X + I_{\theta}(X) - P_{+}^{u}(\theta))) - \kappa(u(w_{0} - X))\right] \\
\geqslant \mathbb{E}\left[\kappa'(u(w_{0} - X + I_{\theta}(X) - P_{+}^{u}(\theta))) \left(u(w_{0} - X + I_{\theta}(X) - P_{+}^{u}(\theta)) - u(w_{0} - X)\right)\right] \\
\geqslant \kappa'(u(w_{0} - x_{0} + \theta x_{0} - P_{+}^{u}(\theta))))\mathbb{E}\left[u(w_{0} - X + I_{\theta}(X) - P_{+}^{u}(\theta)) - u(w_{0} - X)\right] \\
= 0, \tag{A3}$$

where the last equality follows directly from Equation (2.3) and the second inequality is derived by the fact

$$\left\{ \kappa'(u(w_0 - X + I_{\theta}(X) - P_+^u(\theta))) - \kappa'(u(w_0 - x_0 + \theta x_0 - P_+^u(\theta))) \right\} \\
\times \left(u(w_0 - X + I_{\theta}(X) - P_+^u(\theta)) - u(w_0 - X) \right) \\
\geqslant 0.$$

Here, $x_0 := \min\{P_+^u(\theta)/\theta, M\}$. Using Equation (2.3) again, we thus have $P_+^{\tilde{u}}(\theta) \ge P_+^u(\theta)$ for any $\theta \in [0, 1]$. Thus, the set Θ for \tilde{u} is larger than that for u.

Second, we will show that

$$\frac{\mathbb{E}[\tilde{u}'(W_0(X))]}{\mathbb{E}[\tilde{u}(W_0(X))] - \mathbb{E}[\tilde{u}(w_0 - X)]} \le \frac{\mathbb{E}[u'(W_0(X))]}{\mathbb{E}[u(W_0(X))] - \mathbb{E}[u(w_0 - X)]} \tag{A4}$$

for $P \in [0, P_+^u(\theta))$, where $W_0(X) := w_0 - X + I_\theta(X) - P$. More specifically, due to Equation (A3), we have

$$\frac{\mathbb{E}[\tilde{u}'(W_0(X))]}{\mathbb{E}[\tilde{u}(W_0(X))] - \mathbb{E}[\tilde{u}(w_0 - X)]} \leq \frac{\mathbb{E}[\kappa'(u(W_0(X)))u'(W_0(X))]}{\mathbb{E}\left[\kappa'(u(W_0(X)))\left(u(W_0(X)) - u(w_0 - X)\right)\right]}.$$

Further, we have

$$\begin{split} &\left(\frac{\mathbb{E}[u'(W_{0}(X))]}{\mathbb{E}[u(W_{0}(X))]} - \frac{\mathbb{E}[\kappa'(u(W_{0}(X)))u'(W_{0}(X))]}{\mathbb{E}\left[\kappa'(u(W_{0}(X)))u'(W_{0}(X)) - u(w_{0} - X))\right]}\right) \\ &\times (\mathbb{E}[u(W_{0}(X))] - \mathbb{E}[u(w_{0} - X)]) \times \mathbb{E}\left[\kappa'(u(W_{0}(X)))(u(W_{0}(X)) - u(w_{0} - X))\right] \\ &= \mathbb{E}\left[u'(W_{0}(Y))u'(W_{0}(X))(\kappa'(u(W_{0}(X))) - \kappa'(u(W_{0}(Y))))\frac{u(W_{0}(X)) - u(w_{0} - X)}{u'(W_{0}(X))}\right] \\ &= \frac{1}{2}\mathbb{E}\Big[u'(W_{0}(Y))u'(W_{0}(X))(\kappa'(u(W_{0}(X))) - \kappa'(u(W_{0}(Y)))) \\ &\times \left(\frac{u(W_{0}(X)) - u(w_{0} - X)}{u'(W_{0}(X))} - \frac{u(W_{0}(Y)) - u(w_{0} - Y)}{u'(W_{0}(Y))}\right)\Big], \end{split}$$

where Y is an independent copy of X and the second equality is derived by the fact

$$\begin{split} &\mathbb{E}\left[u'(W_0(Y))u'(W_0(X))(\kappa'(u(W_0(X))) - \kappa'(u(W_0(Y))))\frac{u(W_0(X)) - u(w_0 - X)}{u'(W_0(X))}\right] \\ &= \mathbb{E}\left[u'(W_0(Y))u'(W_0(X))(\kappa'(u(W_0(Y))) - \kappa'(u(W_0(X))))\frac{u(W_0(Y)) - u(w_0 - Y)}{u'(W_0(Y))}\right]. \end{split}$$

It is easy to see that $\kappa'(u(W_0(X))) \ge \kappa'(u(W_0(Y)))$ whenever $X \ge Y$. Furthermore, for $x \le \min\{P/\theta, M\}$, we have $W_0(x) \le w_0 - x$ such that $u(w_0 - x) \ge u(W_0(x))$ and $u'(w_0 - x) \le u'(W_0(x))$. Therefore, when $x \le \min\{P/\theta, M\}$, we have

$$\begin{split} &\left(\frac{u(W_{0}(x)) - u(w_{0} - x)}{u'(W_{0}(x))}\right)' \\ &\geqslant \frac{(1 - \theta)}{u'(W_{0}(x))} \left\{ u'(w_{0} - x) - u'(W_{0}(x)) + \mathcal{A}_{u}(W_{0}(x)) \times (u(w_{0} - x) - u(W_{0}(x))) \right\} \\ &= \frac{(1 - \theta) \left(u(w_{0} - x) - u(W_{0}(x))\right)}{u'(W_{0}(x))} \left(\frac{u'(w_{0} - x) - u'(W_{0}(x))}{u(w_{0} - x) - u(W_{0}(x))} + \mathcal{A}_{u}(W_{0}(x))\right) \\ &= \frac{(1 - \theta) \left(u(w_{0} - x) - u(W_{0}(x))\right)}{u'(W_{0}(x))} \left\{ -\mathcal{A}_{u}(w_{0} - x + \alpha(\theta x - P)) + \mathcal{A}_{u}(w_{0} - x + (\theta x - P)) \right\} \\ &\geqslant 0 \end{split}$$

