

APPLICATION OF AN ESTIMATOR-FREE INFORMATION CRITERION (*WIC*) TO APERTURE SYNTHESIS IMAGING

M. ISHIGURO

The Institute of Statistical Mathematics, 4-6-7 Minami Azabu, Minato-ku, Tokyo, 106 Japan

K.I. MORITA

Nobeyama Radio Observatory, National Astronomical observatory, Nobeyama, MinamiSaku Nagano, 384-13

M. ISHIGURO

Nobeyama Radio Observatory, National Astronomical observatory, Nobeyama, Minami Maki-mura, Minami Saku-gun, Nagano-ken, 384-13 Japan

ABSTRACT A statistical criterion for stopping *CLEAN* procedure is proposed. The criterion is called *WIC* and an estimator of Kullback-Leibler information quantity which is a measure of the goodness of statistical models. A numerical example is given.

INTRODUCTION

Visibility Data

Visibility z_j observed at time j can be expressed in terms of “true visibility” defined by the brightness distribution $T(x, y)$ and observation noise ε_j :

$$z_j = \tilde{T}(u_j, v_j) + \varepsilon_j \quad (j = 1, 2, \dots, N) \quad (1)$$

$$\tilde{T}(u, v) = \iint \exp\{2\pi i(xu + yv)\} T(x, y) dx dy, \quad (2)$$

where $\{(u_j, v_j)\}$ is a set of points in the u - v plane. If ε_j can be regarded as a random variable, then the vector $z = (z_1, z_2, \dots, z_N)$ is an N dimensional complex random variable. Let $f(z)$ denote the probability density function (*PDF*) of z .

Statistical Modeling

Assume that a model of $T(x, y)$ is given by

$$T_M(x, y) = \sum_{m=1}^M w_m K(x - x_m, y - y_m), \quad (3)$$

where $K(x, y)$ is a suitably fixed kernel function.

Substituting this equation for $T(x, y)$ in Eq. (1), and assuming $\{\varepsilon_j\}$ be a mutually independent complex normal random numbers with mean 0 and variance σ^2 , a statistical model of data z is defined which is parametrized by $(w_1, x_1, y_1), (w_2, x_2, y_2), \dots, (w_M, x_M, y_M)$ and σ^2 . Let us denote this model by $f(z|(w_1, x_1, y_1), (w_2, x_2, y_2), \dots, (w_M, x_M, y_M), \sigma^2)$.

Image formation

When an estimate $\{(\hat{w}_1, \hat{x}_1, \hat{y}_1), (\hat{w}_2, \hat{x}_2, \hat{y}_2), \dots, (\hat{w}_M, \hat{x}_M, \hat{y}_M)\}$ is given, it is natural to estimate σ^2 by

$$\hat{\sigma}^2 = \frac{1}{2N} \sum_{j=1}^N |z_j - \hat{T}_M(u_j, v_j)|^2, \quad (4)$$

where $\hat{T}_M(u, v)$ is defined by substituting estimate $T_M(x, y)$ in Eq.(2). The model of data z , $f(z|(\hat{w}_1, \hat{x}_1, \hat{y}_1), (\hat{w}_2, \hat{x}_2, \hat{y}_2), \dots, (\hat{w}_M, \hat{x}_M, \hat{y}_M), \hat{\sigma}^2)$ should be an estimate of the true *PDF* $f(z)$.

INFORMATION CRITERION

The goodness of statistical models is measured by the expected log likelihood defined by

$$ELL = \int f(z) \log f(z|(\hat{w}_1, \hat{x}_1, \hat{y}_1), (\hat{w}_2, \hat{x}_2, \hat{y}_2), \dots, (\hat{w}_M, \hat{x}_M, \hat{y}_M), \hat{\sigma}^2) dz. \quad (5)$$

A larger value of *ELL* means a better fit of a model.

Ishiguro & Ishiguro (1980) successfully applied the *ELL* minimization procedure to the aperture synthesis data analysis. They employed the maximum likelihood method to the model fitting, and used *AIC* (Akaike, 1973) as the estimator of *ELL*. The example was the one-dimensional imaging problem. The extension of the proposed method, however, to the two-dimensional imaging was difficult. Two dimensional models are too complex to be fitted by the maximum likelihood method and it is a vital condition for the use of *AIC* that the parameter estimation is done by the maximum likelihood method.