for some $\alpha \in [0, 1]$, where the last equality is obtained by using Cauchy's mean value theorem and the last inequality follows from the DARA property of utility function u. On the other hand, for any x strictly

larger than P/θ , we have $\theta x > P$ and thus $u(W_0(x)) \ge u(w_0 - x)$. Then,

$$\left(\frac{u(W_{0}(x)) - u(w_{0} - x)}{u'(W_{0}(x))}\right)' \\
\geqslant \frac{(1 - \theta)(u(W_{0}(x)) - u(w_{0} - x))}{u'(W_{0}(x))} \left\{ \mathcal{A}_{u}(w_{0} - x + \alpha(\theta x - P)) - \mathcal{A}_{u}(w_{0} - x + (\theta x - P)) \right\} \\
\geqslant 0$$

for some $\alpha \in [0, 1]$. In summary, the function $\frac{u(W_0(x)) - u(w_0 - x)}{u'(W_0(x))}$ is increasing over [0, M] such that

$$\frac{\mathbb{E}[u'(W_0(X))]}{\mathbb{E}[u(W_0(X))] - \mathbb{E}[u(w_0 - X)]} \geqslant \frac{\mathbb{E}[\kappa'(u(W_0(X)))u'(W_0(X))]}{\mathbb{E}\left[\kappa'(u(W_0(X)))\left(u(W_0(X)) - u(w_0 - X)\right)\right]}.$$

Thus, Equation (A4) can be obtained.

Third, recalling that $\Psi(\theta, P)$ is concave in P, we can get from Equations (A2) and (A4) that given an admissible θ , the optimal insurance premium for \tilde{u} is larger than that for u, that is, $P_{\tilde{u}}^*(\theta) \geqslant P_u^*(\theta)$. We further demonstrate that

$$\frac{\mathbb{E}[u'(W_0(X))X]}{\mathbb{E}[u'(W_0(X))]} \leqslant \frac{\mathbb{E}[\tilde{u}'(W_0(X))X]}{\mathbb{E}[\tilde{u}'(W_0(X))]}.$$
(A5)

More specifically, using the similar arguments as in the previous analysis, we have

$$\begin{split} &\left(\frac{\mathbb{E}[u'(W_0(X))X]}{\mathbb{E}[u'(W_0(X))]} - \frac{\mathbb{E}[\tilde{u}'(W_0(X))X]}{\mathbb{E}[\tilde{u}'(W_0(X))]}\right) \times \mathbb{E}[u'(W_0(X))] \times \mathbb{E}[\tilde{u}'(W_0(X))] \\ &= \mathbb{E}\left[u'(W_0(X))u'(W_0(Y))\kappa'(u(W_0(Y)))(X-Y)\right] \\ &= -\frac{1}{2}\mathbb{E}\left[u'(W_0(X))u'(W_0(Y))\left(\kappa'(u(W_0(X))) - \kappa'(u(W_0(Y)))\right) \times (X-Y)\right] \leqslant 0. \end{split}$$

Fourth, we prove that the function $h_2(x)/h_1(x)$ is increasing, where

$$h_2(x) := \mathbb{E}[-v''(W_1 + P - (1+\tau)I_{\theta}(X))|X = x]$$

and

$$h_1(x) := \mathbb{E}[v'(W_1 + P - (1 + \tau)I_{\theta}(X))|X = x].$$

Obviously, the function $h_2(x)$ is positive, while $h_1(x)$ is positive and increasing because v'' < 0 and $-W_1 \uparrow_{hr} X$. For any x_1, x_2 in the support of X and satisfying $x_1 < x_2$,

$$\begin{split} \frac{h_2(x_2)}{h_1(x_2)} &= \frac{\mathbb{E}[\mathcal{A}_{\nu}(W_1 + P - (1 + \tau)\theta x_2)\nu'(W_1 + P - (1 + \tau)\theta x_2)|X = x_2]}{\mathbb{E}[\nu'(W_1 + P - (1 + \tau)\theta x_2)|X = x_2]} \\ &\geqslant \frac{\mathbb{E}[\mathcal{A}_{\nu}(W_1 + P - (1 + \tau)\theta x_1)\nu'(W_1 + P - (1 + \tau)\theta x_2)|X = x_2]}{\mathbb{E}[\nu'(W_1 + P - (1 + \tau)\theta x_1)\nu'(W_1 + P - (1 + \tau)\theta x_1)|X = x_2]} \\ &\geqslant \frac{\mathbb{E}[\mathcal{A}_{\nu}(W_1 + P - (1 + \tau)\theta x_1)\nu'(W_1 + P - (1 + \tau)\theta x_1)|X = x_2]}{\mathbb{E}[\nu'(W_1 + P - (1 + \tau)\theta x_1)\nu'(W_1 + P - (1 + \tau)\theta x_1)|X = x_1]} \\ &\geqslant \frac{\mathbb{E}[\mathcal{A}_{\nu}(W_1 + P - (1 + \tau)\theta x_1)\nu'(W_1 + P - (1 + \tau)\theta x_1)|X = x_1]}{\mathbb{E}[\nu'(W_1 + P - (1 + \tau)\theta x_1)|X = x_1]} \\ &= \frac{h_2(x_1)}{h_1(x_2)}, \end{split}$$

where the first inequality is derived by the DARA property of v, the last inequality follows from the assumption $-W_1 \uparrow_{hr} X$ and Lemma 4 in Chi et al. (2025), while the second inequality can be obtained

by showing that

$$Q(x) := \frac{\mathbb{E}[A_{\nu}(W_1 + P - (1+\tau)\theta x_1)\nu'(W_1 + P - (1+\tau)\theta x)|X = x_2]}{\mathbb{E}[\nu'(W_1 + P - (1+\tau)\theta x)|X = x_2]}$$

is increasing in x over $[x_1, x_2]$. Taking the derivative of Q(x) with respect to x yields

$$\begin{split} &\frac{Q'(x)}{\theta(1+\tau)} \left(\mathbb{E}[v'(W_1+P-(1+\tau)\theta x)|X=x_2] \right)^2 \\ &= -\mathbb{E}[\mathcal{A}_{\nu}(W_1+P-(1+\tau)\theta x_1)v''(W_1+P-(1+\tau)\theta x)|X=x_2] \mathbb{E}[v'(W_1+P-(1+\tau)\theta x)|X=x_2] \\ &+ \mathbb{E}[\mathcal{A}_{\nu}(W_1+P-(1+\tau)\theta x_1)v'(W_1+P-(1+\tau)\theta x)|X=x_2] \mathbb{E}[v''(W_1+P-(1+\tau)\theta x)|X=x_2] \\ &= \mathbb{E}\left[-\mathcal{A}_{\nu}(W_1^{(1)}+P-(1+\tau)\theta x_1)v'(W_1^{(1)}+P-(1+\tau)\theta x)v'(W_1^{(2)}+P-(1+\tau)\theta x) \right] \\ &+ \mathbb{E}\left[\mathcal{A}_{\nu}(W_1^{(1)}+P-(1+\tau)\theta x_1)v'(W_1^{(1)}+P-(1+\tau)\theta x)v''(W_1^{(2)}+P-(1+\tau)\theta x) \right] \\ &= \mathbb{E}\left\{ \mathcal{A}_{\nu}(W_1^{(1)}+P-(1+\tau)\theta x_1)v'(W_1^{(1)}+P-(1+\tau)\theta x)v'(W_1^{(2)}+P-(1+\tau)\theta x) \right. \\ &\times \left. \left(\mathcal{A}_{\nu}(W_1^{(1)}+P-(1+\tau)\theta x) - \mathcal{A}_{\nu}(W_1^{(2)}+P-(1+\tau)\theta x) \right) \right\} \\ &= \frac{1}{2} \mathbb{E}\left[v'(W_1^{(1)}+P-(1+\tau)\theta x) \times \left(\mathcal{A}_{\nu}(W_1^{(1)}+P-(1+\tau)\theta x_1) - \mathcal{A}_{\nu}(W_1^{(2)}+P-(1+\tau)\theta x_1) \right) \\ &\times \left(\mathcal{A}_{\nu}(W_1^{(1)}+P-(1+\tau)\theta x) - \mathcal{A}_{\nu}(W_1^{(2)}+P-(1+\tau)\theta x) \right) \times v'(W_1^{(2)}+P-(1+\tau)\theta x) \right] \\ &\geqslant 0, \end{split}$$

where random variables $W_1^{(1)}$ and $W_1^{(2)}$ are independent and have the same distribution as $[W_1|X=x_2]$. Fifth, we can show that both functions

$$L_{\nu}(P) := -\frac{\mathbb{E}[\nu'(W_1 + P - (1+\tau)I_{\theta}(X))X]}{\mathbb{E}[\nu'(W_1 + P - (1+\tau)I_{\theta}(X))]} \quad \text{and} \quad L_{\nu}(P) := \frac{\mathbb{E}[u'(w_0 - X + I_{\theta}(X) - P)X]}{\mathbb{E}[\nu'(w_0 - X + I_{\theta}(X) - P)]}$$

are increasing in P. More specifically, it is easy to get

$$L'_{v}(P) = -\frac{\mathbb{E}[v''(W_{1}(X))X]\mathbb{E}[v'(W_{1}(X))] - \mathbb{E}[v'(W_{1}(X))X]\mathbb{E}[v''(W_{1}(X))]}{(\mathbb{E}[v'(W_{1}(X))])^{2}}$$

$$= \frac{\mathbb{E}[h_{2}(X)X]\mathbb{E}[h_{1}(X)] - \mathbb{E}[h_{1}(X)X]\mathbb{E}[h_{2}(X)]}{(\mathbb{E}[v'(W_{1}(X))])^{2}}$$

$$= \frac{\mathbb{E}\left[h_{1}(X)h_{1}(Y)(X - Y)\frac{h_{2}(X)}{h_{1}(X)}\right]}{(\mathbb{E}[v'(W_{1}(X))])^{2}}$$

$$= \frac{\mathbb{E}\left[h_{1}(X)h_{1}(Y)(X - Y)\left(\frac{h_{2}(X)}{h_{1}(X)} - \frac{h_{2}(Y)}{h_{1}(Y)}\right)\right]}{2(\mathbb{E}[v'(W_{1}(X))])^{2}}$$

$$\geqslant 0, \tag{A6}$$

where $W_1(X)$ is shorthand for $W_1 + P - (1 + \tau)I_{\theta}(X)$, Y is an independent copy of X, and the last inequality is derived by the increasing property of $h_2(x)/h_1(x)$. Using the similar arguments, we can obtain that $L_u(P)$ is increasing when u is a DARA utility function.

Finally, let $\Phi_u(\theta) := \Psi_u(\theta, P_u^*(\theta))$. Using the similar arguments as in Appendix 7.4, we can get that $\Phi_u(\theta)$ is concave. Further, the optimality of proportion θ_u^* for u implies

$$\begin{split} 0 &= \Phi_{u}'(\theta_{u}^{*}) \times \frac{\mathbb{E}[u(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))] - \mathbb{E}[u(w_{0} - X)]}{(1 - \delta)\mathbb{E}[u'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))]} \\ &= - (1 + \tau) \frac{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))X]}{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))]} + \frac{\mathbb{E}[u'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))X]}{\mathbb{E}[u'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))]} \\ &\leq - (1 + \tau) \frac{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))X]}{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))X]} + \frac{\mathbb{E}[u'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))X]}{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))X]} \\ &\leq - (1 + \tau) \frac{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))X]}{\mathbb{E}[v'(W_{1} + P_{u}^{*}(\theta_{u}^{*}) - (1 + \tau)I_{\theta_{u}^{*}}(X))X]} + \frac{\mathbb{E}[\tilde{u}'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*})X]}{\mathbb{E}[\tilde{u}'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))]} \\ &= \Phi_{\tilde{u}}'(\theta_{u}^{*}) \times \frac{\mathbb{E}[\tilde{u}(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))] - \mathbb{E}[u(w_{0} - X)]}{(1 - \delta)\mathbb{E}[\tilde{u}'(w_{0} - X + I_{\theta_{u}^{*}}(X) - P_{u}^{*}(\theta_{u}^{*}))]}, \end{split}$$