WIC

We propose the use of *WIC*, an estimator-free information criterion, when the model fitting is not done by the maximum likelihood method. *WIC* is an Extension of *AIC* which is defined by

$$\begin{aligned}
 WIC = & -2 \times \log f(z | (\hat{w}_1, \hat{x}_1, \hat{y}_1), (\hat{w}_2, \hat{x}_2, \hat{y}_2), \dots, (\hat{w}_M, \hat{x}_M, \hat{y}_M), \hat{\sigma}^2) \\
 & + 2 \times E_{z^*} \{ \log f(z^* | (\hat{w}_1^*, \hat{x}_1^*, \hat{y}_1^*), (\hat{w}_2^*, \hat{x}_2^*, \hat{y}_2^*), \dots, (\hat{w}_M^*, \hat{x}_M^*, \hat{y}_M^*), \hat{\sigma}^2) \\
 & - \log f(z | (\hat{w}_1^*, \hat{x}_1^*, \hat{y}_1^*), (\hat{w}_2^*, \hat{x}_2^*, \hat{y}_2^*), \dots, (\hat{w}_M^*, \hat{x}_M^*, \hat{y}_M^*), \hat{\sigma}^2) \} \quad (6)
 \end{aligned}$$

where z^* is a simulated data obtained by the resampling technique $\{(\hat{w}_1^*, \hat{x}_1^*, \hat{y}_1^*), (\hat{w}_2^*, \hat{x}_2^*, \hat{y}_2^*), \dots, (\hat{w}_M^*, \hat{x}_M^*, \hat{y}_M^*)\}$ and $\hat{\sigma}^2$ are estimates of parameters based on the data z^* . E_{z^*} denotes the expectation with respect to the variation of z^* , which is practically computed by the Monte Carlo method.

A NUMERICAL EXAMPLE

We applied the *WIC* minization procedure for the choice of the number of *CLEAN* component.

A set of data was produced, assuming the structure shown in Fig. 1a. The *CLEAN* process proceeds as shown in Figs. 1d through 1f. The *WIC* values indicate that we should stop around $M = 600$. *ELL* values which could be computed only when the data is artificially generated supports this choice. Visual inspection also endorses this choice.

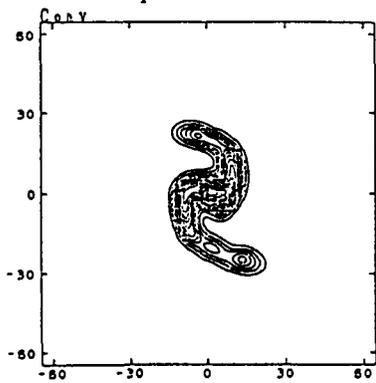


Fig. 1a. Source Model.

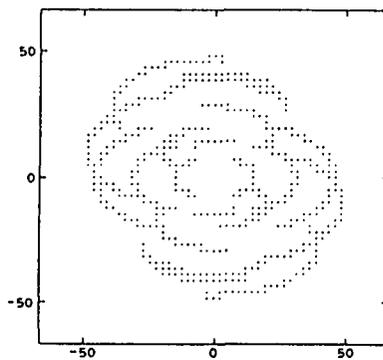


Fig. 1b. $U - V$ plane.

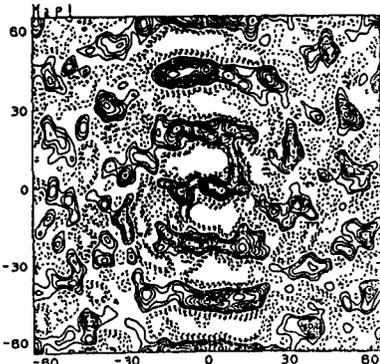


Fig. 1c. Synthesized Map.