where the first inequality is derived by the increasing property of $L_{\nu}(P)$ and $L_{u}(P)$ and the fact $P_{\tilde{u}}^{*}(\theta) \geqslant P_{u}^{*}(\theta)$, while the last inequality follows from Equation (A5). Noting that $\Phi_{\tilde{u}}(\theta)$ is concave, we thus have $\theta_{\tilde{u}}^{*} \geqslant \theta_{u}^{*}$. Further, it is easy to see that $\frac{\partial \Psi_{\tilde{u}}(\theta,P)}{\partial \theta}$ is increasing in P, then

$$P_{\tilde{u}}^{*'}(\theta) = -\frac{\partial^2 \Psi_{\tilde{u}}(\theta, P_{\tilde{u}}^*(\theta))}{\partial \theta \partial P} / \frac{\partial^2 \Psi_{\tilde{u}}(\theta, P_{\tilde{u}}^*(\theta))}{\partial P^2} \geqslant 0,$$

which implies

$$P_{\tilde{u}}^* = P_{\tilde{u}}^*(\theta_{\tilde{u}}^*) \geqslant P_{\tilde{u}}^*(\theta_{u}^*) \geqslant P_{u}^*(\theta_{u}^*) = P_{u}^*.$$

That is, the optimal proportion and insurance premium increase as the insured becomes more risk averse in the Arrow-Pratt sense.

A.7 Proof of Proposition 7

First, we analyze the effect of the change of the insurer's bargaining power δ on the welfare increments of the insured and the insurer. The proof of this part is a slight modification to that of Proposition 4 in Chi *et al.* (2024). More specifically, it is trivial that the optimal solution (θ^* , P^*) to Problem (3.4) relies heavily upon the bargaining power δ . To emphasize this dependence, we rewrite (θ^* , P^*) by (θ^*_{δ} , P^*_{δ}). For any $0 < \delta_2 < \delta_1 < 1$, the optimality of (θ^*_{δ} , P^*_{δ}) can imply

$$\begin{split} &\delta_{1} \ln \left(\mathbb{E}[v(W_{1} + P_{\delta_{1}}^{*} - (1 + \tau)I_{\theta_{\delta_{1}}^{*}}(X))] - \mathbb{E}[v(W_{1})] \right) \\ &+ (1 - \delta_{1}) \ln \left(\mathbb{E}[u(w_{0} - X + I_{\theta_{\delta_{1}}^{*}}(X) - P_{\delta_{1}}^{*})] - \mathbb{E}[u(w_{0} - X)] \right) \\ &\geqslant \delta_{1} \ln \left(\mathbb{E}[v(W_{1} + P_{\delta_{2}}^{*} - (1 + \tau)I_{\theta_{\delta_{2}}^{*}}(X))] - \mathbb{E}[v(W_{1})] \right) \\ &+ (1 - \delta_{1}) \ln \left(\mathbb{E}[u(w_{0} - X + I_{\theta_{\delta_{2}}^{*}}(X) - P_{\delta_{2}}^{*})] - \mathbb{E}[u(w_{0} - X)] \right) \end{split}$$

and

$$\begin{split} &\delta_{2} \ln \left(\mathbb{E}[\nu(W_{1} + P_{\delta_{2}}^{*} - (1 + \tau)I_{\theta_{\delta_{2}}^{*}}(X))] - \mathbb{E}[\nu(W_{1})] \right) \\ &+ (1 - \delta_{2}) \ln \left(\mathbb{E}[u(w_{0} - X + I_{\theta_{\delta_{2}}^{*}}(X) - P_{\delta_{2}}^{*})] - \mathbb{E}[u(w_{0} - X)] \right) \\ &\geqslant \delta_{2} \ln \left(\mathbb{E}[\nu(W_{1} + P_{\delta_{1}}^{*} - (1 + \tau)I_{\theta_{\delta_{1}}^{*}}(X))] - \mathbb{E}[\nu(W_{1})] \right) \\ &+ (1 - \delta_{2}) \ln \left(\mathbb{E}[u(w_{0} - X + I_{\theta_{\delta_{1}}^{*}}(X) - P_{\delta_{1}}^{*})] - \mathbb{E}[u(w_{0} - X)] \right). \end{split}$$

Recalling from Equations (3.6) and (3.8) that $P_{\delta_i}^*$ satisfies

$$\begin{split} \delta_{i} \mathbb{E}[v'(W_{1} + P_{\delta_{i}}^{*} - (1+\tau)I_{\theta_{\delta_{i}}^{*}}(X))] \left(\mathbb{E}[u(w_{0} - X + I_{\theta_{\delta_{i}}^{*}}(X) - P_{\delta_{i}}^{*})] - \mathbb{E}[u(w_{0} - X)] \right) \\ = (1 - \delta_{i}) \mathbb{E}[u'(w_{0} - X + I_{\theta_{\delta_{i}}^{*}}(X) - P_{\delta_{i}}^{*})] \left(\mathbb{E}[v(W_{1} + P_{\delta_{i}}^{*} - (1+\tau)I_{\theta_{\delta_{i}}^{*}}(X))] - \mathbb{E}[v(W_{1})] \right), \end{split}$$
(A7)

we can get from the above two inequalities that

$$(1 - \delta_{1}) \ln \left(\frac{1 - \delta_{1}}{\delta_{1}}\right) + \ln \left(\mathbb{E}[\nu(W_{1} + P_{\delta_{1}}^{*} - (1 + \tau)I_{\theta_{\delta_{1}}^{*}}(X))] - \mathbb{E}[\nu(W_{1})]\right) \\ + (1 - \delta_{1}) \ln \frac{\mathbb{E}[u'(w_{0} - X + I_{\theta_{\delta_{1}}^{*}}(X) - P_{\delta_{1}}^{*})]}{\mathbb{E}[\nu'(W_{1} + P_{\delta_{1}}^{*} - (1 + \tau)I_{\theta_{\delta_{1}}^{*}}(X))]}$$

$$\geq (1 - \delta_{1}) \ln \left(\frac{1 - \delta_{2}}{\delta_{2}}\right) + \ln \left(\mathbb{E}[\nu(W_{1} + P_{\delta_{2}}^{*} - (1 + \tau)I_{\theta_{\delta_{2}}^{*}}(X))] - \mathbb{E}[\nu(W_{1})]\right) \\ + (1 - \delta_{1}) \ln \frac{\mathbb{E}[u'(w_{0} - X + I_{\theta_{\delta_{2}}^{*}}(X) - P_{\delta_{2}}^{*})]}{\mathbb{E}[\nu'(W_{1} + P_{\delta_{2}}^{*} - (1 + \tau)I_{\theta_{\delta_{1}}^{*}}(X))]},$$

which in turn implies

$$\begin{split} & \ln \frac{\mathbb{E}[\nu(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))] - \mathbb{E}[\nu(W_1)]}{\mathbb{E}[\nu(W_1 + P_{\delta_2}^* - (1 + \tau)I_{\theta_{\delta_2}^*}(X))] - \mathbb{E}[\nu(W_1)]} \\ & \geqslant (1 - \delta_1) \left\{ \ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)]}{\mathbb{E}[\nu'(W_1 + P_{\delta_2}^* - (1 + \tau)I_{\theta_{\delta_2}^*}(X))]} - \ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_1}^*}(X) - P_{\delta_1}^*)]}{\mathbb{E}[\nu'(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))]} \right\} \\ & + (1 - \delta_1) \left(\ln \left(\frac{1 - \delta_2}{\delta_2} \right) - \ln \left(\frac{1 - \delta_1}{\delta_1} \right) \right). \end{split}$$

Similarly, we can obtain

$$\begin{split} & \ln \frac{\mathbb{E}[\nu(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))] - \mathbb{E}[\nu(W_1)]}{\mathbb{E}[\nu(W_1 + P_{\delta_2}^* - (1 + \tau)I_{\theta_{\delta_2}^*}(X))] - \mathbb{E}[\nu(W_1)]} \\ & \leqslant (1 - \delta_2) \left\{ \ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)]}{\mathbb{E}[\nu'(W_1 + P_{\delta_2}^* - (1 + \tau)I_{\theta_{\delta_2}^*}(X))]} - \ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_1}^*}(X) - P_{\delta_1}^*)]}{\mathbb{E}[\nu'(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))]} \right\} \\ & + (1 - \delta_2) \left(\ln \left(\frac{1 - \delta_2}{\delta_2} \right) - \ln \left(\frac{1 - \delta_1}{\delta_1} \right) \right). \end{split}$$

Then the following results can naturally be drawn:

$$\frac{1-\delta_2}{\delta_2} \frac{\mathbb{E}[u'(w_0 - X + I_{\theta^*_{\delta_2}}(X) - P^*_{\delta_2})]}{\mathbb{E}[v'(W_1 + P^*_{\delta_2} - (1+\tau)I_{\theta^*_{\delta_2}}(X))]} \geqslant \frac{1-\delta_1}{\delta_1} \frac{\mathbb{E}[u'(w_0 - X + I_{\theta^*_{\delta_1}}(X) - P^*_{\delta_1})]}{\mathbb{E}[v'(W_1 + P^*_{\delta_1} - (1+\tau)I_{\theta^*_{\delta_2}}(X))]}$$

and

$$\mathbb{E}[\nu(W_1 + P_{\delta_1}^* - (1+\tau)I_{\theta_{\delta_1}^*}(X))] - \mathbb{E}[\nu(W_1)] \geqslant \mathbb{E}[\nu(W_1 + P_{\delta_2}^* - (1+\tau)I_{\theta_{\delta_2}^*}(X))] - \mathbb{E}[\nu(W_1)].$$

In other words, the increment of the insurer's expected utility becomes larger as he has more power in bargaining.

Using Equation (A7) again, we have

$$\ln\left(\mathbb{E}[v(W_{1} + P_{\delta_{i}}^{*} - (1+\tau)I_{\theta_{\delta_{i}}^{*}}(X))] - \mathbb{E}[v(W_{1})]\right) + \ln\frac{1-\delta_{i}}{\delta_{i}}$$

$$= \ln\left(\mathbb{E}[u(w_{0} - X + I_{\theta_{\delta_{i}}^{*}}(X) - P_{\delta_{i}}^{*})] - \mathbb{E}[u(w_{0} - X)]\right) - \ln\frac{\mathbb{E}[u'(w_{0} - X + I_{\theta_{\delta_{i}}^{*}}(X) - P_{\delta_{i}}^{*})]}{\mathbb{E}[v'(W_{1} + P_{\delta_{i}}^{*} - (1+\tau)I_{\theta_{\delta_{i}}^{*}}(X))]},$$

which together with the first inequality of this proof implies

$$\begin{split} &\ln\left(\mathbb{E}[u(w_0 - X + I_{\theta^*_{\delta_1}}(X) - P^*_{\delta_1})] - \mathbb{E}[u(w_0 - X)]\right) \\ &- \delta_1 \ln\frac{\mathbb{E}[u'(w_0 - X + I_{\theta^*_{\delta_1}}(X) - P^*_{\delta_1})]}{\mathbb{E}[v'(W_1 + P^*_{\delta_1} - (1 + \tau)I_{\theta^*_{\delta_1}}(X))]} - \delta_1 \ln\left(\frac{1 - \delta_1}{\delta_1}\right) \\ &\geqslant \ln\left(\mathbb{E}[u(w_0 - X + I_{\theta^*_{\delta_2}}(X) - P^*_{\delta_2})] - \mathbb{E}[u(w_0 - X)]\right) \\ &- \delta_1 \ln\frac{\mathbb{E}[u'(w_0 - X + I_{\theta^*_{\delta_2}}(X) - P^*_{\delta_2})]}{\mathbb{E}[v'(W_1 + P^*_{\delta_2} - (1 + \tau)I_{\theta^*_{\delta_3}}(X))]} - \delta_1 \ln\left(\frac{1 - \delta_2}{\delta_2}\right). \end{split}$$