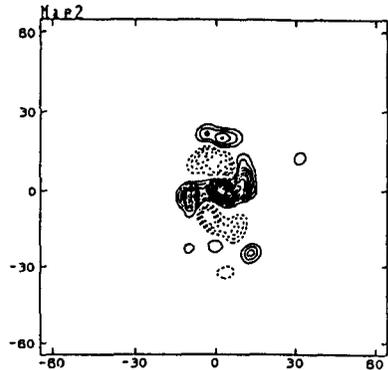


Fig. 1d. Clean Map ($M=200$).

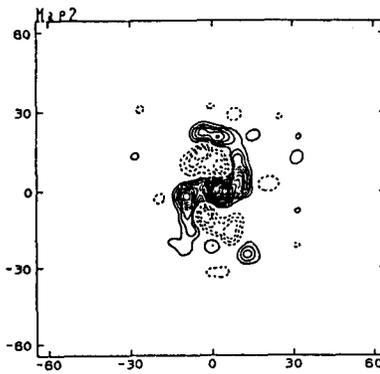


Fig. 1e. Clean Map ($M=600$).

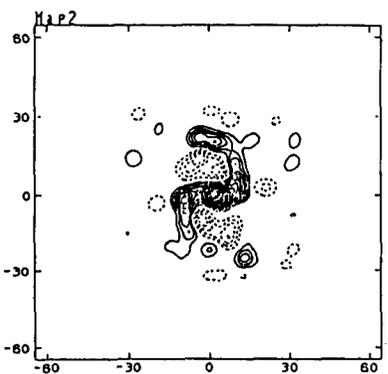


Fig. 1f. Clean Map ($M=1000$).

TABLE II Analysis of complex data

M	WIC	$-2ELL$
100	29647.9	29643.5
200	29424.2	29411.0
300	29334.6	29312.8
400	29298.2	29269.0
500	29284.2	29248.1
600	29280.6	29238.2
700	29281.6	29235.0
800	29284.3	29235.3
900	29287.5	29235.2
1000	29291.1	29237.2

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REFERENCES

- Akaike, H. 1973, in *Proc. 2nd Inter. Symp. on Information Theory* ed B. N. Petyov and F. Csai, Akademiai Kiado, Budapest, pp.267-281.
- Efron, B. 1979, *The Annals of Statistics*, 7, 1, pp.1-26.
- Sakamoto, Y., Ishiguro, M. & Kitagawa, G. 1986, *Akaike Information Criterion Statistics*, D.Reidel Publishing Company, Dordrecht/Tokyo.
- Schwartz, U. J. 1977, *Main Journal*.
- Wong, W. H. 1983, *JASA*, 78, 382, pp.461-463.

APPENDIX: A SHORT DICTIONARY OF STATISTICAL TERMS

AIC: An estimate of Information Criterion proposed by Akaike. It is defined by

$$AIC = -2 \log \max_{\theta} f(x | \theta) + 2 \{ \text{number of free parameters} \},$$

where f is a statistical model of data x with parameter θ . It is a measure of the badness of the statistical models.

Expected log likelihood: Let $f(x)$ and $f(x|\theta)$ denotes the true *PDF* of data x and its model, respectively. Then the expected log likelihood is defined by

$$\int f(x) \log f(x|\theta) dx.$$

The larger value of the expected log likelihood means a better fit of the model to the real *PDF*.

Log likelihood: Let $f(x|\theta)$ be a statistical model. If x is fixed at the value of given data, f can be regarded as the function of parameter θ only. This function is called *likelihood function* of θ . *Log likelihood function* is the logarithm of likelihood function.

MLE: Maximum likelihood estimate. The estimated parameter of a statistical model obtained by maximizing the (log) likelihood function.

Resampling: When we have a set of data, and need to generate similar set of data, there are two possible way. One is to fit a *PDF* model to the data and generate simulation data using this fitted *PDF*. The other way is the "resampling". Suppose that we have the set of data

$$\{d_1, d_2, \dots, d_N\}$$

then the resampled data is obtained by randomly choosing N values from this set of data replacing the chosen value every time. Formal expression of the resampled data could be

$$\{d_{j(1)}, d_{j(2)}, \dots, d_{j(N)}\},$$

where $\{j(1), j(2), \dots, j(N)\}$ is a sequence of independent random numbers uniformly distributed on the interval $[1, N]$.

Statistical Model: Usually a parametrized family of probability (density) function.