Equivalently, we have

$$\begin{split} & \ln \frac{\mathbb{E}[u(w_0 - X + I_{\theta_{\delta_1}^*}(X) - P_{\delta_1}^*)] - \mathbb{E}[u(w_0 - X)]}{\mathbb{E}[u(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)] - \mathbb{E}[u(w_0 - X)]} \\ & \geqslant \delta_1 \left[\ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)]}{\mathbb{E}[v'(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))]} - \ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)]}{\mathbb{E}[v'(W_1 + P_{\delta_2}^* - (1 + \tau)I_{\theta_{\delta_2}^*}(X))]} + \ln \left(\frac{\frac{1 - \delta_1}{\delta_1}}{\frac{1 - \delta_2}{\delta_2}} \right) \right]. \end{split}$$

In a similar way, we have

$$\begin{split} & \ln \frac{\mathbb{E}[u(w_0 - X + I_{\theta_{\delta_1}^*}(X) - P_{\delta_1}^*)] - \mathbb{E}[u(w_0 - X)]}{\mathbb{E}[u(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)] - \mathbb{E}[u(w_0 - X)]} \\ & \leqslant \delta_2 \left[\ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_1}^*}(X) - P_{\delta_1}^*)]}{\mathbb{E}[v'(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))]} - \ln \frac{\mathbb{E}[u'(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)]}{\mathbb{E}[v'(W_1 + P_{\delta_1}^* - (1 + \tau)I_{\theta_{\delta_1}^*}(X))]} + \ln \left(\frac{\frac{1 - \delta_1}{\delta_1}}{\frac{1 - \delta_2}{\delta_2}}\right) \right]. \end{split}$$

As a consequence, it can naturally get that

$$\mathbb{E}[u(w_0 - X + I_{\theta_{\delta_1}^*}(X) - P_{\delta_1}^*)] - \mathbb{E}[u(w_0 - X)] \leqslant \mathbb{E}[u(w_0 - X + I_{\theta_{\delta_2}^*}(X) - P_{\delta_2}^*)] - \mathbb{E}[u(w_0 - X)].$$

That is to say, the amount of the insured's welfare improvement is reduced as the bargaining power δ becomes larger.

Next, we will analyze how the change of the insurer's bargaining power δ affects the Nash bargaining solution (θ^*, P^*) . Since it is assumed that $-W_1 \uparrow_{hr} X$, then condition (3.7) is met such that (θ^*, P^*) satisfies the first-order condition. Taking the derivative of the first-order conditions in Equation (A2) with respect to δ yields

$$\begin{cases} 0 = \frac{\partial^2 \Psi(\theta^*, P^*)}{\partial \theta \partial P} \frac{\partial \theta^*}{\partial \delta} + \frac{\partial^2 \Psi(\theta^*, P^*)}{\partial P^2} \frac{\partial P^*}{\partial \delta} + \frac{\mathbb{E}[v'(W_1^*(X))]}{\mathbb{E}[v(W_1^*(X))] - \mathbb{E}[v(W_1)]} + \frac{\mathbb{E}[u'(W_0^*(X))]}{\mathbb{E}[u(W_0^*(X))] - \mathbb{E}[u(w_0 - X)]}, \\ 0 = \frac{\partial^2 \Psi(\theta^*, P^*)}{\partial \theta^2} \frac{\partial \theta^*}{\partial \delta} + \frac{\partial^2 \Psi(\theta^*, P^*)}{\partial \theta \partial P} \frac{\partial P^*}{\partial \delta} - \frac{(1 + \tau)\mathbb{E}[v'(W_1^*(X))X]}{\mathbb{E}[v(W_1^*(X))] - \mathbb{E}[v(W_1)]} - \frac{\mathbb{E}[u'(W_0^*(X))X]}{\mathbb{E}[u(W_0^*(X))] - \mathbb{E}[u(w_0 - X)]}, \end{cases}$$

where $W_0^*(X) := w_0 - X + I_{\theta^*}(X) - P^*$ and $W_1^*(X) := W_1 + P^* - (1 + \tau^*)I_{\theta^*}(X)$. The above equation together with Equation (3.6) can imply

$$\begin{split} &\left\{\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta^{2}}\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P^{2}} - \left(\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta\partial P}\right)^{2}\right\}\frac{\partial\theta^{*}}{\partial\delta} \\ &= \frac{1}{\delta\left(\mathbb{E}[u(W_{0}^{*}(X))] - \mathbb{E}[u(w_{0}-X)]\right)}\left\{\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta\partial P}\mathbb{E}[u'(W_{0}^{*}(X))] + \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P^{2}}\mathbb{E}[u'(W_{0}^{*}(X))X]\right\} \\ &= \frac{(1-\delta)\mathbb{E}[u'(W_{0}^{*}(X))]\mathbb{E}[u'(W_{0}^{*}(X))X]}{\delta\left(\mathbb{E}[u(W_{0}^{*}(X))] - \mathbb{E}[u(w_{0}-X)]\right)^{2}} \\ &\times \left\{\frac{\mathbb{E}[u''(W_{0}^{*}(X))] - \mathbb{E}[u(w_{0}-X)]}{\mathbb{E}[u'(W_{0}^{*}(X))X]} + \frac{\mathbb{E}[v''(W_{1}^{*}(X))]}{\mathbb{E}[v'(W_{1}^{*}(X))]} - \frac{\mathbb{E}[v''(W_{1}^{*}(X))X]}{\mathbb{E}[v'(W_{1}^{*}(X))X]}\right\} \end{split}$$

and

$$\begin{split} &\left\{ \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta^{2}} \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P^{2}} - \left(\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta\partial P} \right)^{2} \right\} \frac{\partial P^{*}}{\partial\delta} \\ &= - \frac{\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta^{2}} \mathbb{E}[u'(W_{0}^{*}(X))] + \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P\partial\theta} \mathbb{E}[u'(W_{0}^{*}(X))X]}{\delta \left(\mathbb{E}[u(W_{0}^{*}(X))] - \mathbb{E}[u(w_{0} - X)] \right)} \\ &= \frac{(1 - \delta)\mathbb{E}[u'(W_{0}^{*}(X))] \mathbb{E}[u'(W_{0}^{*}(X))X]}{\delta \left(\mathbb{E}[u(W_{0}^{*}(X))X] - \mathbb{E}[u(w_{0} - X)] \right)^{2}} \\ &\times \left\{ \frac{\mathbb{E}[u''(W_{0}^{*}(X))X]}{\mathbb{E}[u'(W_{0}^{*}(X))X]} - \frac{\mathbb{E}[u''(W_{0}^{*}(X))X^{2}]}{\mathbb{E}[u'(W_{0}^{*}(X))X]} + (1 + \tau) \left(\frac{\mathbb{E}[v''(W_{1}^{*}(X))X]}{\mathbb{E}[v'(W_{1}^{*}(X))X]} - \frac{\mathbb{E}[v''(W_{1}^{*}(X))X^{2}]}{\mathbb{E}[v'(W_{1}^{*}(X))X]} \right) \right\}. \end{split}$$

Under the DARA assumption of the utility function u, we have

$$\begin{split} &\mathbb{E}[u''(W_0^*(X))]\mathbb{E}[u'(W_0^*(X))X] - \mathbb{E}[u''(W_0^*(X))X]\mathbb{E}[u'(W_0^*(X))] \\ &= \mathbb{E}\left[\mathcal{A}_u(W_0^*(X))u'(W_0^*(X))u'(W_0^*(Y))(X-Y)\right] \\ &= \frac{1}{2}\mathbb{E}\left[u'(W_0^*(X))u'(W_0^*(Y))(X-Y)(\mathcal{A}_u(W_0^*(X))) - \mathcal{A}_u(W_0^*(Y))\right] \geqslant 0, \end{split}$$

where Y is an independent copy of X and the last inequality is derived by the fact that $W_0^*(X)$ is decreasing in X. Using the similar arguments, we have

$$\mathbb{E}[u''(W_0^*(X))X^2]\mathbb{E}[u'(W_0^*(X))] \leqslant \mathbb{E}[u''(W_0^*(X))X]\mathbb{E}[u'(W_0^*(X))X].$$

Furthermore, we define

$$h_2^*(x) := \mathbb{E}[-v''(W_1^*(X))|X=x]$$
 and $h_1^*(x) := \mathbb{E}[v'(W_1^*(X))|X=x]$.

Using the similar arguments as in the proof of Proposition 6 in Appendix A.6, we get that both $\frac{h_2^*(x)}{h_1^*(x)}$ and $h_1^*(x)$ are positive and increasing under the assumptions of a DARA v and $-W_1 \uparrow_{hr} X_1$. Thus, similar to Equation (A6), we have

$$\mathbb{E}[\nu''(W_1^*(X))X]\mathbb{E}[\nu'(W_1^*(X))] \leqslant \mathbb{E}[\nu''(W_1^*(X))]\mathbb{E}[\nu'(W_1^*(X))X] \tag{A8}$$

and

$$\begin{split} &\mathbb{E}[v''(W_1^*(X))X^2]\mathbb{E}[v'(W_1(X))] - \mathbb{E}[v''(W_1^*(X))X]\mathbb{E}[v'(W_1^*(X))X] \\ &= -\mathbb{E}\left[h_2^*(X)X^2\right]\mathbb{E}[h_1^*(Y)] + \mathbb{E}\left[h_2^*(X)X\right]\mathbb{E}[h_1^*(Y)Y] \\ &= \mathbb{E}\left[h_1^*(X)h_1^*(Y)\left(Y - X\right) \times \frac{h_2^*(X)X}{h_1^*(X)}\right] \\ &= -\frac{1}{2}\mathbb{E}\left[h_1^*(X)h_1^*(Y)\left(X - Y\right) \times \left(\frac{h_2^*(X)X}{h_1^*(X)} - \frac{h_2^*(Y)Y}{h_1^*(Y)}\right)\right] \\ &\leqslant 0. \end{split}$$

Recalling that

$$\frac{\partial^2 \Psi(\theta^*,P^*)}{\partial \theta^2} \frac{\partial^2 \Psi(\theta^*,P^*)}{\partial P^2} - \left(\frac{\partial^2 \Psi(\theta^*,P^*)}{\partial \theta \partial P}\right)^2 > 0,$$

we thus have $\frac{\partial \theta^*}{\partial \delta} \geqslant 0$ and $\frac{\partial P^*}{\partial \delta} \geqslant 0$. The proof is thus completed.

A.8 Proof of Proposition 8

Noticing that $W_1 = w_1$ almost surely, we get that condition (3.7) is met, then (θ^* , P^*) satisfies the first-order condition. Taking the derivative of the first-order conditions in Equation (A2) with respect to w_1 yields

$$\begin{cases} 0 = \frac{\partial^2 \Psi(\theta^*, P^*)}{\partial \theta \partial P} \frac{\partial \theta^*}{\partial w_1} + \frac{\partial^2 \Psi(\theta^*, P)}{\partial P^2} \frac{\partial P^*}{\partial w_1} \\ + \frac{\delta \left\{ \mathbb{E}[v''(W_1^*(X))](\mathbb{E}[v(W_1^*(X))] - v(w_1)) - \mathbb{E}[v'(W_1^*(X))](\mathbb{E}[v'(W_1^*(X))] - v'(w_1)) \right\}}{(\mathbb{E}[v(W_1^*(X))] - v(w_1))^2}, \\ 0 = \frac{\partial^2 \Psi(\theta^*, P^*)}{\partial \theta^2} \frac{\partial \theta^*}{\partial w_1} + \frac{\partial^2 \Psi(\theta^*, P)}{\partial \theta \partial P} \frac{\partial P^*}{\partial w_1} \\ - \frac{\delta (1 + \tau) \left\{ \mathbb{E}[v''(W_1^*(X))X](\mathbb{E}[v(W_1^*(X))] - v(w_1)) - \mathbb{E}[v'(W_1^*(X))X](\mathbb{E}[v'(W_1^*(X))] - v'(w_1)) \right\}}{(\mathbb{E}[v(W_1^*(X))] - v(w_1))^2}. \end{cases}$$

Since it is assumed that the insured has a CARA utility function u, then there exists a $\lambda > 0$ such that $u''(w) = -\lambda u'(w)$, which in turn implies

$$\lambda = -\frac{\mathbb{E}[u''(W_0^*(X))X]}{\mathbb{E}[u'(W_0^*(X))X]} = -\frac{\mathbb{E}[u''(W_0^*(X))]}{\mathbb{E}[u'(W_0^*(X))]}.$$

Thus, we have

$$\begin{split} &\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P^{2}} \\ &= \frac{\delta\mathbb{E}[v'(W_{1}^{*}(X))]}{\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1})} \left\{ \frac{\mathbb{E}[v''(W_{1}^{*}(X))]}{\mathbb{E}[v'(W_{1}^{*}(X))]} + \frac{\mathbb{E}[u''(W_{0}^{*}(X))]}{\mathbb{E}[u'(W_{0}^{*}(X))]} - \frac{\mathbb{E}[v'(W_{1}^{*}(X))]}{(1 - \delta)(\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1}))} \right\} \\ &= \frac{\delta\mathbb{E}[v'(W_{1}^{*}(X))]}{\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1})} \left\{ \frac{\mathbb{E}[v''(W_{1}^{*}(X))]}{\mathbb{E}[v'(W_{1}^{*}(X))]} - \lambda - \frac{\mathbb{E}[v'(W_{1}^{*}(X))]}{(1 - \delta)(\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1}))} \right\} \end{split}$$

and

$$\begin{split} &\frac{\partial^2 \Psi(\theta^*, P^*)}{\partial \theta \partial P} \\ &= \frac{\delta(1+\tau) \mathbb{E}[v'(W_1^*(X))X]}{\mathbb{E}[v(W_1^*(X))] - v(w_1)} \left\{ -\frac{\mathbb{E}[v''(W_1^*(X))X]}{\mathbb{E}[v'(W_1^*(X))X]} - \frac{\mathbb{E}[u''(W_0^*(X))X]}{\mathbb{E}[u'(W_0^*(X))X]} + \frac{\mathbb{E}[v'(W_1^*(X))]}{(1-\delta)(\mathbb{E}[v(W_1^*(X))] - v(w_1))} \right\} \\ &= \frac{\delta(1+\tau) \mathbb{E}[v'(W_1^*(X))X]}{\mathbb{E}[v(W_1^*(X))] - v(w_1)} \left\{ -\frac{\mathbb{E}[v''(W_1^*(X))X]}{\mathbb{E}[v'(W_1^*(X))X]} + \lambda + \frac{\mathbb{E}[v'(W_1^*(X))]}{(1-\delta)(\mathbb{E}[v(W_1^*(X))] - v(w_1))} \right\}. \end{split}$$

The above equations together with Equation (3.6) can imply

$$\begin{split} &\frac{1}{\delta} \left\{ \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta^{2}} \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P^{2}} - \left(\frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta\partial P} \right)^{2} \right\} \frac{\partial\theta^{*}}{\partial w_{1}} \\ &= \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial\theta\partial P} \frac{\left\{ \mathbb{E}[v''(W_{1}^{*}(X))](\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1})) - \mathbb{E}[v'(W_{1}^{*}(X))](\mathbb{E}[v'(W_{1}^{*}(X))] - v'(w_{1})) \right\}}{(\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1}))^{2}} \\ &+ \frac{(1+\tau) \left\{ \mathbb{E}[v''(W_{1}^{*}(X))X](\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1})) - \mathbb{E}[v'(W_{1}^{*}(X))X](\mathbb{E}[v'(W_{1}^{*}(X))] - v'(w_{1})) \right\}}{(\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1}))^{2}} \\ &\times \frac{\partial^{2}\Psi(\theta^{*},P^{*})}{\partial P^{2}} \\ &= \frac{\delta(1+\tau)}{(\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1}))^{3}} \left\{ \mathbb{E}[v''(W_{1}^{*}(X))]\mathbb{E}[v'(W_{1}^{*}(X))X] - \mathbb{E}[v''(W_{1}^{*}(X))X]\mathbb{E}[v'(W_{1}^{*}(X))] \right\} \\ &\times \left\{ \lambda(\mathbb{E}[v(W_{1}^{*}(X))] - v(w_{1})) + \frac{\delta}{1-\delta}\mathbb{E}[v'(W_{1}^{*}(X))] + v'(w_{1}) \right\} \\ &\geqslant 0, \end{split}$$

where the last inequality is derived by Equation (A8). Recalling that $\frac{\partial^2 \Psi(\theta^*,P^*)}{\partial \theta^2} \frac{\partial^2 \Psi(\theta^*,P^*)}{\partial P^2} - \left(\frac{\partial^2 \Psi(\theta^*,P^*)}{\partial \theta \partial P}\right)^2 > 0$, we thus obtain $\frac{\partial \theta^*}{\partial w_1} \ge 0$. In other words, the insurance demand is increasing in the insurer's constant initial wealth when the insured's risk preference exhibits CARA.